

Title: Numerical Methods Lecture

Speakers: Erik Schnetter

Collection: Numerical Methods 2023/24

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Markov Chain Monte Carlo

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PSI Numerical Methods 2024

Borrowing heavily from Dan Foreman-Mackey's slides
<https://speakerdeck.com/dfm/data-analysis-with-mcmc>

These slides are available at
<https://github.com/dstndstn/MCMC-talk>

1

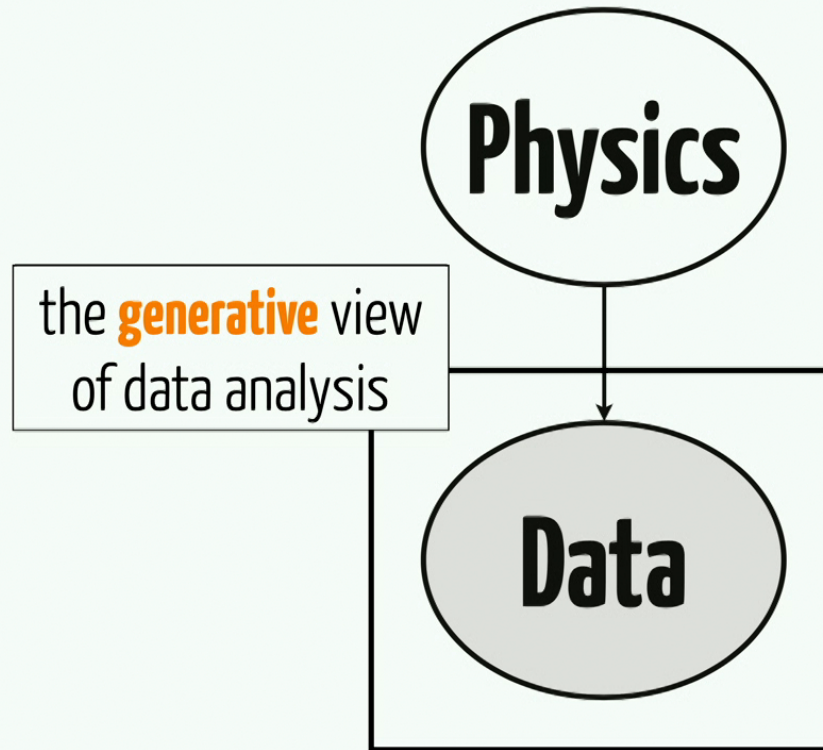
data analysis with

Markov chain Monte Carlo

Dan Foreman-Mackey

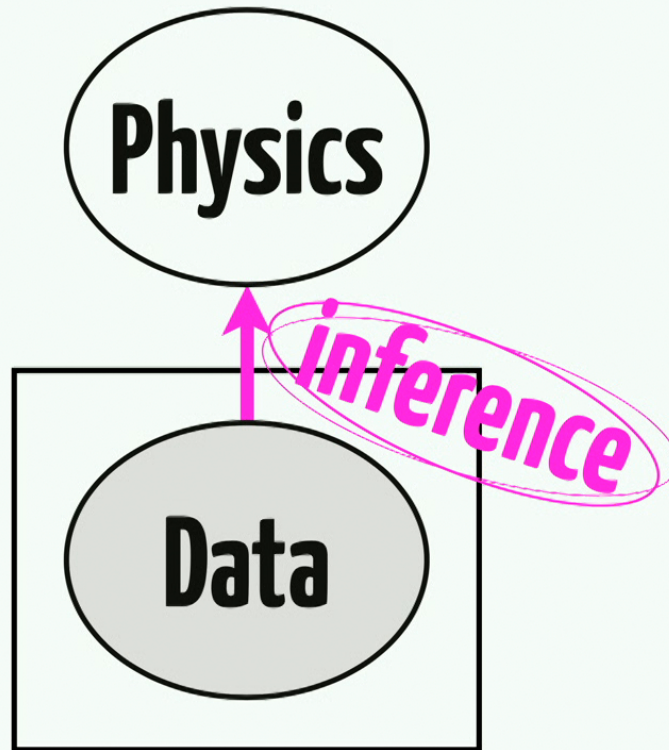
CCPP@NYU

2



a sketch of
The graphical model of my research.

3



a sketch of
The graphical model of my research.

4

$p(\text{data} \mid \text{physics})$
likelihood function/generative model

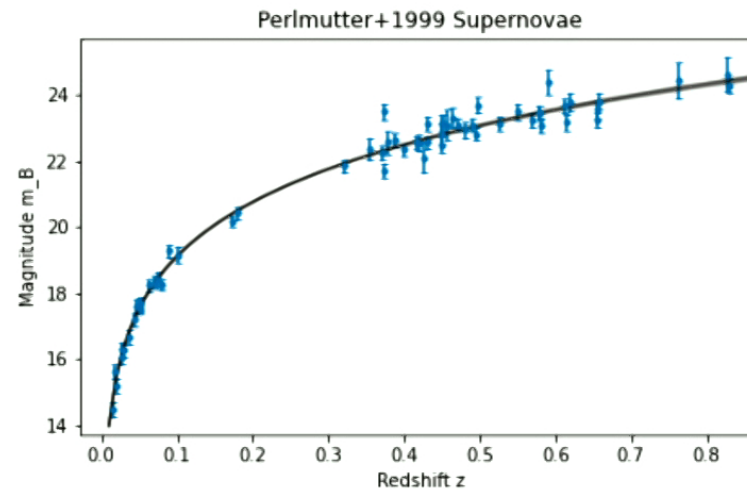


$p(\text{physics} \mid \text{data}) \propto p(\text{physics}) p(\text{data} \mid \text{physics})$
posterior probability

5

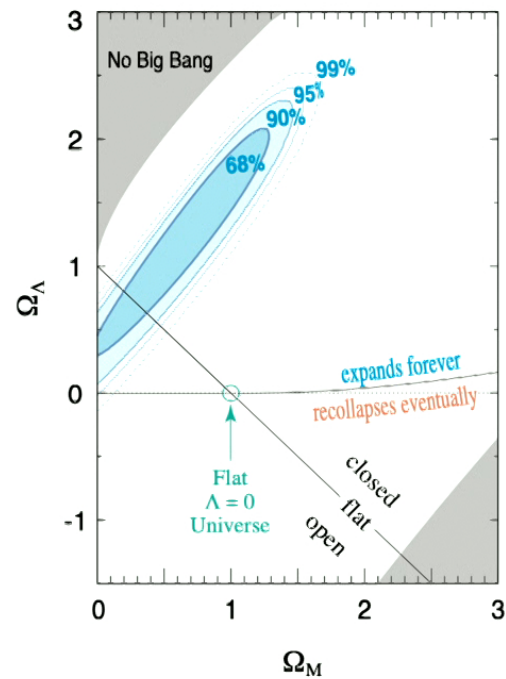
An example

- ▶ Use Bayes' theorem to convert data likelihoods into constraints on the model parameters $\theta = \{\Omega_M, \Omega_\Lambda\}$
- ▶ $p(\theta | \text{data}) \propto p(\theta) p(\text{data} | \theta)$
- ▶ $p(\Omega_M, \Omega_\Lambda | \{\text{mag}_i\}) \propto p(\Omega_M, \Omega_\Lambda) \prod_i \mathcal{N}(\text{mag}_i | \text{mag}_{\text{int}} + D_L(z_i, \Omega_M, \Omega_\Lambda), \sigma_i^2)$



An example

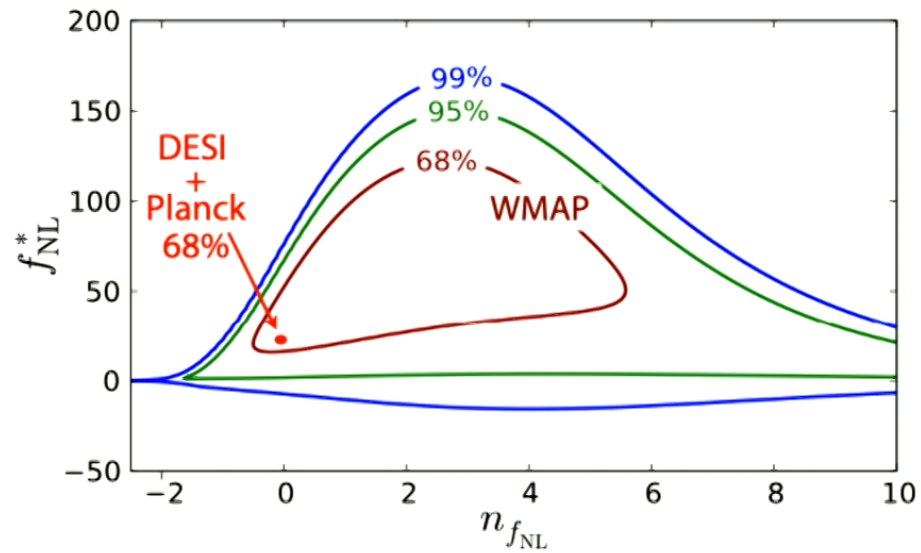
- ▶ Then they ran MCMC...
- ▶ Resulting parameter constraints (blue ellipse):



Why we often need MCMC

- ▶ Real-life models and likelihoods are often complex
- ▶ ... so the resulting **constraints** have complicated distributions (not Gaussians!)
- ▶ ... but we can represent them with **samplings**
- ▶ MCMC is used for drawing samples from probability distributions that we can compute numerically but cannot solve analytically

Samplings to represent constraints - examples



► From <https://arxiv.org/abs/1611.00036>

MCMC

draws samples from a probability function ↗

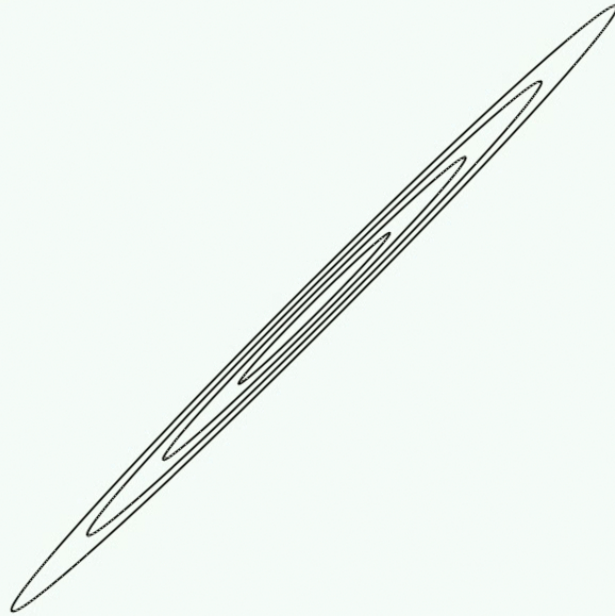
and all you need to be able to do is

evaluate

the function

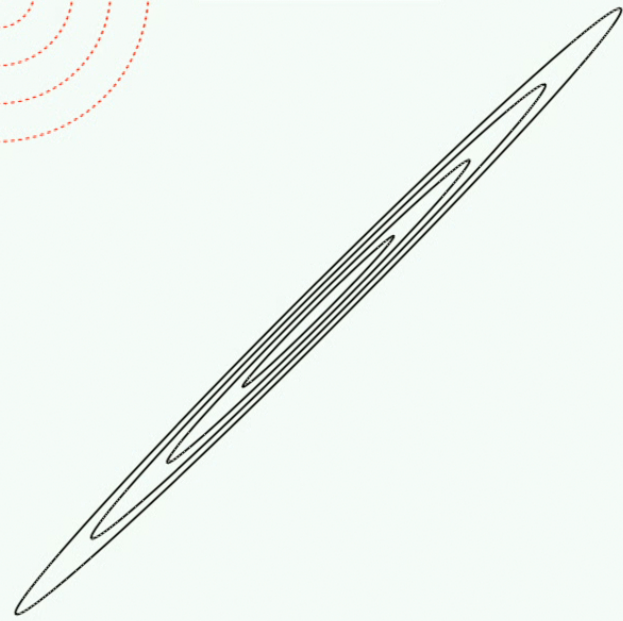
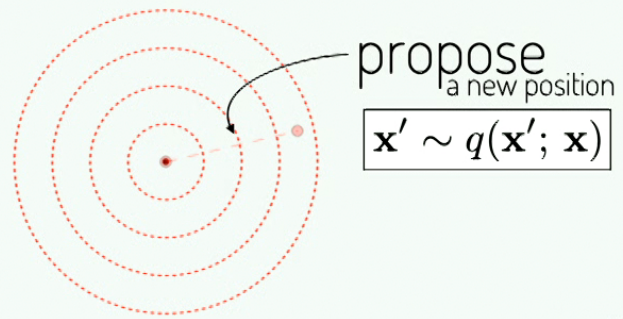
(up to a constant)

13

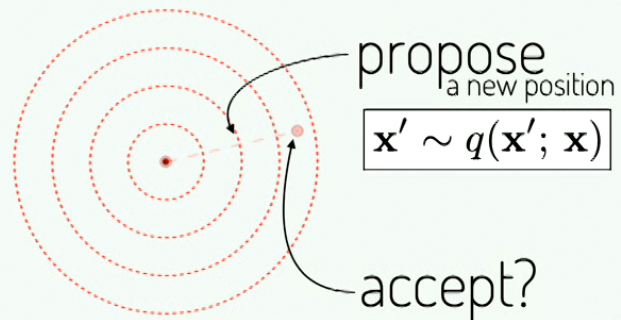


Metropolis-Hastings

14



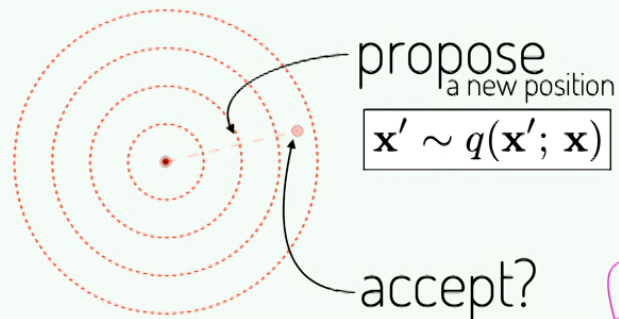
Metropolis–Hastings
in an ideal world



$$p(\text{accept}) = \min \left(1, \frac{p(\mathbf{x})}{p(\mathbf{x}')} \frac{q(\mathbf{x}; \mathbf{x}')}{q(\mathbf{x}'; \mathbf{x})} \right)$$

Metropolis-Hastings
in an ideal world

18



$$p(\text{accept}) = \min \left(1, \frac{p(\mathbf{x})}{p(\mathbf{x}')} \frac{q(\mathbf{x}; \mathbf{x}')}{q(\mathbf{x}'; \mathbf{x})} \right)$$

only relative probabilities

definitely.

Metropolis-Hastings

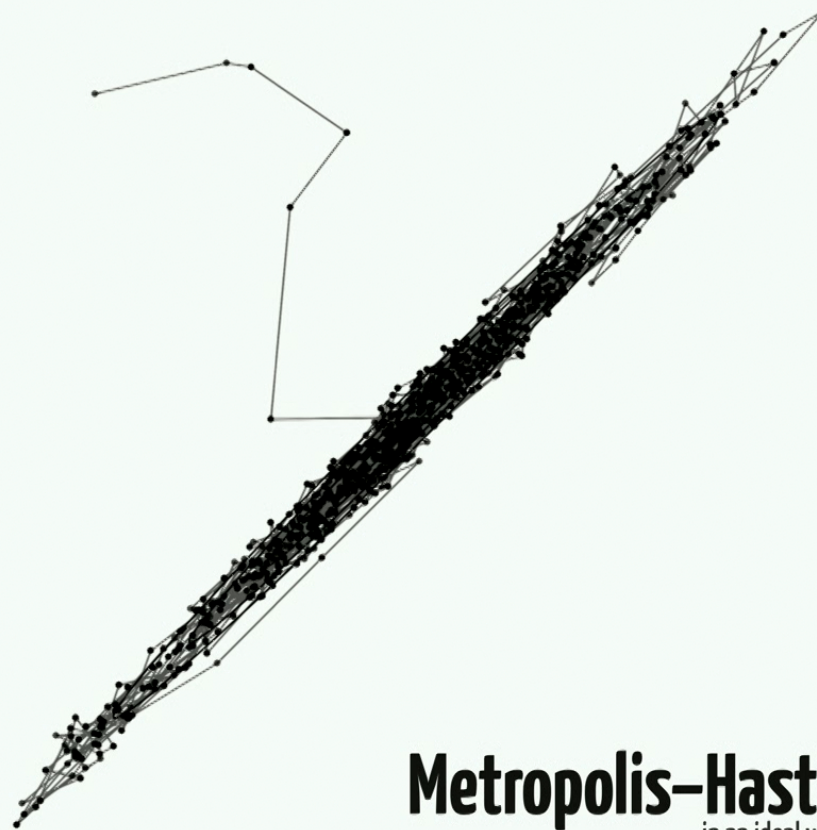
in an ideal world



double count!

Metropolis-Hastings
in an ideal world

26



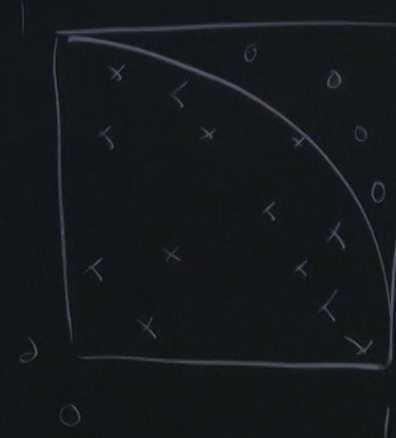
Metropolis-Hastings

in an ideal world

28

About the name

- ▶ **Monte Carlo**: a reference to the famous Monte Carlo Casino in Monaco, alluding to the randomness used in the algorithm
- ▶ **Markov Chain**: a list of samples, where each one is generated by a process that only looks at the previous one.
- ▶ **Markov**: a 19th-century Russian mathematician and impressive-moustache-haver with an [extensive list of things named after him](#)
- ▶ **Metropolis–Hastings**: lead authors of 1953 and 1970 papers (resp.) giving the algorithm with symmetric and general proposal distributions (resp.)



$$\frac{N(x)}{N(x) + N(0)} = \frac{\pi}{4}$$

The Algorithm (1)

```
function mcmc(prob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    prob = prob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        # ...
        # compute probability at new position
        # ...
        # decide whether to jump to the new position
        if # ...
            # ...
            # ...
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```

31

The Algorithm (2)

```
function mcmc(prob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    prob = prob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        prob_new = prob_func(p_new)
        # decide whether to jump to the new position
        if prob_new / prob > uniform_random()
            p = p_new
            prob = prob_new
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```

32

The Algorithm (3)

```
function mcmc(logprob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    logprob = logprob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        logprob_new = logprob_func(p_new)
        # decide whether to jump to the new position
        if exp(logprob_new - logprob) > uniform_random()
            p = p_new
            logprob = logprob_new
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```

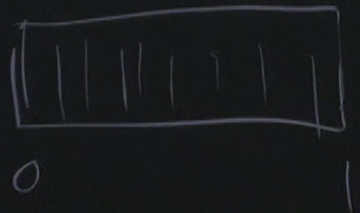
33

The Algorithm (4)

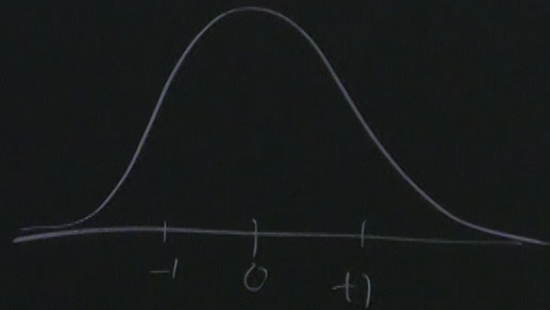
```
function mcmc(logprob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    logprob = logprob_func(p)
    chain = []
    naccept = 0
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        logprob_new = logprob_func(p_new)
        # decide whether to jump to the new position
        if exp(logprob_new - logprob) > uniform_random():
            p = p_new
            logprob = logprob_new
            naccept += 1
        end
        # save the position
        append!(chain, p)
    end
    return chain, naccept/nsteps
```

34

rand(1000)



randn(1000)

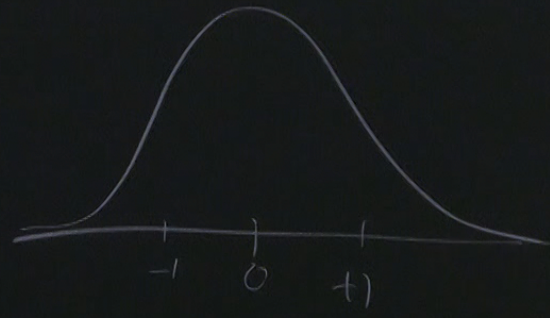


$$\overline{(0)} = \frac{\pi}{4}$$

rand(1000)



randn(1000)

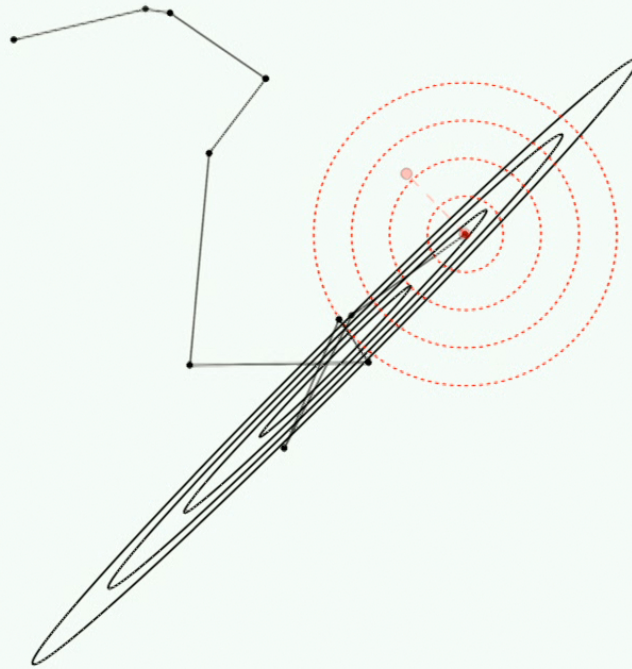


$$P(0) = \frac{\pi}{4}$$

$\text{prob}(p) = 20\%$
 p-new

Practicalities

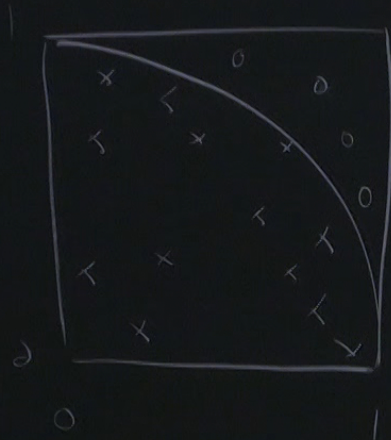
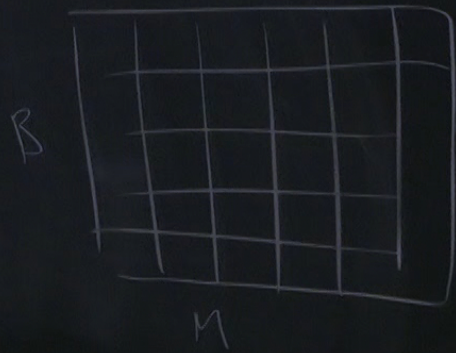
- ▶ How do I choose a proposal distribution?
- ▶ How many steps do I have to take?



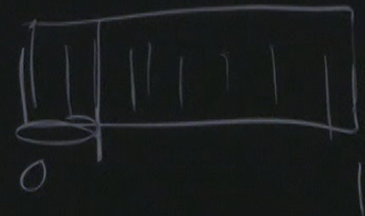
Metropolis-Hastings

in the real world

36



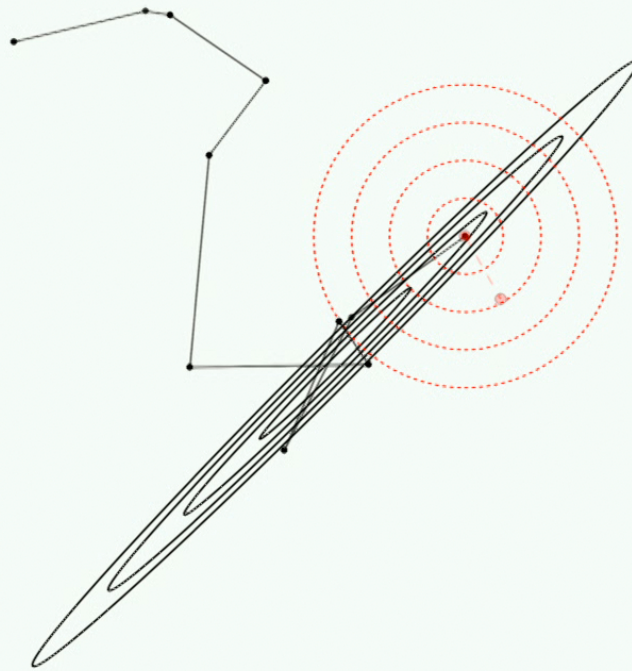
rand(1000)



$$\frac{N(x)}{N(x) + N(o)} = \frac{\pi}{4}$$

$$\frac{\text{prob}(p_{\text{new}} = 20\%)}{\text{prob}(p_{\text{old}})}$$

p-new



Metropolis-Hastings

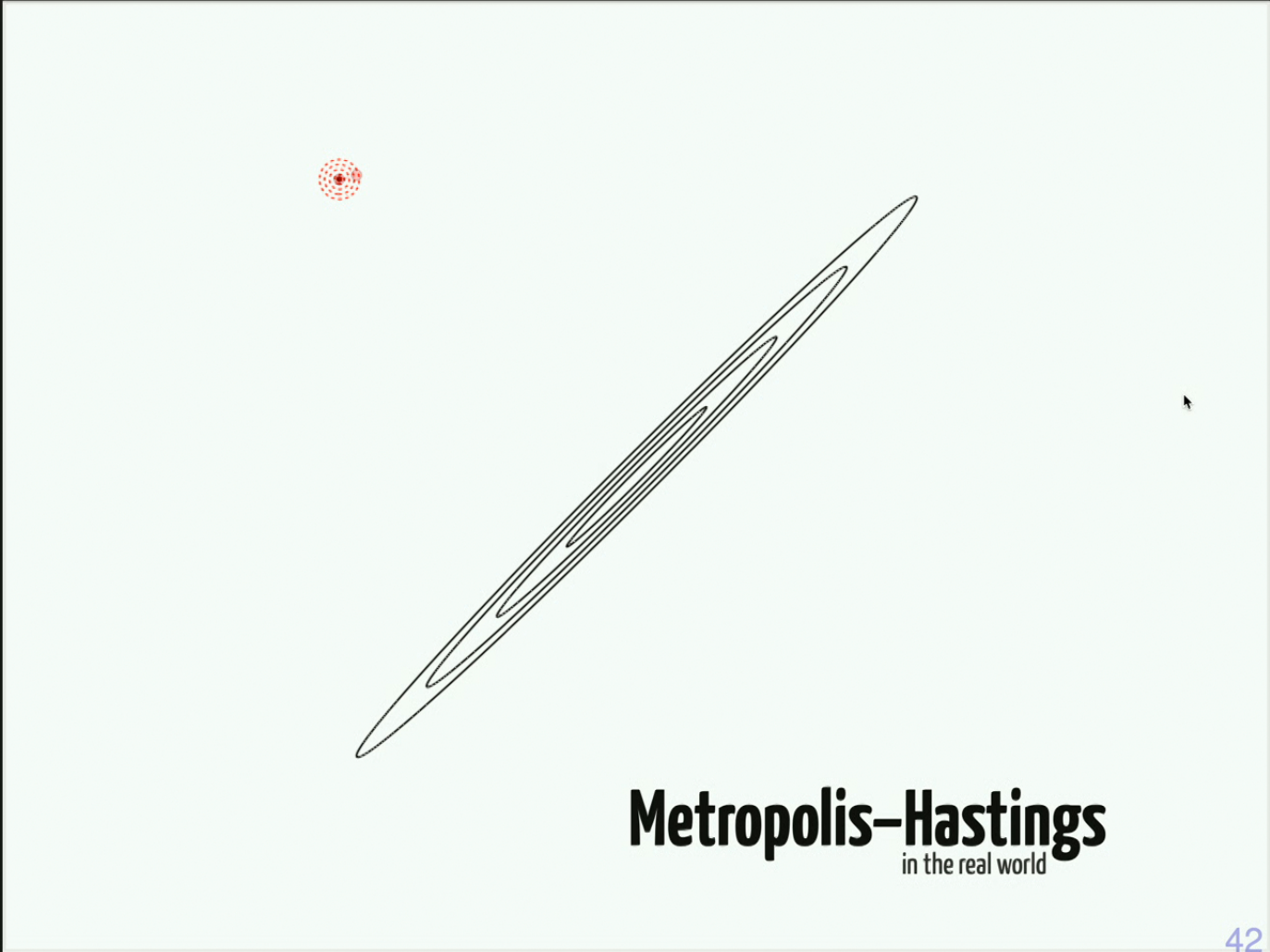
in the real world

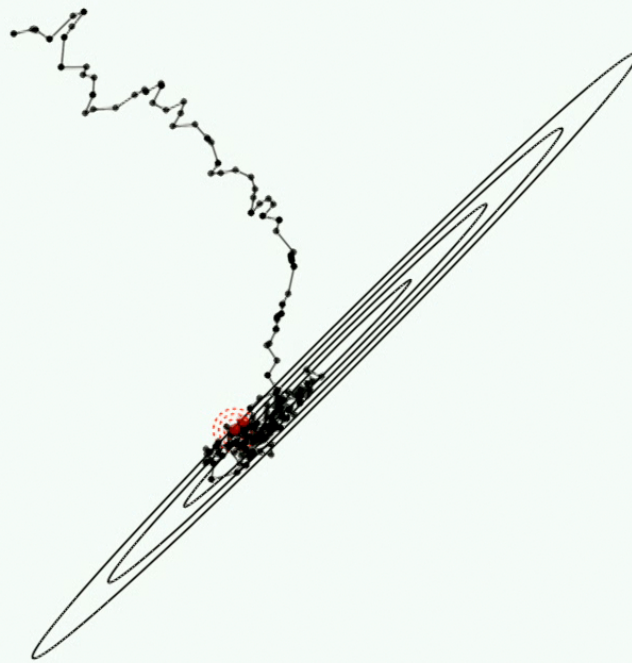
40



the
Small Acceptance Fraction
problem

Metropolis-Hastings
in the real world

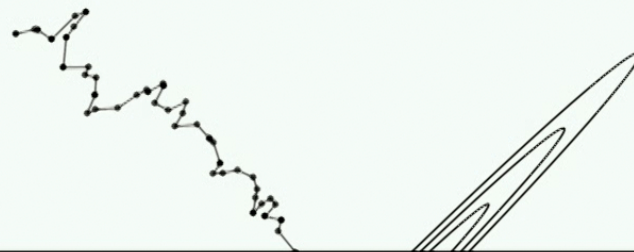




Metropolis-Hastings

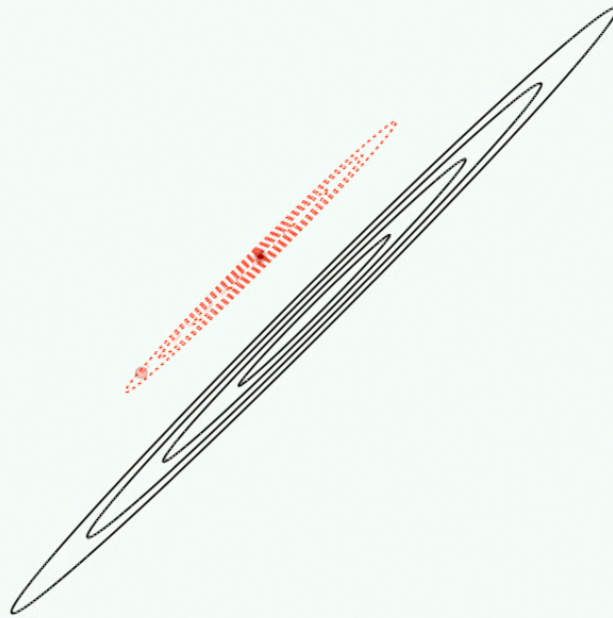
in the real world

43



the
Huge Acceptance Fraction
problem

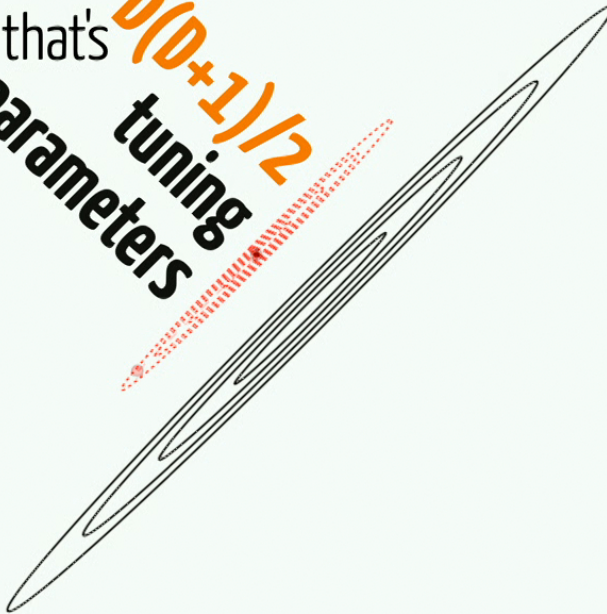
Metropolis–Hastings
in the real world



Metropolis-Hastings
in the real world

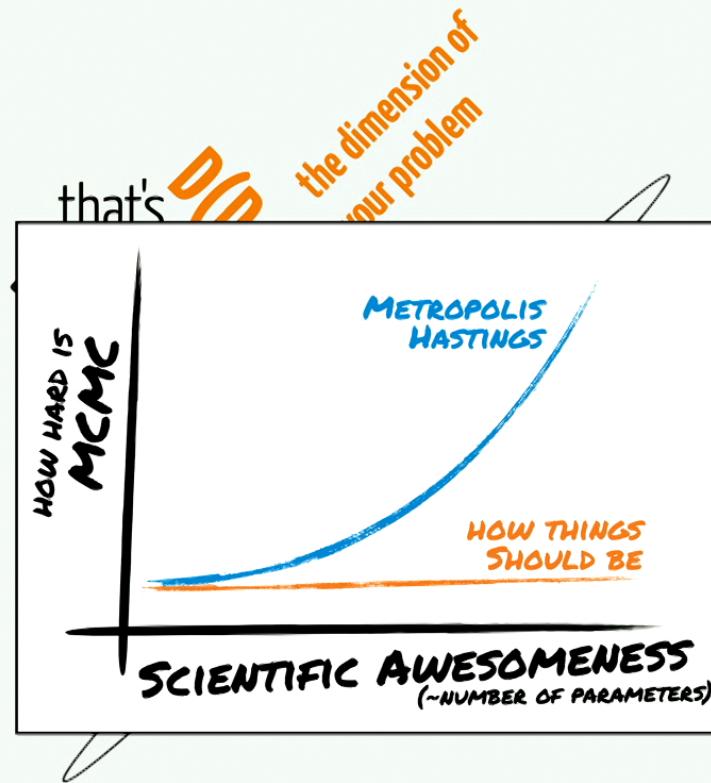
45

that's $O(D+1)/2$
tuning
parameters



Metropolis–Hastings
in the real world

46

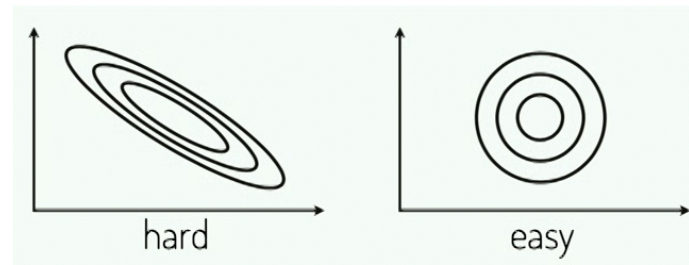


Metropolis-Hastings

in the real world

A connection to symmetries

- ▶ In Metropolis–Hastings MCMC, the *proposal distribution* needs **tuning parameters**, especially as dimensionality increases
- ▶ Can be seen as a lack of **symmetry** in the algorithm—the algorithm is sensitive to the parameterization of the problem
- ▶ For example, it's not invariant to an **affine** transformation
- ▶ **Next lecture**, I'll show you an alternative algorithm that **does** have affine invariance



How many samples do I need?

- ▶ Burn-in — skip the first N samples
- ▶ *Has my chain converged?*
- ▶ MCMC produces **correlated** samples, so
 - ▶ How correlated are my samples?
 - ▶ Can measure the *autocorrelation time* τ
 - ▶ Keep $1/\tau$ of the MCMC samples
 - ▶ eg <https://github.com/dfm/acor>
 - ▶ How many uncorrelated samples do I need?
 - ▶ No easy general answer to this question!
 - ▶ “How many can you afford?”

Conclusions

- ▶ MCMC remains an essential tool for probabilistic inference
- ▶ For science: lets us constrain model parameters based on data (Bayesian inference)
- ▶ Beguilingly simple algorithm, but difficult practicalities
- ▶ MCMC has beautiful theoretical guarantees... as compute time $\rightarrow \infty$