

Title: Latent variable justifies the stronger instrumental variable bounds

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Abstract: For binary instrumental variable models, there seems to be a long-standing gap between two sets of bounds on the average treatment effect: the stronger Balke-Pearl ("sharp") bounds versus the weaker Robins-Manski ("natural") bounds. In the literature, the Balke-Pearl bounds are typically derived under stronger assumptions, i.e., either individual exclusion or joint exogeneity, which are untestable cross-world statements, while the natural bounds only require testable assumptions. In this talk, I show that the stronger bounds are justified by the existence of a latent confounder. In fact, the Balke-Pearl bounds are sharp under latent confounding and stochastic exclusion. The "secret sauce" that closes this gap is a set of CHSH-type inequalities that generalize Bell's (1964) inequality.

Latent Variable Justifies the Stronger IV Bounds

F. Richard Guo

April 19, 2023

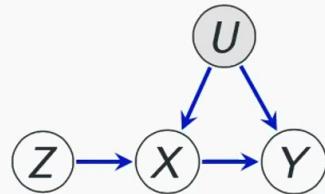
Causal Inference & Quantum Foundations Workshop
Perimeter Institute

Statistical Laboratory, Cambridge
based on joint work w. Thomas Richardson

Question: *Why should statisticians care about CHSH/Bell-type inequalities?*

- ☞ This comes up in studying partial identification when there could be latent variables.

Instrumental variable



- ▶ Suppose $X, Y, Z \in \{0, 1\}$:

Z : instrument, X : treatment, Y : outcome.

$Y(x, z)$ denotes the counterfactual outcome.

- ▶ U is a latent confounder. ↗ No assumption on its state space.
- ↗ We want **sharp bounds** on the average treatment effect

$$\tau \equiv \mathbb{E} Y(x = 1, z = 0) - \mathbb{E} Y(x = 0, z = 0)$$

in terms of $P(X, Y, Z)$.

Balke–Pearl bounds

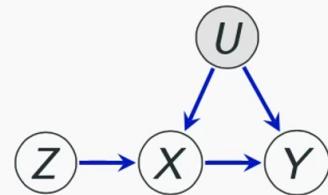
Balke and Pearl (1997) obtained bounds on τ :

$$\max \left\{ \begin{array}{l} -P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 1) - 1 \\ P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 0) - 1 \\ P(X = 1, Y = 1 | Z = 0) - P(Y = 1 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ P(X = 1, Y = 1 | Z = 1) - P(Y = 1 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 1) - P(Y = 0 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 0) - P(Y = 0 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \end{array} \right\}$$

$$\leq \tau \leq$$

$$\min \left\{ \begin{array}{l} -P(X = 1, Y = 1 | Z = 0) - P(X = 0, Y = 0 | Z = 0) \\ -P(X = 1, Y = 1 | Z = 1) - P(X = 0, Y = 0 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 0) + P(X = 0, Y = 1 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 1) + P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 0) + P(Y = 0 | Z = 1) + P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) + P(Y = 0 | Z = 0) + P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 1) \\ -P(X = 0, Y = 1 | Z = 1) + P(Y = 1 | Z = 0) + P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 1) \\ -P(X = 0, Y = 1 | Z = 0) + P(Y = 1 | Z = 1) + P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 0) \end{array} \right\}$$

Balke–Pearl bounds



The Balke–Pearl bounds are **sharp** under the following two **cross-world** assumptions.

- ▶ Individual-level exclusion (**ind.Excl**)

$$Y(x, z = 0) = Y(x, z = 1), \quad x = 0, 1.$$

(so one can write $Y(x) = Y(x, z)$) ↗ Cannot be consistently tested by randomizing (Z, X) .

- ▶ Joint exogeneity (**joint.Exo**)

$$Z \perp\!\!\!\perp Y(0, 0), Y(0, 1), Y(1, 0), Y(1, 1)$$

(i.e., $Z \perp\!\!\!\perp Y(0), Y(1)$ under **(ind.Excl)**)

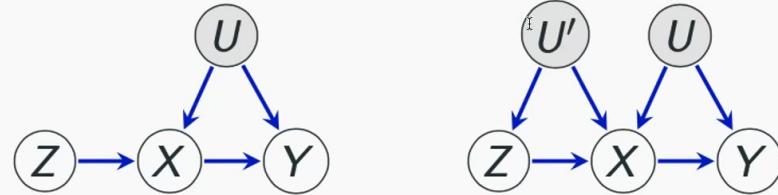
↗ Dubious in observational studies.

Question: *What bounds can we obtain with single-world assumptions?*

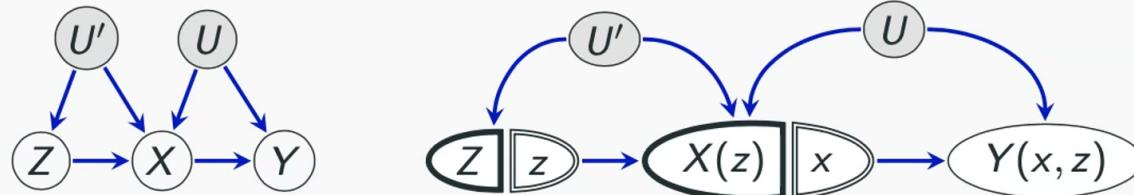
Latent variable DAG

Following the convention in the literature,

- we introduce counterfactual $X(z)$,
 - and allow a separate latent variable U' between Z and X for $U' \perp\!\!\!\perp U$.
- ☞ But these are not essential for our derivation.



SWIG interpretation of the DAG



(Population-level) Single World Intervention Graph (SWIG) (Richardson and Robins, 2013) encodes the following assumptions.

1. $Y(x, z) \perp\!\!\!\perp z \mid U$ on the SWIG encodes the **latent-variable stochastic exclusion**

$$\mathbb{E}[Y(x, 0) \mid U] = \mathbb{E}[Y(x, 1) \mid U], \quad x \in \{0, 1\} \quad (\text{LV.sto.Excl})$$

☞ Compare with **(ind.Excl)**: $Y(x, 0) = Y(x, 1)$.

2. **Latent-variable marginal exogeneity**

$$Z \perp\!\!\!\perp U, \quad Y(x, z) \perp\!\!\!\perp Z, X(z) \mid U \quad (\text{LV.marg.Exo})$$

☞ Compare with **(joint.Exo)**: $Z \perp\!\!\!\perp Y(0,0), Y(0,1), Y(1,0), Y(1,1)$.

Deriving bounds

We derive the bounds on $\tau \equiv \mathbb{E} Y(1,0) - \mathbb{E} Y(0,0)$ with **linear programming**.

1. Let $\{P(X(0), X(1), Y(0,0), Y(0,1), Y(1,0), Y(1,1) \mid Z = z) : z = 0, 1\}$ and $\{P(X, Y \mid Z = z) : z = 0, 1\}$ be variables. U, U' are not part of the program.
2. Add constraints:
 - (1) Simplex
 - (2) Consistency: $P(X = x, Y = y \mid Z = z) = P(X(z) = x, Y(x,z) = y \mid Z = z)$.
 - (3) Single-world assumptions

$$\mathbb{E}[Y(x,0) \mid U] = \mathbb{E}[Y(x,1) \mid U], \quad x \in \{0,1\} \quad (\text{LV.sto.Excl})$$

$$Z \perp\!\!\!\perp U, \quad Y(x,z) \perp\!\!\!\perp Z, X(z) \mid U \quad (\text{LV.marg.Exo})$$

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 - (3) Single-world assumptions

$$\boxed{\mathbb{E}[Y(x, 0) \mid U] = \mathbb{E}[Y(x, 1) \mid U]} \quad (\text{LV.sto.Excl}) \Rightarrow \boxed{\mathbb{E} Y(x, 0) = \mathbb{E} Y(x, 1)} \quad (\text{sto.Excl})$$

$$\boxed{Z \perp\!\!\!\perp U, \quad Y(x, z) \perp\!\!\!\perp Z, X(z) \mid U} \quad (\text{LV.marg.Exo}) \Rightarrow \boxed{Z \perp\!\!\!\perp Y(x, z)} \quad (\text{marg.Exo})$$

Deriving bounds

We derive the bounds on $\tau \equiv \mathbb{E} Y(1,0) - \mathbb{E} Y(0,0)$ with **linear programming**.

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 - (3) Single-world assumptions

$$\mathbb{E} Y(x,0) = \mathbb{E} Y(x,1), \quad x \in \{0,1\} \quad (\text{sto.Excl})$$

$$Z \perp\!\!\!\perp Y(x,z), \quad x, z \in \{0,1\} \quad (\text{marg.Exo})$$

3. Also add τ as a programming variable. \Leftrightarrow Degenerate convex polytope.
4. Project onto coordinates $\{P(X, Y \mid Z = z) : z = 0, 1\}$ and τ .

Robins–Manski bounds

We get the “natural” bounds (Manski, 1990; Robins, 1989):

$$\max \left\{ \begin{array}{l} -P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 1) - 1 \\ P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 0) - 1 \end{array} \right\}$$
$$\leq \tau \leq$$
$$\min \left\{ \begin{array}{l} -P(X = 1, Y = 1 | Z = 0) - P(X = 0, Y = 0 | Z = 0) \\ -P(X = 1, Y = 1 | Z = 1) - P(X = 0, Y = 0 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 0) + P(X = 0, Y = 1 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 1) + P(X = 0, Y = 1 | Z = 0) \end{array} \right\}$$

Compare with the Balke–Pearl bounds

$$\max \left\{ \begin{array}{c} -P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 1) - 1 \\ P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 0) - 1 \\ P(X = 1, Y = 1 | Z = 0) - P(Y = 1 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ P(X = 1, Y = 1 | Z = 1) - P(Y = 1 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 1) - P(Y = 0 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 0) - P(Y = 0 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \end{array} \right\}$$

$$\leq \tau \leq$$

$$\min \left\{ \begin{array}{c} -P(X = 1, Y = 1 | Z = 0) - P(X = 0, Y = 0 | Z = 0) \\ -P(X = 1, Y = 1 | Z = 1) - P(X = 0, Y = 0 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 0) + P(X = 0, Y = 1 | Z = 1) \\ 1 - P(X = 1, Y = 0 | Z = 1) + P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 0) + P(Y = 0 | Z = 1) + P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) + P(Y = 0 | Z = 0) + P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 1) \\ -P(X = 0, Y = 1 | Z = 1) + P(Y = 1 | Z = 0) + P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 1) \\ -P(X = 0, Y = 1 | Z = 0) + P(Y = 1 | Z = 1) + P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 0) \end{array} \right\}$$

☞ **Question:** Is this gap fundamental to the single-world assumptions?

What caused the gap?

In terms of models, we have

$$(\text{ind.Excl}) \subset (\text{LV.sto.Excl}) \subset (\text{sto.Excl})$$

and

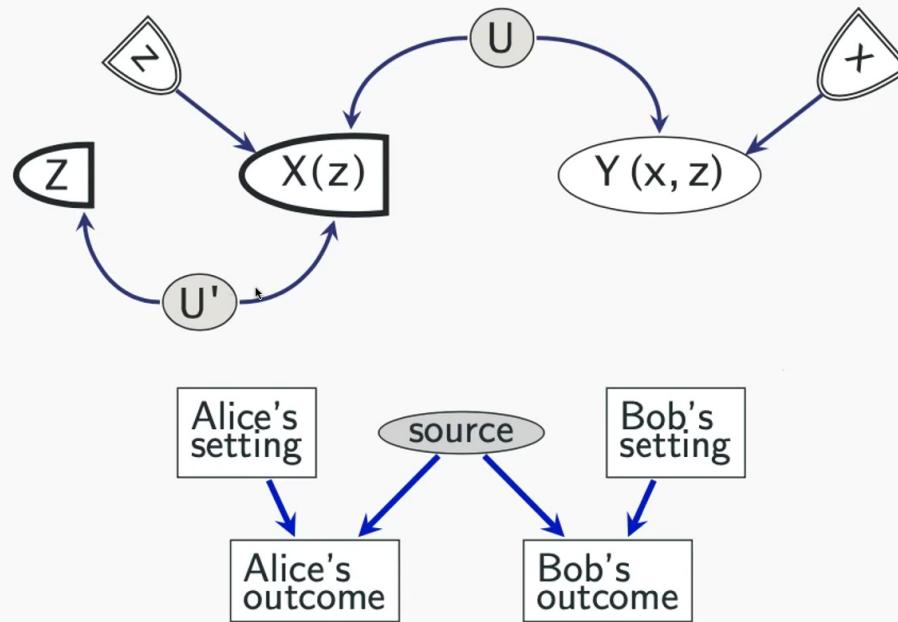
$$(\text{joint.Exo}) \subset (\text{LV.marg.Exo}) \subset (\text{marg.Exo}).$$

☞ The gap stems from

$(\text{LV.sto.Excl}) \Rightarrow (\text{sto.Excl})$ and $(\text{LV.marg.Exo}) \Rightarrow (\text{marg.Exo})$ by integrating out U ,

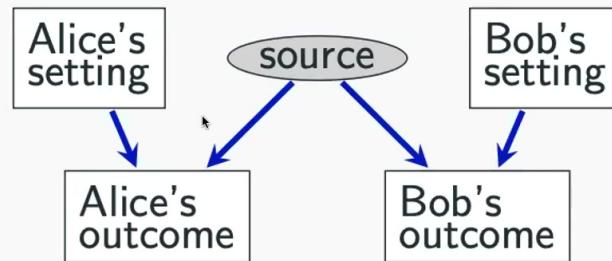
which misses the CHSH inequality implied by the sheer **existence** of U !

The missing piece



☞ Works similarly if $X(z)$ is not introduced.

CHSH



CHSH is postulated on the correlations:

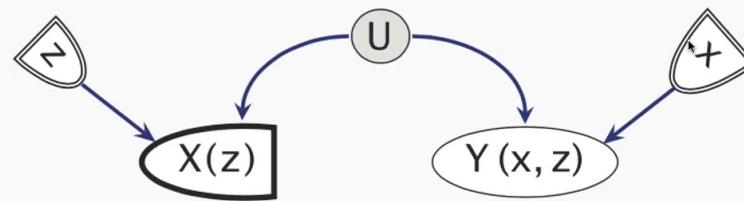
$$\langle A, B \rangle + \langle A', B \rangle + \langle A, B' \rangle - \langle A', B' \rangle \leq 2.$$

Alice either observes $A = \pm 1$ or $A' = \pm 1$; Bob observes either $B = \pm 1$ or $B' = \pm 1$.

CHSH

Lemma The latent variable U implies the Bell–CHSH inequalities

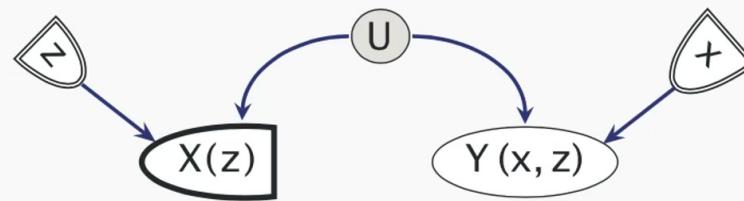
$$\begin{aligned} 0 \leq & P(X = 1, Y(x, z) = 1 \mid Z = z) + P(X = 0, Y(1 - x, z) = 0 \mid Z = z) \\ & + P(X = 0, Y(x, 1 - z) = 0 \mid Z = 1 - z) \\ & - P(X = 0, Y(1 - x, 1 - z) = 0 \mid Z = 1 - z) \leq 1, \quad x, z \in \{0, 1\}. \end{aligned}$$



CHSH

Lemma The latent variable U implies the Bell–CHSH inequalities

$$\begin{aligned} 0 \leq & P(X = 1, Y(x, z) = 1 \mid Z = z) + P(X = 0, Y(1 - x, z) = 0 \mid Z = z) \\ & + P(X = 0, Y(x, 1 - z) = 0 \mid Z = 1 - z) \\ & - P(X = 0, Y(1 - x, 1 - z) = 0 \mid Z = 1 - z) \leq 1, \quad x, z \in \{0, 1\}. \end{aligned}$$



☞ Affine in $\{P(X(0), X(1), Y(0, 0), Y(0, 1), Y(1, 0), Y(1, 1) \mid Z = z) : z = 0, 1\}$.

Closing the gap

1. Let $\{P(X(0), X(1), Y(0,0), Y(0,1), Y(1,0), Y(1,1) \mid Z = z) : z = 0, 1\}$ and $\{P(X, Y \mid Z = z) : z = 0, 1\}$ be variables.
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$$\mathbb{E} Y(x, 0) = \mathbb{E} Y(x, 1), \quad x \in \{0, 1\} \quad (\text{sto.Excl})$$

$$Z \perp\!\!\!\perp Y(x, z), \quad x, z \in \{0, 1\} \quad (\text{marg.Exo})$$

- (4) **CHSH inequalities**

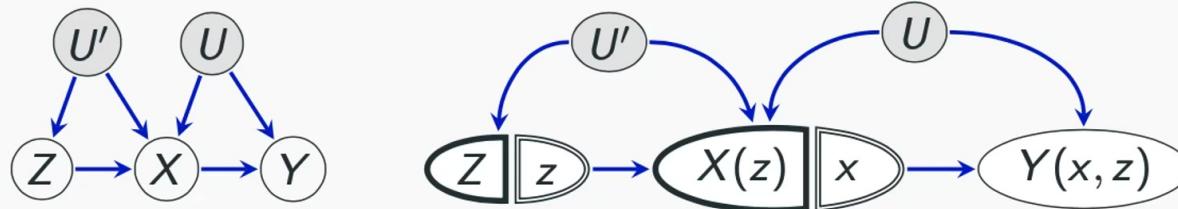
We recover the Balke–Pearl bounds!

$$\max \left\{ \begin{array}{l} -P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ -P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 1, Y = 1 | Z = 0) + P(X = 0, Y = 0 | Z = 1) - 1 \\ P(X = 1, Y = 1 | Z = 1) + P(X = 0, Y = 0 | Z = 0) - 1 \\ P(X = 1, Y = 1 | Z = 0) - P(Y = 1 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \\ P(X = 1, Y = 1 | Z = 1) - P(Y = 1 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 1) - P(Y = 0 | Z = 0) - P(X = 1, Y = 0 | Z = 1) - P(X = 0, Y = 1 | Z = 1) \\ P(X = 0, Y = 0 | Z = 0) - P(Y = 0 | Z = 1) - P(X = 1, Y = 0 | Z = 0) - P(X = 0, Y = 1 | Z = 0) \end{array} \right\}$$

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Result



Theorem Suppose $X, Y, Z \in \{0, 1\}$. Under **(LV.sto.Excl)**+**(LV.marg.Exo)**, which are encoded by SWIG for the latent variable DAG, Balke–Pearl bounds hold and are sharp for τ .

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Further interpretation...

Imagine you are a statistician (or econometrician) who is only willing to conceptualize $\{Z, X(z), Y(z, x) : x, z = 0, 1\}$.

Further, you only want to put down single-world assumptions

$$\mathbb{E} Y(x, 0) = \mathbb{E} Y(x, 1), \quad x \in \{0, 1\} \quad (\text{sto.Excl})$$

$$Z \perp\!\!\!\perp Y(x, z), \quad x, z \in \{0, 1\} \quad (\text{marg.Exo})$$

Question: What is the minimal extra condition on $\{Z, X(z), Y(z, x) : x, z = 0, 1\}$ that is necessary to yield the Balke–Pearl bounds?

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Answer: CHSH!

Summary

- Balke–Pearl bounds need not rely on untestable, cross-world assumptions such as **(ind.Excl)** and **(joint.Exo)**.
 - They are justified by a SWIG with latent variable.
 - That being said, CHSH is cross-world.
 - ☞ But it is a consequence of the existence of U , rather than assumptions that we impose on!
- The CHSH inequalities fill the gap between Balke–Pearl and Robins–Manski.
 - ☞ Dawid (2003) showed the Balke–Pearl bounds (and their tightness) with a certain latent variable formulation that does not use counterfactuals.
- To obtain sharp bounds, we need to understand the inequality constraints in a latent variable formulation.
 - ☞ Good to consult a graph over both observed and counterfactual variables (e.g., SWIGs).

Summary

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- To obtain sharp bounds, we need to understand the inequality constraints in a latent variable formulation.
 - ☞ Good to consult a graph over both observed and counterfactual variables (e.g., SWIGs).
- Linear/convex programming is a useful tool.

^I ► Ref: F. R. Guo, Likelihood Analysis of Causal Models, §5.3. University of Washington, 2021. 20

THANKS

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