

Title: Bounding counterfactual distributions in discrete structural causal models

Speakers: Jin Tian

Collection: Causal Inference & Quantum Foundations Workshop

Date: April 19, 2023 - 3:30 PM

URL: <https://pirsa.org/23040118>

Abstract: We investigate the problem of bounding counterfactual queries from an arbitrary collection of observational and experimental distributions and qualitative knowledge about the underlying data-generating model represented in the form of a causal diagram. We show that all counterfactual distributions in an arbitrary structural causal model (SCM) with finite discrete endogenous variables could be generated by a family of SCMs with the same causal diagram where unobserved (exogenous) variables are discrete with a finite domain. Utilizing this family of SCMs, we translate the problem of bounding counterfactuals into that of polynomial programming whose solution provides optimal bounds for the counterfactual query.

Counterfactuals

- Goal: inferring counterfactual queries from
 - ◆ observational/experimental data
 - ◆ causal diagram: qualitative knowledge about the underlying data-generating model
- E.g., investigating the gender discrimination in admission:
“Would admission outcome for female applicant change had she been a male?”

Counterfactuals

- Goal: inferring counterfactual queries from
 - ◆ observational/experimental data
 - ◆ causal diagram: qualitative knowledge about the underlying data-generating model
- E.g., investigating the gender discrimination in admission:
“Would admission outcome for female applicant change had she been a male?”

Partial Counterfactual Identifica.

- Identifying counterfactual distributions from data and causal diagram
 - ◆ (Halpern, 1998, Shpitser and Pearl 2007, Correa et al. 2021)
- Often non-identifiable
- Partial identification: deriving informative bounds (Manski, 1990; Robins, 1989; Balke & Pearl, 1994; 1997; Tian & Pearl, 2000; Evans, 2012; Richardson et al., 2014; Zhang & Bareinboim, 2017; Kallus & Zhou, 2018; Finkelstein & Shpitser, 2020; Kilbertus et al., 2020; Zhang & Bareinboim, 2021; ...)

Structural Causal Models

A structural causal model (SCM) is a tuple $\langle V, U, \mathcal{F}, P \rangle$ where

- V is a set of endogenous variables
- U is a set of mutually independent exogenous variables
- \mathcal{F} is a set of functions where each $f_V \in \mathcal{F}$ decides values of an endogenous variable $V \in V$

$$v \leftarrow f_V(pa_V, u_V), PA_V \subseteq V, U_V \subseteq U$$

- $P(U)$ is a distribution

We consider acyclic SCMs.

Pearl's Causal Hierarchy - L1&2

- \mathcal{F} can be seen as a mapping from $U \rightarrow V$
- An SCM M induces distribution $P(V)$, called the *observational distribution*
- An intervention $\text{do}(X = x)$ induces a submodel M_x

$x \leftarrow f_X(\text{pa}_X, u_X)$ replaced by $x \leftarrow x$

- M_x induces an *interventional distribution* $P(V|\text{do}(x))$
- Causal effect $P(y|\text{do}(x))$: how the outcome Y will respond if we take an action $X = x$

Pearl's Causal Hierarchy - L3

- The *potential response* $Y_x(\mathbf{u})$ is defined as the solution of Y in the submodel M_x given $U = \mathbf{u}$.
- $P(U)$ induces a *counterfactual variable* Y_x

$$P(Y_x = y) = P(\mathbf{y}|\text{do}(\mathbf{x})) = \int_{\Omega_U} \mathbb{1}_{Y_x(\mathbf{u})=y} dP(\mathbf{u})$$

- A *counterfactual distribution*

$$P(\mathbf{y}_x, \dots, \mathbf{z}_w) = \int_{\Omega_U} \mathbb{1}_{Y_x(\mathbf{u})=y, \dots, Z_w(\mathbf{u})=z} dP(\mathbf{u})$$

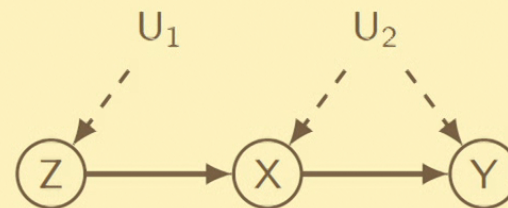
Causal Diagram

- Every SCM M induces a causal diagram

$$z \leftarrow f_Z(u_1)$$

$$x \leftarrow f_X(z, u_2)$$

$$y \leftarrow f_Y(x, u_2)$$



- Researchers may know the scope of the functions, but not the details about the underlying mechanisms
- A graph is compatible with infinitely many SCMs

Causal Inference Tasks

- Model knowledge: causal diagram
- Data: $P(v)$, $P(v|do(z))$
- Query: $Q = P(y|do(x))$; $Q = P(y_x, z_w)$
- Develop CI algorithms:
 - ◆ Identifiable? $P(y|do(x)) = \sum_z P(y|x, z)P(z)$
 - ◆ Partially identifiable? $Q \in [a, b]$

CI by Optimization

- Task: given the observational distribution $P(v)$ in an arbitrary causal diagram \mathcal{G} , bound $P(y|do(x))$ or $P(Y_x, Z_w)$
- Assume endogenous variables \mathbf{V} are discrete and finite
- let $\mathcal{M}(\mathcal{G})$ be the set of all SCMs associated with \mathcal{G}

$$\begin{aligned} \min / \max_{M \in \mathcal{M}(\mathcal{G})} & P_M(\mathbf{y}_x, \dots, \mathbf{z}_w) \\ \text{s.t.} & P_M(\mathbf{V}) = P(\mathbf{V}) \end{aligned}$$

- Solving this optimization is difficult since we do not have access to the parametric forms of structural functions f_V nor $P(\mathbf{U})$

Canonical Partitioning (Balke&Pearl 94)

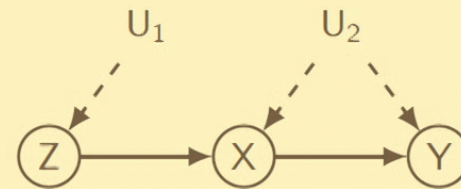
- IV model: binary X (treatment realized), Y (outcome), Z (treatment assigned)
- The *canonical functions* $x \leftarrow h_X^{(i)}(z)$

$$h_X^{(1)}(z) = 0, \quad \text{Never-taker}$$

$$h_X^{(2)}(z) = z, \quad \text{Complier}$$

$$h_X^{(3)}(z) = 1 - z, \quad \text{Defier}$$

$$h_X^{(4)}(z) = 1. \quad \text{Always-taker}$$



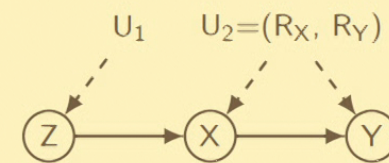
- For any $x \leftarrow f_X(z, u_2)$, there exists a *canonical partition* $\mathcal{U}_X^{(i)}$, $i = 1, 2, 3, 4$ over the domain of U_2 such that $u_2 \in \mathcal{U}_X^{(i)}$ if and only if $f_X(\cdot, u_2) = h_X^{(i)}$

Canonical IV Model (Balke&Pearl 94)

- A canonical IV model with $X, Y, Z \in \{0, 1\}$:
 $U_2 = (R_X, R_Y)$ where $R_X, R_Y \in \{1, 2, 3, 4\}$

$$x \leftarrow f_X(z, R_X) = h_X^{(R_X)}(z),$$

$$y \leftarrow f_Y(x, R_Y) = h_Y^{(R_Y)}(x).$$



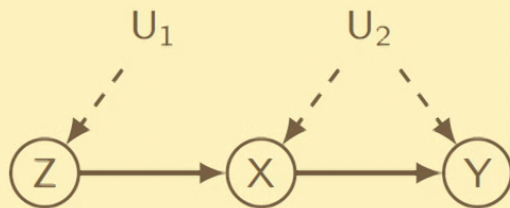
- For any IV model M there exists a canonical IV model N such that

$$P_M(x, y, z) = P_N(x, y, z), \quad P_M(y|do(x)) = P_N(y|do(x))$$

$$P_M(Z, X_{z_0}, X_{z_1}, Y_{x_0}, Y_{x_1}) = P_N(Z, X_{z_0}, X_{z_1}, Y_{x_0}, Y_{x_1})$$

CI by Optimization (Balke&Pearl 94)

- Task: given the observational distribution $P(x, y, z)$, bound the causal effect $P(y|do(x))$



- Parameters in a canonical IV model: $P(R_X, R_Y)$
- $P(x, y, z)$ imposes constraints on $P(R_X, R_Y)$
- Optimize $P(y|do(x))$ expressed in terms of $P(R_X, R_Y)$

LP Formulation (Balke&Pearl 94)

- 16 parameters

$$q_{jk} = P(R_X = j, R_Y = k)$$

- Express $p_{i.j.k} = P(X = i, Y = j|Z = k)$ as linear functions of q_{jk}

$$p_{00.0} = P(Y = 0, X = 0|Z = 0) = q_{00} + q_{01} + q_{10} + q_{11}$$

- Express objective function $P(y|do(x))$ as linear functions of q_{jk}

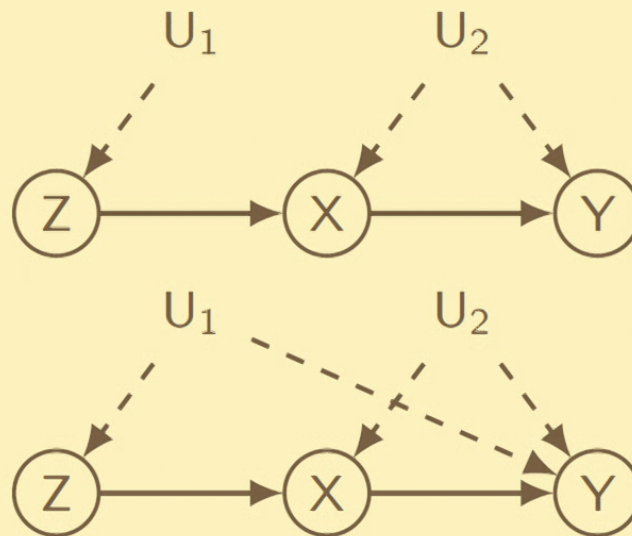
$$\begin{aligned} &P(Y = 1|do(X = 0)) \\ &= q_{02} + q_{12} + q_{22} + q_{32} + q_{03} + q_{13} + q_{23} + q_{33} \end{aligned}$$

Beyond IV Model

- Goal: extend the canonical IV model to general causal diagrams
- Assume endogenous variables V are discrete and finite
- Evans et al. (2018) showed observational distributions in geared graphs could be generated by a model with exogenous variables of discrete domains
- Rosset et al. (2018), Fraser (2020) showed observational distribution in an arbitrary causal diagram can be generated by a model with finite-state exogenous variables

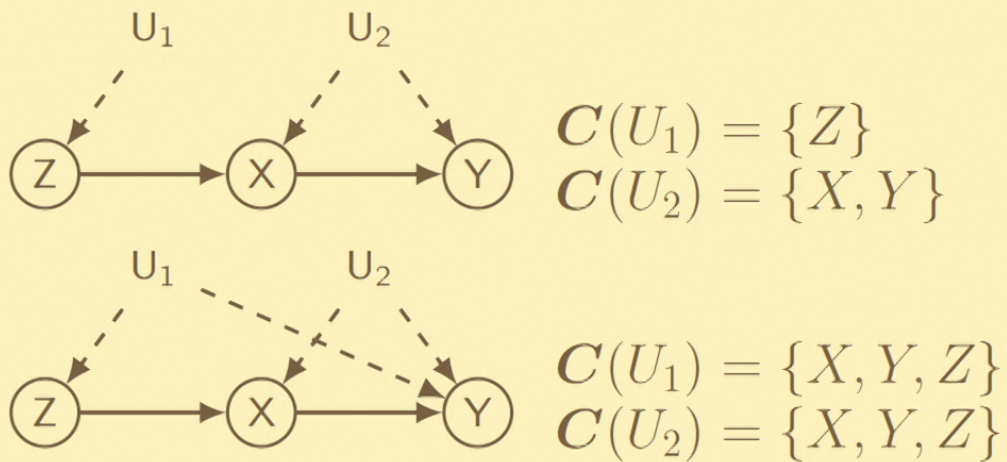
C-Components

Definition (C-component): Two endogenous variables are in the same c-component if and only if they are connected by a bidirected path, a path composed entirely of bi-directed edges.



C-Components

- We denote by $C(U)$ the maximal c-component covering U in \mathcal{G} , i.e., $U \in \bigcup_{V \in C(U)} U_V$.



Canonical SCMs

Definition: An SCM $M = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P \rangle$ is said to be a canonical SCM over discrete endogenous $V \in \mathbf{V}$ with finite domain Ω_V if

- Every exogenous $U \in \mathbf{U}$ has a finite domain Ω_U with cardinality

$$|\Omega_U| = \prod_{V \in \mathcal{C}(U)} |\Omega_{PA_V \mapsto \Omega_V}|,$$

$|\Omega_{PA_V \mapsto \Omega_V}| = |\Omega_V|^{|\Omega_{PA_V}|}$ is the number of possible (canonical) functions mapping domain Ω_{PA_V} to Ω_V .

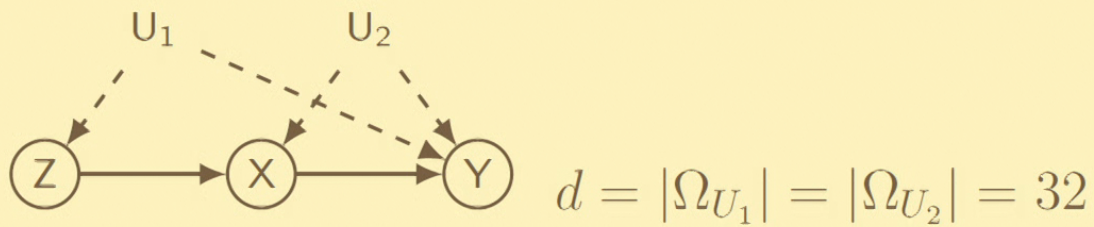
$$v \leftarrow f_V(pa_V, u_V)$$

Canonical SCMs

Theorem For an arbitrary SCM $M = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P \rangle$ over discrete and finite endogenous variables \mathbf{V} , there exists a canonical SCM N such that

1. M and N are associated with the same causal diagram, i.e., $\mathcal{G}_M = \mathcal{G}_N$.
2. For any set of counterfactual variables $\mathbf{Y}_x, \dots, \mathbf{Z}_w$,
 $P_M(\mathbf{Y}_x, \dots, \mathbf{Z}_w) = P_N(\mathbf{Y}_x, \dots, \mathbf{Z}_w)$.

Example



$$P(x_{z'}, y_{x'}) = \sum_{u_1, u_2=1}^d \mathbb{1}_{f_X(z', u_2)=x} \mathbb{1}_{f_Y(x', u_1, u_2)=y} P(u_1) P(u_2).$$

$$P(\mathbf{y}_x, \dots, \mathbf{z}_w) = \sum_{U \in \mathcal{U}: u=1}^{d_U} \mathbb{1}_{Y_x(u)=y, \dots, Z_w(u)=z} \prod_{U \in \mathcal{U}} P(u).$$

Canonical SCMs - Int. Dist.

Theorem For an arbitrary SCM $M = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P \rangle$ over discrete and finite endogenous variables \mathbf{V} , there exists a SCM N over \mathbf{V} with discrete exogenous \mathbf{U} having cardinality

$$|\Omega_U| = \prod_{V \in \mathcal{C}(U)} |\Omega_{PA_V}| \times |\Omega_V|$$

such that

1. M and N are associated with the same causal diagram, i.e., $\mathcal{G}_M = \mathcal{G}_N$.
2. M and N generate the same set of interventional distributions, i.e., for any subsets $\mathbf{X}, \mathbf{Y} \subseteq \mathbf{V}$, $P_M(\mathbf{Y}_x) = P_N(\mathbf{Y}_x)$.

Canonical SCMs - Obs. Dist.

Theorem For an arbitrary SCM $M = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P \rangle$ over discrete and finite endogenous variables \mathbf{V} , there exists a SCM N over \mathbf{V} with discrete exogenous \mathbf{U} having cardinality

$$|\Omega_U| = \prod_{V \in Pa(\mathbf{C}(U))} |\Omega_V|$$

such that

1. M and N are associated with the same causal diagram, i.e., $\mathcal{G}_M = \mathcal{G}_N$.
2. M and N generate the same observational distribution, i.e., $P_M(\mathbf{V}) = P_N(\mathbf{V})$.

This has been shown in (Rosset et al., 2018, Fraser 2020).

Bounding Counterfactuals

- Given a collection of observational and interventional distributions $\{P(\mathbf{V}|do(\mathbf{z})) \mid \mathbf{z} \in \mathbb{Z}\}$
- Qualitative assumption: causal diagram \mathcal{G}
- Assume endogenous variables \mathbf{V} are discrete and finite
- Query: $P(\mathbf{y}_x, \dots, \mathbf{z}_w)$
- let $\mathcal{N}(\mathcal{G})$ be the set of all canonical SCMs associated with \mathcal{G}

$$\begin{aligned} \min / \max_{N \in \mathcal{N}(\mathcal{G})} & P_N(\mathbf{y}_x, \dots, \mathbf{z}_w) \\ \text{s.t.} & P_N(\mathbf{V}_z) = P(\mathbf{V}_z) \quad \forall \mathbf{z} \in \mathbb{Z} \end{aligned}$$

Polynomial Optimization

- For every $U \in \mathbf{U}$, let parameters $\theta_u = P(U = u)$

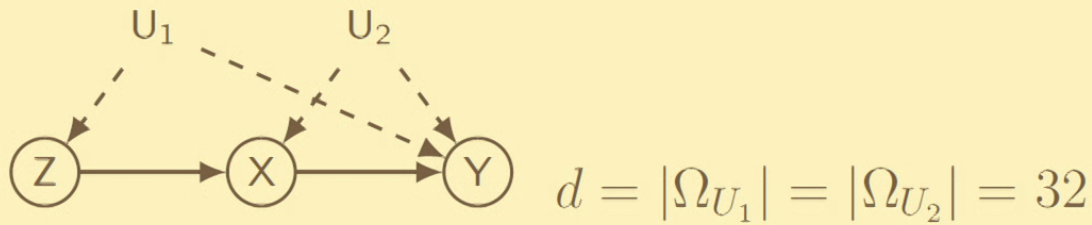
$$\theta_u \in [0, 1], \quad \sum_{u \in \Omega_U} \theta_u = 1.$$

- For every $V \in \mathbf{V}$, we represent the output of function $f_V(pa_V, u_V)$ given input pa_V, u_V using an indicator vector $\mu_V^{(pa_V, u_V)} = \left(\mu_v^{(pa_V, u_V)} \mid \forall v \in \Omega_V \right)$ such that

$$\mu_v^{(pa_V, u_V)} \in \{0, 1\}, \quad \sum_{v \in \Omega_V} \mu_v^{(pa_V, u_V)} = 1.$$

- Write any counterfactual probability as a polynomial function of parameters $\mu_v^{(pa_V, u_V)}$ and θ_u

Example



■ Objective function

$$P(x_{z'}, y_{x'}) = \sum_{u_1, u_2=1}^d \mu_x^{(z', u_2)} \mu_y^{(x', u_1, u_2)} \theta_{u_1} \theta_{u_2}$$

■ Observational constraints

$$P(x, y, z) = \sum_{u_1, u_2=1}^d \mu_z^{(u_1)} \mu_x^{(z, u_2)} \mu_y^{(x, u_1, u_2)} \theta_{u_1} \theta_{u_2}$$

Quasi-Markovian Models

- Given an interventional distribution $P_z(\mathbf{V}) = P(\mathbf{V}|do(\mathbf{z}))$
- Consider $G_{\overline{\mathbf{Z}}}$

$$\sum_{u_i=1}^{d_{u_i}} \prod_{V_j \in C_i} \mathbb{1}_{f_{V_j}(pa_j, u_i)=v_j} \theta_{u_i} = \prod_{V_j \in C_i} P_z(v_j | pa_j^+)$$

- If the target query also factorizes, e.g, $\mathbf{W} = \{Y, \dots, Z\}$

$$P(y_{pa_y}, \dots, z_{pa_z}) = \prod_i \sum_{u_i=1}^{d_{u_i}} \prod_{V_j \in \mathbf{W} \cap C_i} \mathbb{1}_{f_{V_j}(pa_j, u_i)=v_j} \theta_{u_i}$$

Bayesian Approach

- Solving polynomial optimization problems is generally hard.
- Duarte et al. (2021) presented an algorithm for bounding causal effects given data
- We develop a MCMC algorithm to approximate the bound given finite data

Bayesian Approach

- Given i.i.d. samples $\bar{\mathbf{v}} = \left\{ \mathbf{V}^{(n)} \right\}_{n=1}^N$ from $\{P(\mathbf{V}|do(\mathbf{z})) \mid \mathbf{z} \in \mathbb{Z}\}$
- Query $\theta_{\text{ctf}} = P(\mathbf{y}_x, \dots, \mathbf{z}_w)$
- Prior: For every V , $\forall pa_V, u_V$, $\mu_V^{(pa_V, u_V)}$ are drawn uniformly over domain Ω_V . For every U , θ_u are drawn from a Dirichlet distribution
- Sample the posterior distribution $P(\theta_{\text{ctf}} \mid \bar{\mathbf{v}})$ given data $\bar{\mathbf{v}}$
 1. Draw $(\boldsymbol{\mu}, \boldsymbol{\theta}) \sim P(\boldsymbol{\mu}, \boldsymbol{\theta} \mid \bar{\mathbf{v}})$ by Gibbs sampling
 2. Given parameters $\boldsymbol{\theta}, \boldsymbol{\mu}$, compute the counterfactual probability $\theta_{\text{ctf}} = P(\mathbf{y}_x, \dots, \mathbf{z}_w)$

Estimating Bounds

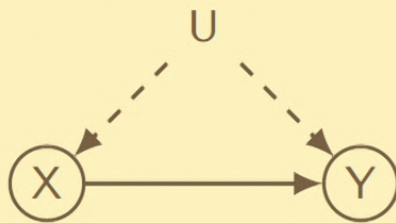
A $100(1 - \alpha)\%$ *credible interval* $[l_\alpha, r_\alpha]$: any counterfactual probability θ_{ctf} that is compatible with observational data $\bar{\mathbf{v}}$ lies between the interval l_α and r_α with probability $1 - \alpha$.

1. **Input:** Credible level α , tolerance level δ, ϵ .
2. **Output:** A credible interval $[l_\alpha, h_\alpha]$ for θ_{ctf} .
3. Draw $T = \lceil 2\epsilon^{-2} \ln(4/\delta) \rceil$ samples $\{\theta^{(1)}, \dots, \theta^{(T)}\}$ from the posterior distribution $P(\theta_{\text{ctf}} \mid \bar{\mathbf{v}})$.
4. Return interval $[\hat{l}_\alpha(T), \hat{r}_\alpha(T)]$.

$$l_0 = \min_t \theta^{(t)}, \quad r_0 = \max_t \theta^{(t)}$$

$[l_0, r_0]$ converges almost surely to the optimal bound as $N, T \rightarrow \infty$

Simulation Example



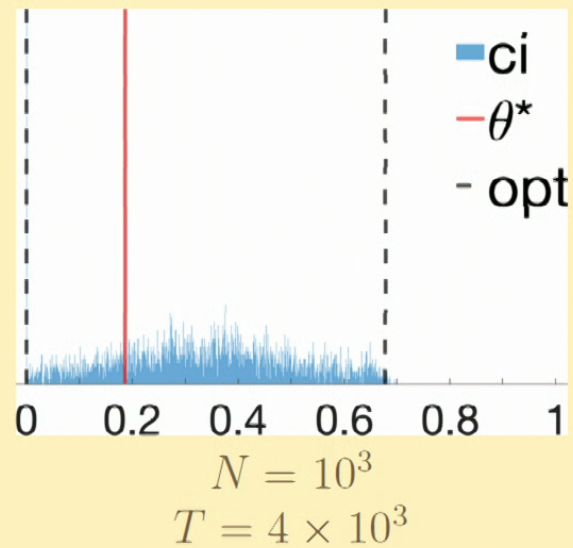
$X, Y \in \{0, 1\}$

$d = |\Omega_U| = 8$

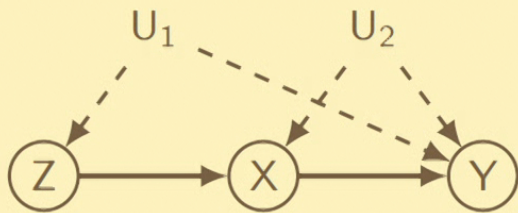
Data: $P(X, Y)$

Query: PNS

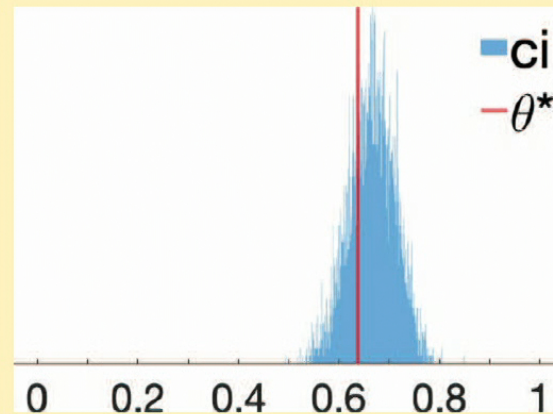
$PNS = P(Y_{x=1} = 1, Y_{x=0} = 0)$ opt: sharp bound (Tian&Pearl 2000)



Simulation Example



$X, Y, Z \in \{0, \dots, 9\}$
 $d = |\Omega_{U_1}| = |\Omega_{U_2}| = 10^{21}$
Data: $P(X, Y, Z)$,
 $P(X_z, Y_z)$ for $z = 0, \dots, 9$
Query $P(z, x_{z'}, y_{x'})$



$N = 10^3$
 $T = 4 \times 10^3$
 $\theta = P(Z + X_{z=0} + Y_{x=0} \geq 14)$

Conclusion

- We study the problem of bounding counterfactual probabilities from observational and experimental data given a causal diagram
- We introduce a family of canonical SCMs over discrete endogenous variables with discrete exogenous variables
- We show canonical SCMs could represent *all* counterfactual distributions over discrete observed variables in *any* causal diagram.
- We reduce the partial identification problem into a polynomial program
- We develop an MCMC algorithm to approximate the optimal bounds from finite samples.