

Title: Peaks, excursion sets and the distribution of dark matter haloes

Date: Sep 03, 2013 11:00 AM

URL: <http://pirsa.org/13090062>

Abstract: I will review recent developments in our theoretical understanding of the abundance and clustering of dark matter haloes. In the first part of this talk, I will discuss a toy model based on the statistics of peaks of Gaussian random field (Bardeen et al 1986) and show how the clustering properties of such a point set can be easily derived from a generalised local bias expansion. In the second part, I will explain how this peak formalism relates to the excursion set approach and present parameter-free predictions for the mass function and bias of dark matter halos.

OUTLINE

- *Galaxy clustering*
- *The peak formalism: BBKS and beyond*
- *Local bias approach to BBKS peaks*
- *Excursion set peaks and the halo mass function*

Galaxies positions now routinely measured out to high redshift ...

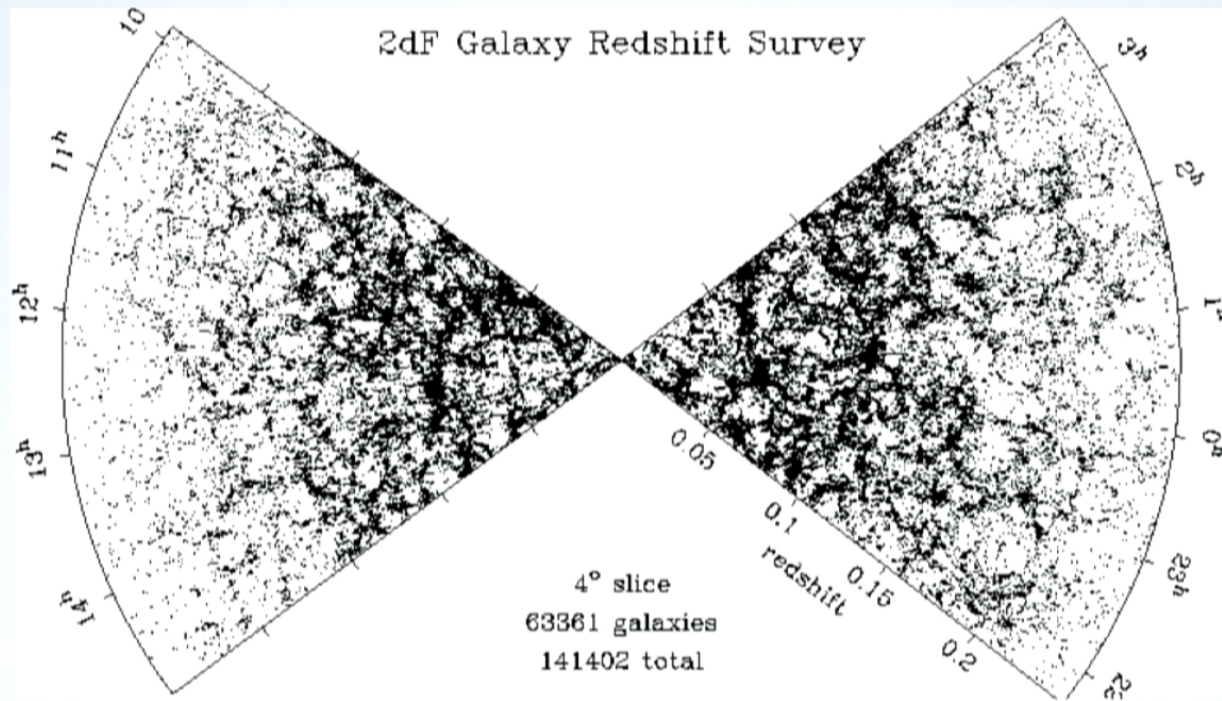


Image credit: 2dF GRS

SCIENCE WITH LARGE GALAXY SURVEYS

A plethora of ongoing and planned galaxy surveys, e.g. BOSS, DES, EUCLID, Pan-STARR, VIPERS, WiggleZ ..., to measure (among others)

- Matter power spectrum: matter content, neutrino masses etc.
- Baryon acoustic oscillation (BAO): dark energy equation of state
- Redshift space distortions: growth rate of structure (probe of gravity)
- Galaxy bispectrum: primordial non-Gaussianity (probe of inflation) etc.

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Require solid understanding of galaxy clustering !

BIAS

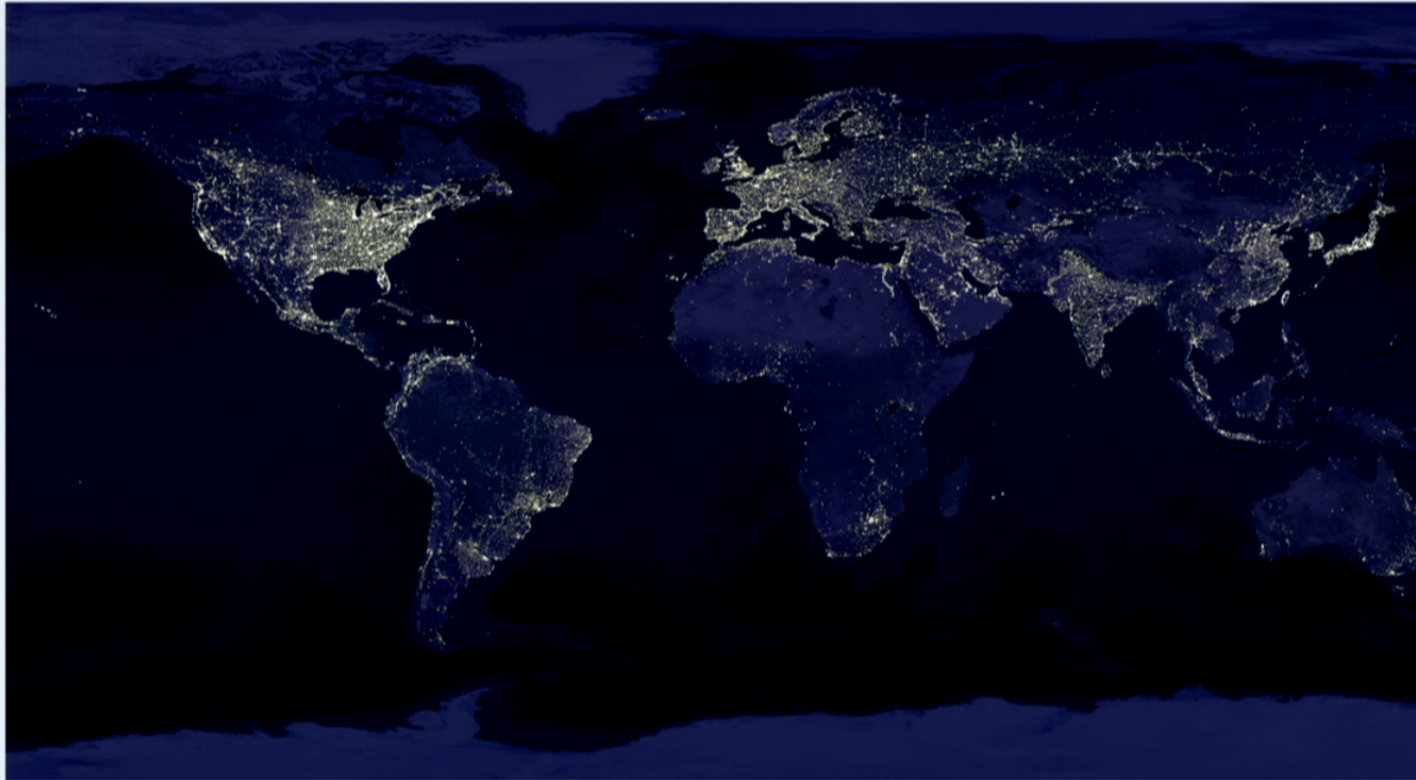


Figure credit: NASA

BIAS

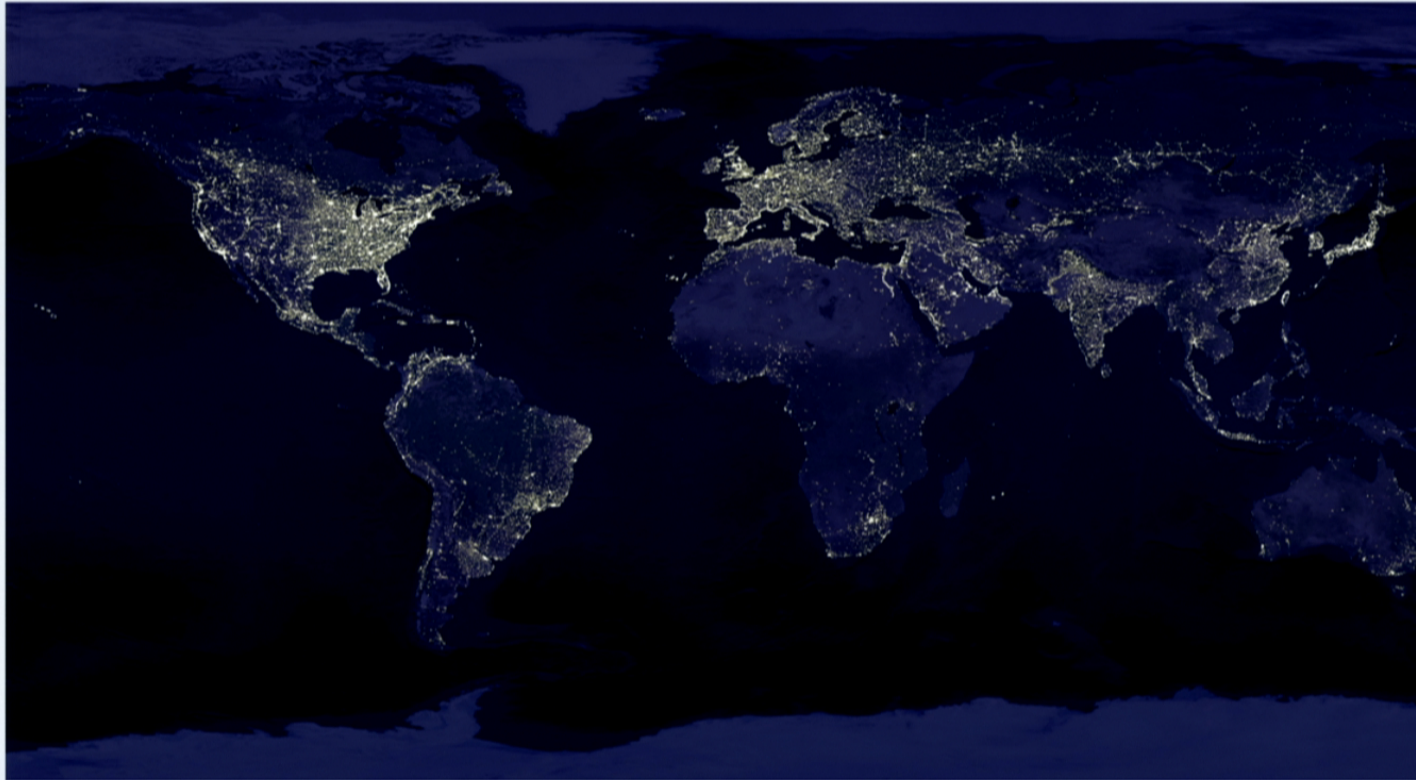


Figure credit: NASA

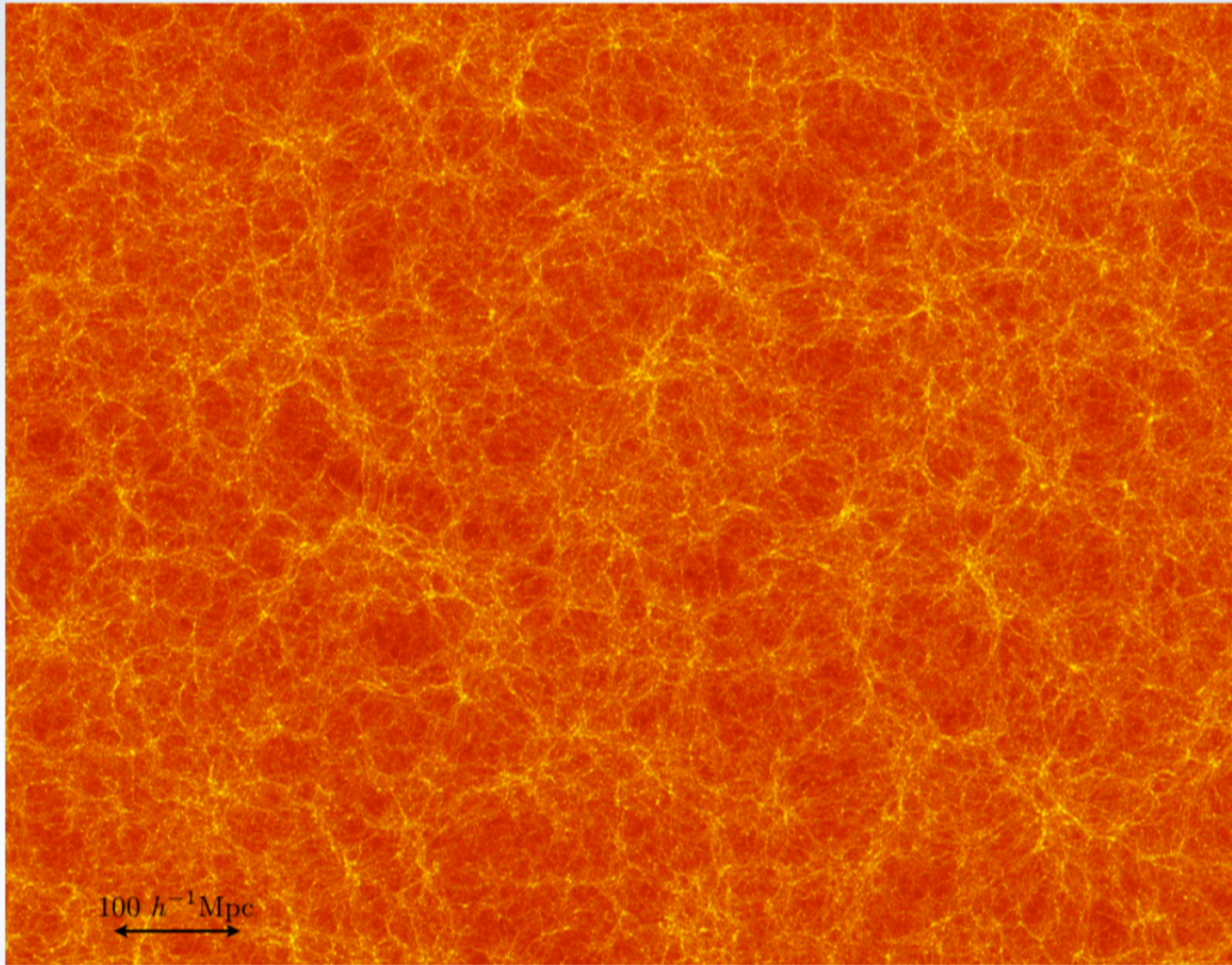


Figure courtesy Ilian T. Iliev

HALO BIASING

In CDM cosmologies, galaxies form in the dense, virialized regions of the dark matter distribution: dark matter halos

$$\delta_h = b \cdot \delta$$

halo overdensity

dark matter overdensity

linear halo bias (Kaiser 1984)

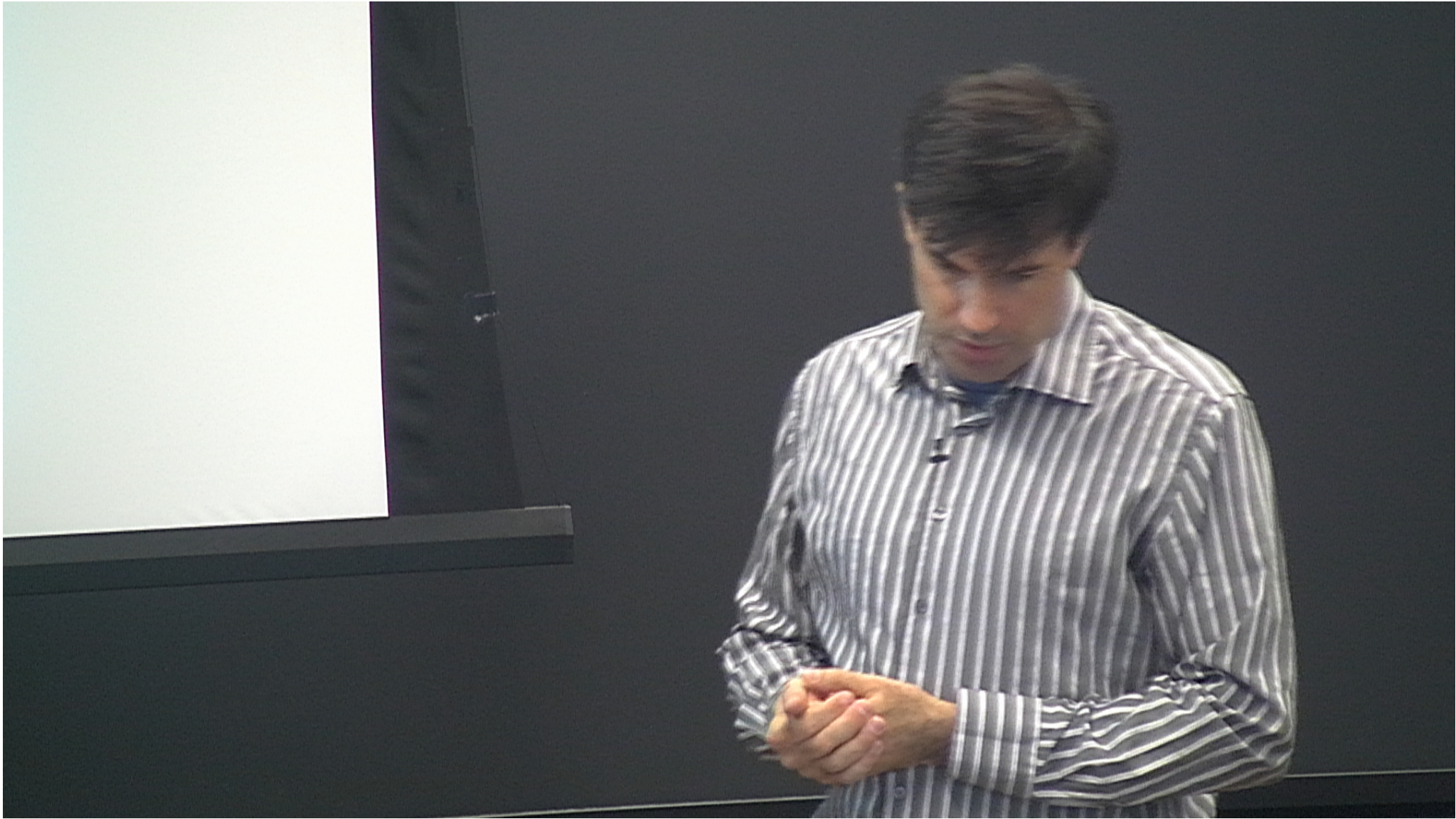
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LOCAL LAGRANGIAN BIAS

Simplest bias scheme: dependence on mass density field only (Szalay 1988; Fry & Gaztanaga 1993, ...):

$$\delta_h(\mathbf{x}) = b_1\delta_R(\mathbf{x}) + \frac{1}{2}b_2\delta_R^2(\mathbf{x}) + \frac{1}{6}b_3\delta_R^3(\mathbf{x}) + \dots$$

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For example, 2-point correlation is computed as:

$$\xi_h(|\mathbf{x}_2 - \mathbf{x}_1|) = \langle \delta_h(\mathbf{x}_1) \delta_h(\mathbf{x}_2) \rangle = b_1^2 \xi_R(|\mathbf{x}_2 - \mathbf{x}_1|) + \dots$$

TWO REMARKS

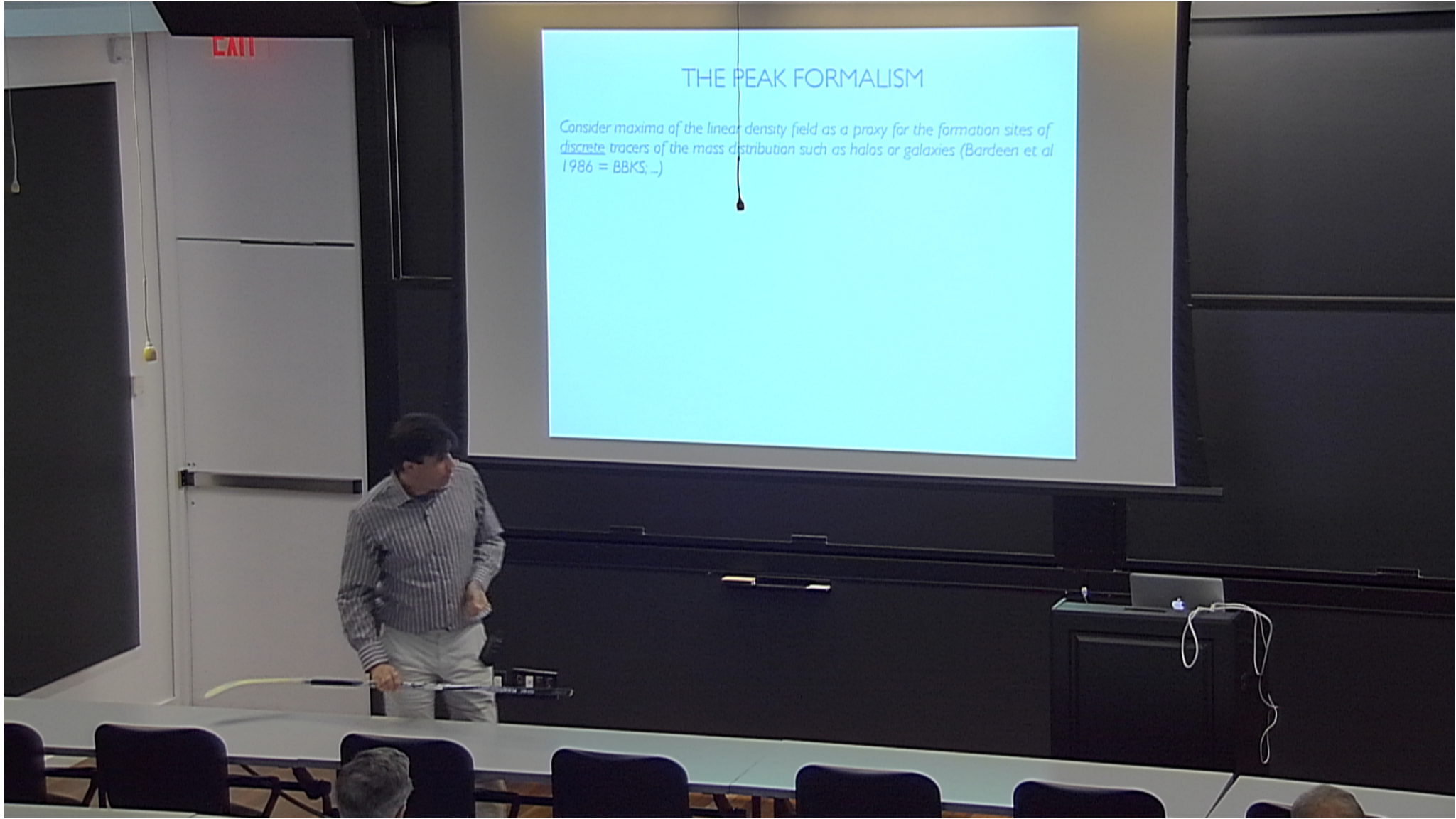
- *One could think of a more general expansion:*

$$\delta_h(\mathbf{x}) = b_1 \delta_R(\mathbf{x}) + c_1 \nabla^2 \delta_R(\mathbf{x}) + \frac{1}{2} b_2 \delta_R^2(\mathbf{x}) + c_2 (\nabla \delta_R)^2(\mathbf{x}) + c_3 \left(\partial_i \partial_j \delta_R - \frac{1}{3} \delta_{ij} \nabla^2 \delta_R \right)^2(\mathbf{x}) + \dots$$

- *Galaxies and dark matter halos are pointlike objects as far as their 3D position is concerned. Does this matter (beyond a Poisson shot noise correction to the power spectrum) ?*

THE PEAK FORMALISM

Consider maxima of the linear density field as a proxy for the formation sites of discrete tracers of the mass distribution such as halos or galaxies (Bardeen et al 1986 = BBKS; ...)



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$$\nu(\mathbf{x}) \equiv \frac{1}{\sigma_0} \delta_s(\mathbf{x}), \quad \eta_i(\mathbf{x}) = \frac{1}{\sigma_1} \partial_i \delta_s(\mathbf{x}), \quad \zeta_{ij} \equiv \frac{1}{\sigma_2} \partial_i \partial_j \delta_s(\mathbf{x}), \quad \delta_s(\mathbf{x}) = \delta_{R=R_s}(\mathbf{x})$$

$$\sigma_n^2 = \frac{1}{2\pi^2} \int_0^\infty dk k^{2(n+1)} P_s(k)$$

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$$\sigma_n^2 = \frac{1}{2\pi^2} \int_0^\infty dk k^{2n+1} P_\delta(k)$$

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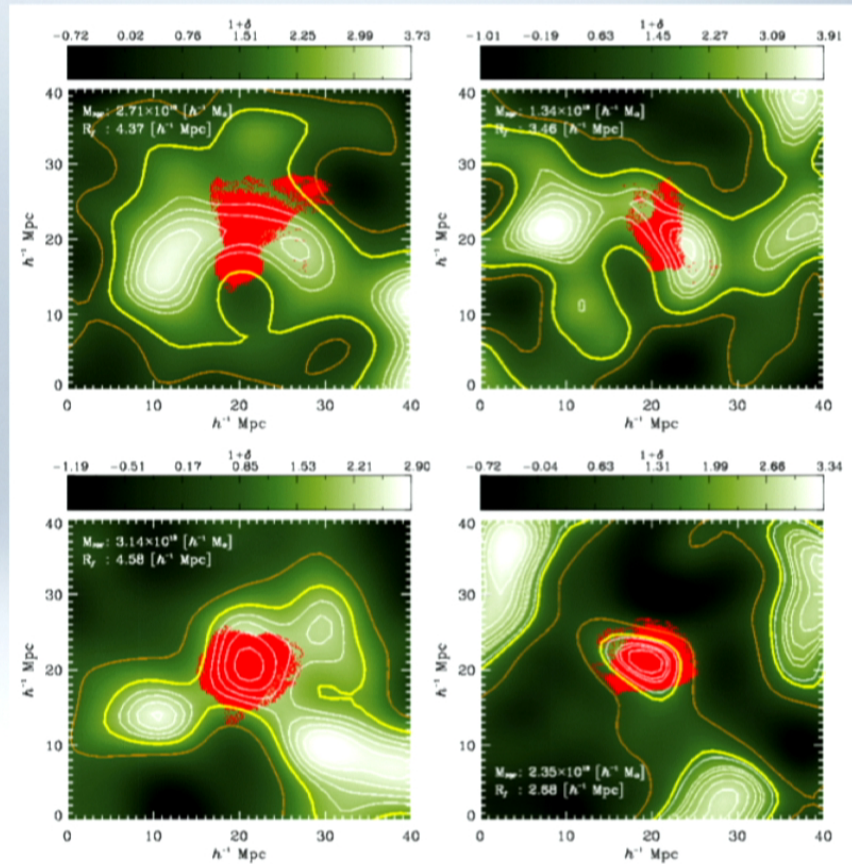
$$(i) \quad \nu(\mathbf{x}_p) = \frac{\delta_c}{\sigma_0} = \frac{1.68}{\sigma_0} \equiv \nu_c$$

$$(ii) \quad \eta_i(\mathbf{x}_p) = 0$$

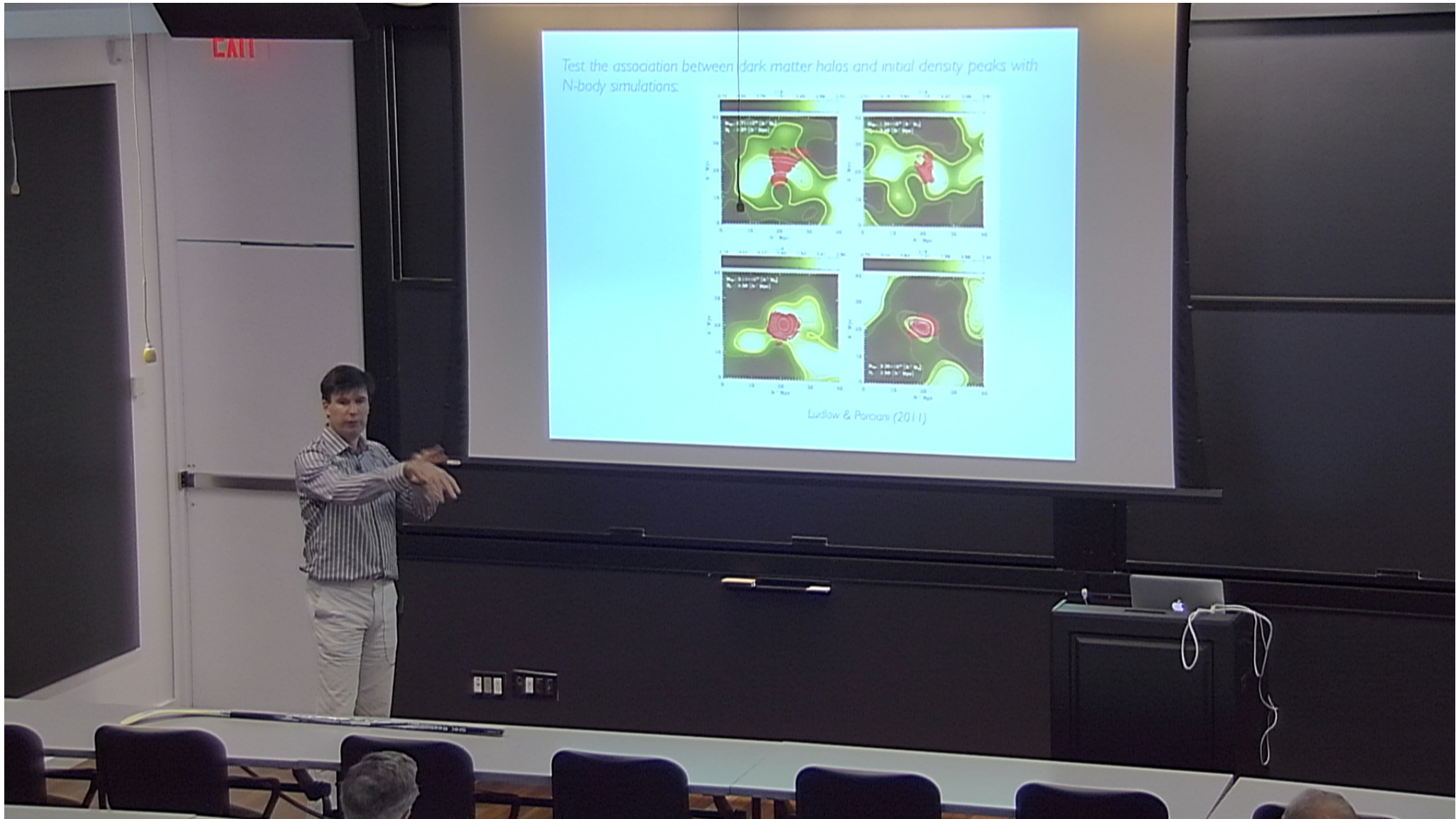
$$(iii) \quad \lambda_1(\mathbf{x}_p) \geq \lambda_2(\mathbf{x}_p) \geq \lambda_3(\mathbf{x}_p) > 0$$

$$\lambda_a(\mathbf{x}) = \text{eigenvalues of } -\zeta_{ij}(\mathbf{x})$$

Test the association between dark matter halos and initial density peaks with N-body simulations:



Ludlow & Porciani (2011)



PEAK CORRELATIONS

The “localized” peak number density is (Kac 1943; Rice 1951; BBKS):

$$n_{\text{pk}}(\mathbf{x}) = \sum_{\mathbf{x}_p} \delta_D(\mathbf{x} - \mathbf{x}_p) = \frac{3^{3/2}}{R_*^3} |\det \zeta(\mathbf{x})| \delta_D[\boldsymbol{\eta}(\mathbf{x})] \theta_H[\lambda_3(\mathbf{x})] \delta_D(\nu(\mathbf{x}) - \nu_c)$$



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$$n_{\text{pk}}(\mathbf{x}) = \sum_{\mathbf{x}_0} \delta_D(\mathbf{x} - \mathbf{x}_0) = \frac{3^{3/2}}{4\pi^2} |\det \zeta(\mathbf{x})| \delta_D[\eta(\mathbf{x})] \theta_H[\lambda_1(\mathbf{x})] \delta_D(\nu(\mathbf{x}) - \nu_c)$$

MEAN NUMBER DENSITY

First computed in BBKS:

$$\langle n_{\text{pk}}(\mathbf{x}) \rangle \equiv \bar{n}_{\text{pk}}(\nu_c) = V_{\star}^{-1} G_0^{(1)}(\gamma_1, \gamma_1 \nu_c) \frac{e^{-\nu_c^2/2}}{\sqrt{2\pi}}$$

$$V_{\star} \equiv (2\pi)^{3/2} R_{\star}^3, \quad \gamma_1 \equiv \frac{\sigma_1^2}{\sigma_0 \sigma_2}$$

$$G_n^{(\alpha)}(\gamma_1, \omega) = \int_0^{\infty} du u^n f(u, \alpha) \frac{\exp\left[-\frac{(u-\omega)^2}{2(1-\gamma_1^2)}\right]}{\sqrt{2\pi(1-\gamma_1^2)}}$$

$$f(u, \alpha) \equiv \frac{3^2 5^{5/2}}{\sqrt{2\pi}} \left\{ \int_0^{u/4} dv \int_{-v}^{+v} dw + \int_{u/4}^{u/2} dv \int_{3v-w}^v dw \right\} F(u, v, w) e^{-\frac{5\alpha}{2}(3v^2+w^2)}$$

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$u \equiv -\text{tr}\zeta = \text{peak curvature}$

$v, w \sim \text{peak ellipticity and prolateness}$

2-POINT CORRELATION

At 2nd order by VD, Crocce, Scoccimarro & Sheth (2010):

$$\begin{aligned} \xi_{pk}(\nu_c, r) = & (\mathbf{b}_I^2 \xi_0^{(0)}) + \frac{1}{2} (\xi_0^{(0)} \mathbf{b}_{II}^2 \xi_0^{(0)}) - \frac{3}{\sigma_1^2} (\xi_1^{(1/2)} \mathbf{b}_{II} \xi_1^{(1/2)}) - \frac{5}{\sigma_2^2} (\xi_2^{(1)} \mathbf{b}_{II} \xi_2^{(1)}) \left(1 + \frac{2}{5} \partial_\alpha \ln G_0^{(\alpha)}(\gamma_1, \gamma_1 \nu_c) \Big|_{\alpha=1} \right) \\ & + \frac{5}{2\sigma_2^4} \left[(\xi_0^{(2)})^2 + \frac{10}{7} (\xi_2^{(2)})^2 + \frac{18}{7} (\xi_4^{(2)})^2 \right] \left(1 + \frac{2}{5} \partial_\alpha \ln G_0^{(\alpha)}(\gamma_1, \gamma_1 \nu_c) \Big|_{\alpha=1} \right)^2 \\ & + \frac{3}{2\sigma_1^4} \left[(\xi_0^{(1)})^2 + 2(\xi_2^{(1)})^2 \right] + \frac{3}{\sigma_1^2 \sigma_2^2} \left[3(\xi_3^{(3/2)})^2 + 2(\xi_1^{(3/2)})^2 \right] \left(1 + \frac{2}{5} \partial_\alpha \ln G_0^{(\alpha)}(\gamma_1, \gamma_1 \nu_c) \Big|_{\alpha=1} \right) \end{aligned}$$

$$\xi_\ell^{(n)}(r) = \frac{1}{2\pi^2} \int_0^\infty dk k^{2(n+1)} P_s(k) j_\ell(kr)$$

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$$\xi_0^{(n)}(r) = \frac{1}{2\pi^2} \int_0^\infty dk k^{2n-1} P_s(k) J_1(kr)$$

PEAK BIAS FACTORS

1st order:

$$\mathfrak{b}_I(k) = b_{10} + b_{01}k^2, \quad b_{10} = \frac{1}{\sigma_0} \left(\frac{\nu_c - \gamma_1 \bar{u}}{1 - \gamma_1^2} \right), \quad b_{01} = \frac{1}{\sigma_2} \left(\frac{\bar{u} - \gamma_1 \nu_c}{1 - \gamma_1^2} \right)$$

2nd order :

$$\mathfrak{b}_{II}(k_1, k_2) = b_{20} + b_{11}(k_1^2 + k_2^2) + b_{02} k_1^2 k_2^2, \quad b_{20} = \frac{1}{\sigma_0^2} \left[\frac{\nu_c^2 - 2\gamma_1 \nu_c \bar{u} + \gamma_1^2 \bar{u}^2}{(1 - \gamma_1^2)^2} - \frac{1}{(1 - \gamma_1^2)} \right]$$

$$b_{11} = \frac{1}{\sigma_0 \sigma_2} \left[\frac{(1 + \gamma_1^2) \nu_c \bar{u} - \gamma_1 [\nu_c^2 + \bar{u}^2]}{(1 - \gamma_1^2)^2} + \frac{\gamma_1}{(1 - \gamma_1^2)} \right]$$

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$$b_{11}(k_1, k_2) = b_{20} + b_{11}(k_1^2 + k_2^2) + b_{12}k_1^2k_2^2, \quad b_{20} = \frac{1}{\sigma_0^2} \left[\frac{\mu_c^2 - 2\gamma_1 \mu_c u + \gamma_1^2 u^2}{(1 - \gamma_1^2)^2} - \frac{1}{(1 - \gamma_1^2)} \right]$$

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EFFECTIVE LOCAL BIAS EXPANSION (I)

$$(b_{11}^2 \xi_0^{(0)}) + \frac{1}{2} (\xi_0^{(0)} b_{11}^2 \xi_0^{(0)}) :$$

Can be thought of as arising from the effective continuous, deterministic local bias expansion:

$$\begin{aligned} \delta_{pk}(\mathbf{x}) &= b_{10} \delta_s(\mathbf{x}) - b_{01} \nabla^2 \delta_s(\mathbf{x}) + \frac{1}{2} b_{20} \delta_s^2(\mathbf{x}) - b_{11} \delta_s(\mathbf{x}) \nabla^2 \delta_s(\mathbf{x}) + \frac{1}{2} b_{02} [\nabla^2 \delta_s(\mathbf{x})]^2 + \dots \\ &= \sigma_0 b_{10} \nu(\mathbf{x}) + \sigma_2 b_{01} u(\mathbf{x}) + \frac{1}{2} \sigma_0^2 b_{20} \nu^2(\mathbf{x}) + \sigma_0 \sigma_2 b_{11} \nu(\mathbf{x}) u(\mathbf{x}) + \frac{1}{2} \sigma_2^2 b_{02} u^2(\mathbf{x}) + \dots \end{aligned}$$

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SYMMETRIES

- *Peak number density invariant under rotations, so must depend on:*

$$\nu(\mathbf{x})$$

$$u(\mathbf{x}) \equiv -\text{tr}\zeta(\mathbf{x})$$

VD (2013), 1211.4128

2-POINT CORRELATION

At 2nd order by VD, Crocce, Scoccimarro & Sheth (2010):

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SYMMETRIES

- *Peak number density invariant under rotations, so must depend on:*

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$$\zeta^2(\mathbf{x}) \equiv \frac{3}{2}\text{tr}(\tilde{\zeta}^2)(\mathbf{x}), \quad \tilde{\zeta}_{ij} = \zeta_{ij} + \frac{1}{3}u\delta_{ij}$$

$$\det\zeta(\mathbf{x})$$

VD (2013), 1211.4128

CORRELATORS

Define $y_\alpha \equiv y(\mathbf{x}_\alpha)$

$$\langle \eta_1^2 \eta_2^2 \rangle = 1 + \frac{2}{3\sigma_1^4} \left[(\xi_0^{(1)})^2 + 2(\xi_2^{(1)})^2 \right]$$

$$\langle \eta_1^2 \nu_2^2 \rangle = 1 + \frac{2}{\sigma_0^2 \sigma_1^2} \xi_1^{(1/2)} \xi_1^{(1/2)}$$

$$\langle \zeta_1^2 \eta_2^2 \rangle = 1 + \frac{2}{5\sigma_1^2 \sigma_2^2} \left[3(\xi_3^{(3/2)})^2 + 2(\xi_1^{(3/2)})^2 \right]$$

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PROBABILITY DENSITIES

Probability distribution $P_1(\mathbf{y})$ needed to compute \bar{n}_{pk} is the multivariate Gaussian

$$P_1(\mathbf{y})d^{10}\mathbf{y} = \frac{1}{(2\pi)^5|\det M|^{1/2}}e^{-Q_1(\mathbf{y})}d^{10}\mathbf{y}$$

The quadratic form is

$$Q_1(\mathbf{y}) = \frac{\nu^2 + u^2 - 2\gamma_1\nu u}{2(1 - \gamma_1^2)} + \frac{3}{2}\eta^2 + \frac{5}{2}\zeta^2$$

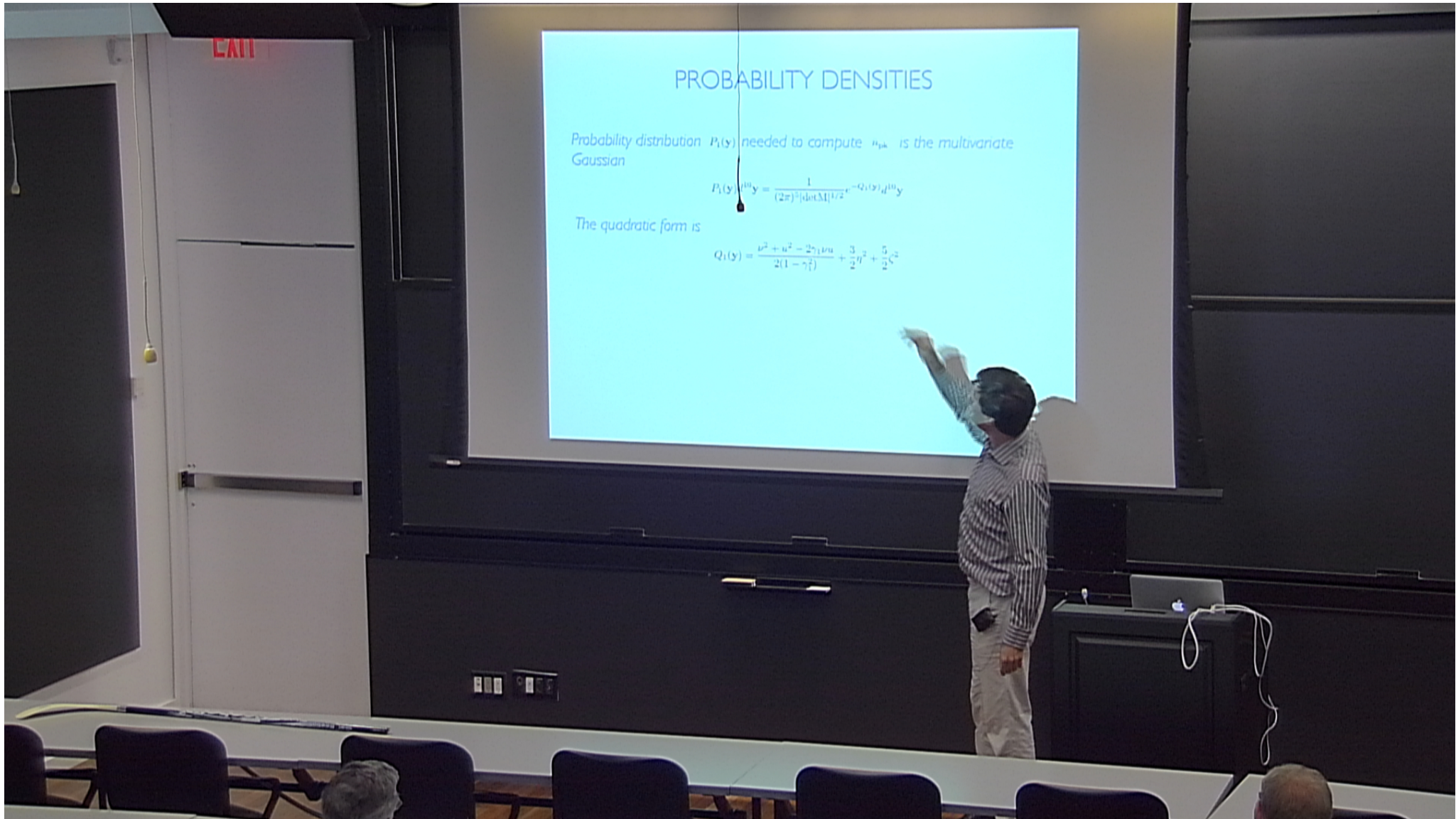
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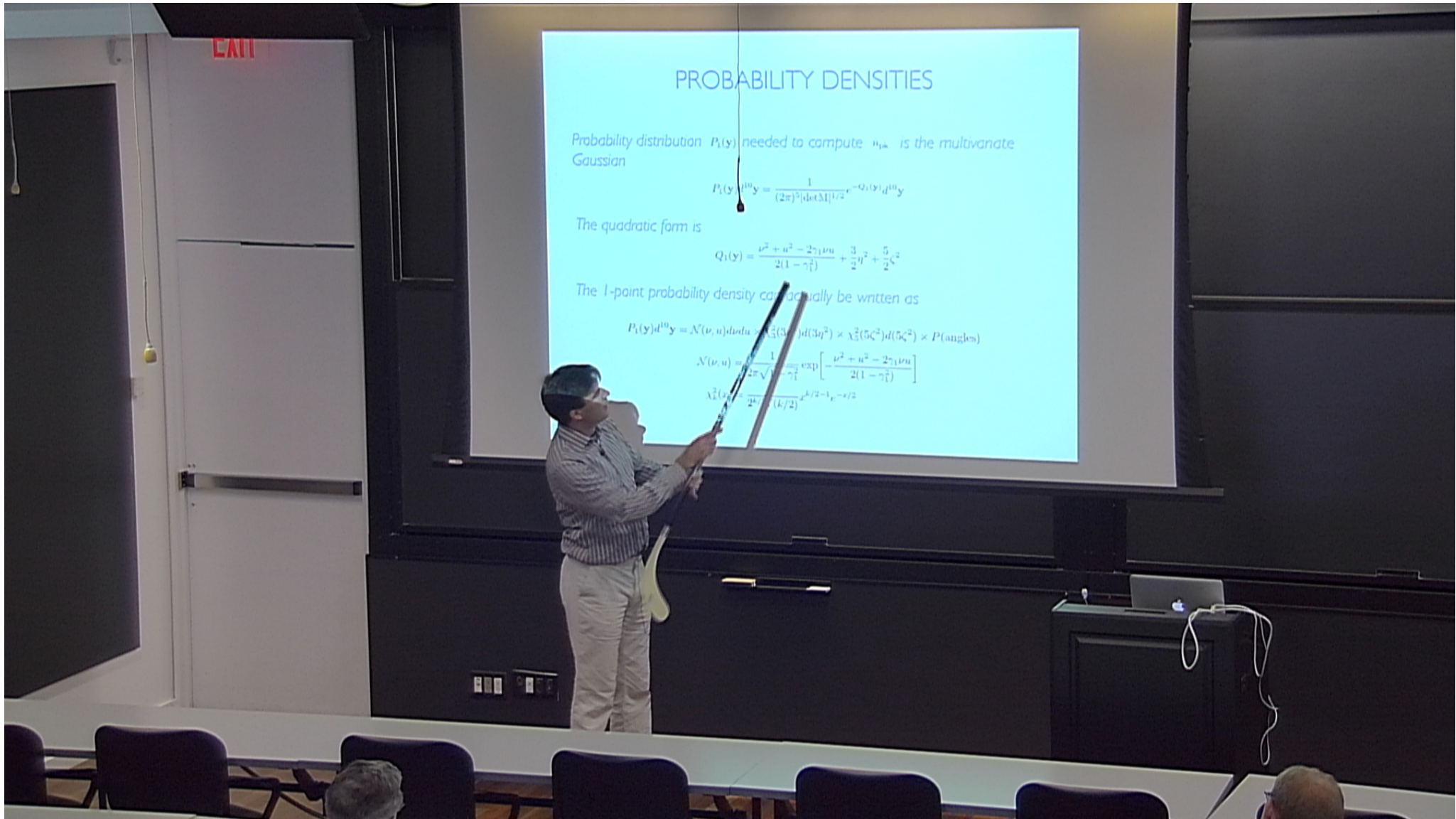
PROBABILITY DENSITIES

Probability distribution $P_1(\mathbf{y})$ needed to compute μ_{jk} is the multivariate Gaussian

$$P_1(\mathbf{y}) = \frac{1}{(2\pi)^3 |\det \Sigma|^{1/2}} e^{-Q_1(\mathbf{y})} \mu^{(1)} \mathbf{y}$$

The quadratic form is

$$Q_1(\mathbf{y}) = \frac{y^2 + u^2 - 2\gamma_1 yu}{2(1 - \gamma_1^2)} + \frac{3}{2}y^2 + \frac{5}{2}u^2$$



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The quadratic form is

$$Q_1(\mathbf{y}) = \frac{v^2 + u^2 - 2\gamma_1 uv}{2(1 - \gamma_1^2)} + \frac{3}{2}v^2 + \frac{5}{2}u^2$$

The 1-point probability density can actually be written as

$$P_1(\mathbf{y}) d^{10}\mathbf{y} = \mathcal{N}(v, u) dv du \times \chi_3^2(3v^2) d(3v^2) \times \chi_5^2(5u^2) d(5u^2) \times P(\text{angles})$$

$$\mathcal{N}(v, u) = \frac{1}{2\pi\sqrt{1 - \gamma_1^2}} \exp\left[-\frac{v^2 + u^2 - 2\gamma_1 uv}{2(1 - \gamma_1^2)}\right]$$

$$\chi_k^2(x) = \frac{1}{2^{k/2} \Gamma(k/2)} x^{k/2 - 1} e^{-x/2}$$

GENERALIZING PBS (I)

- Bias factors b_{ij} can be written as derivatives of $\mathcal{N}(\nu, u)$, i.e. they are bivariate Hermite polynomials averaged over all the possible peak configurations:

$$\sigma_0^i \sigma_2^j b_{ij} \equiv \frac{1}{\bar{n}_{\text{pk}}} \int d^{10} \mathbf{y} n_{\text{pk}}(\mathbf{y}) H_{ij}(\nu, u) P_1(\mathbf{y})$$

$$H_{ij}(\nu, u) = \mathcal{N}(\nu, u)^{-1} \left(-\frac{\partial}{\partial \nu} \right)^i \left(-\frac{\partial}{\partial u} \right)^j \mathcal{N}(\nu, u)$$

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- The same logic must apply to the bias factors associated with $\eta^2(\mathbf{x})$ and $\zeta^2(\mathbf{x})$. We thus expect ensemble averages of generalized Laguerre polynomials:

$$\sigma_1^{2k} \chi_{k0} \equiv \frac{(-1)^k}{\bar{n}_{\text{pk}}} \int d^{10} \mathbf{y} n_{\text{pk}}(\mathbf{y}) L_k^{(1/2)} \left(\frac{3\eta^2}{2} \right) P_1(\mathbf{y})$$

$$\sigma_2^{2k} \chi_{0k} \equiv \frac{(-1)^k}{\bar{n}_{\text{pk}}} \int d^{10} \mathbf{y} n_{\text{pk}}(\mathbf{y}) L_k^{(3/2)} \left(\frac{5\zeta^2}{2} \right) P_1(\mathbf{y})$$

$$L_n^{(\alpha)}(x) \equiv \frac{x^{-\alpha} e^x}{n!} \frac{d^n}{dx^n} (e^{-x} x^{n+\alpha})$$

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GENERALIZING PBS (II)

- A straightforward calculation shows that

$$\chi_{10} = \frac{1}{2\sigma_1^2} \left(\langle 3\eta^2 | \text{pk} \rangle - 3 \right) = -\frac{3}{2\sigma_1^2}$$

$$\chi_{01} = \frac{1}{2\sigma_2^2} \left(\langle 5\zeta^2 | \text{pk} \rangle - 5 \right) = -\frac{5}{2\sigma_2^2} \left(1 + \frac{2}{5} \partial_\alpha \ln G_0^{(\alpha)}(\gamma_1, \gamma_1 \nu_c) \Big|_{\alpha=1} \right)$$

- Physical interpretation: “long”-wavelength perturbations locally modulate the mean of the distributions $\mathcal{N}(\nu, u)$, $\chi_3^2(3\eta^2)$ and $\chi_5^2(5\zeta^2)$; e.g.

$$\chi_k'^2(x; \epsilon) = \chi_k^2(x) \sum_{j=0}^{\infty} \frac{\Gamma(\frac{1}{2}k)}{\Gamma(\frac{1}{2}k + j)} \left(-\frac{\epsilon}{2}\right)^j L_j^{(\alpha)}\left(\frac{x}{2}\right)$$

χ_{01} : PEAK ASPHERICITY

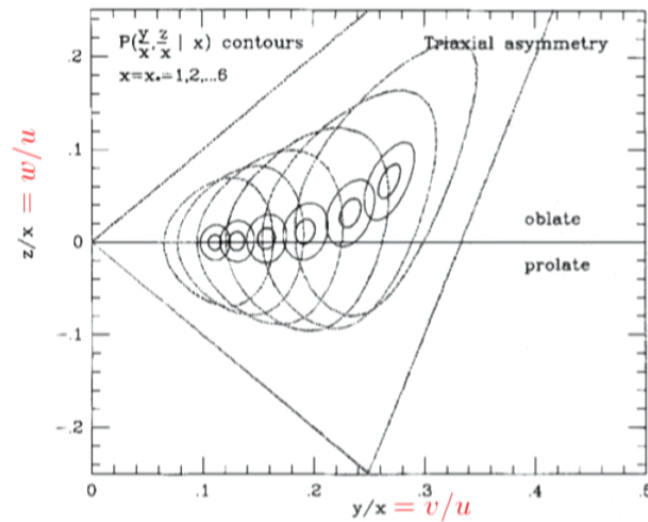


FIG. 7.—The 95%, 90%, and 50% contours of the conditional probability for ellipticity $e = y/x$ and prolateness $p = z/x$ subject to the constraint of given x for peaks (eq. [7.6]). (The x and x_0 used here are $1.58 \approx \gamma^{-1}$ times those used in the text, so $v = 1, 2, \dots, 6$ corresponds to the different curves.) This figure demonstrates that, even for high v , the shapes are triaxial. The values of e and p are constrained to lie in the triangle.

Bardeen et. al. (1986)

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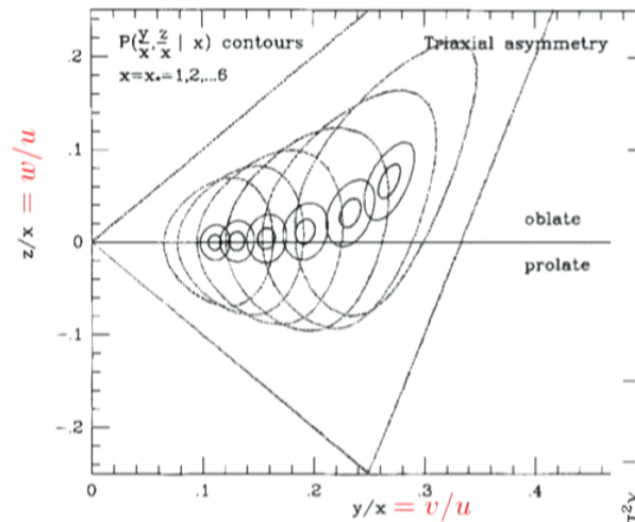
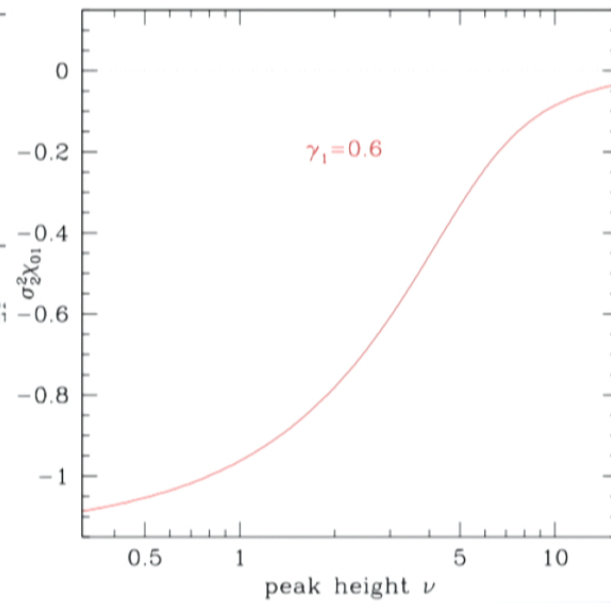
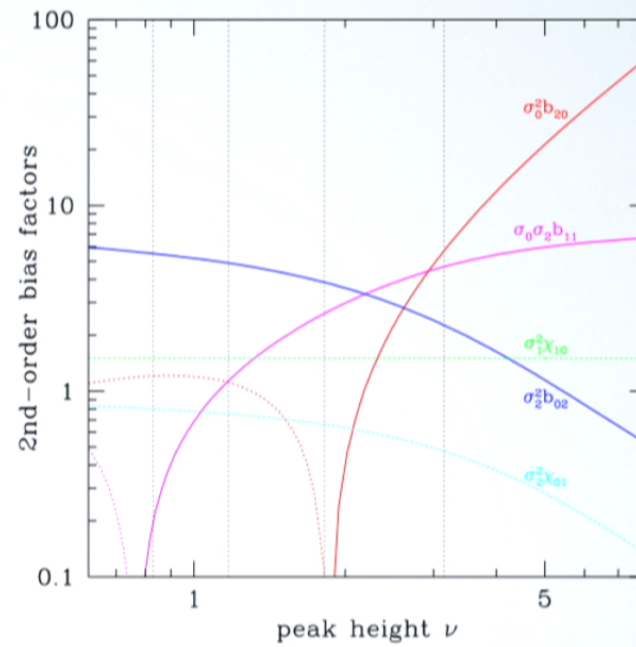
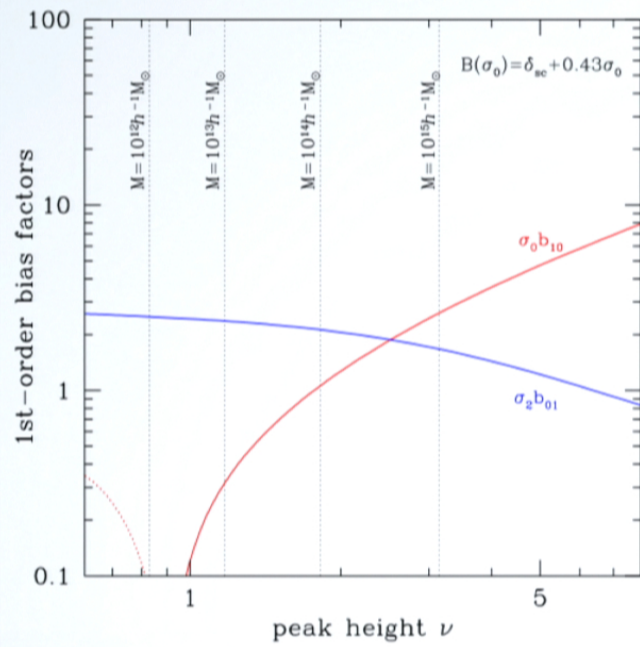


FIG. 7.—The 95%, 90%, and 50% contours of the conditional probability for ellipticity $e = y/x$ and prolateness $p = z/x$ subject to the peaks (eq. [7.6]). (The x and x_i used here are $1.58 = \gamma^{-1}$ times those used in the text, so $v = 1, 2, \dots, 6$ corresponds to the different curves.) That, even for high v , the shapes are triaxial. The values of e and p are constrained to lie in the triangle.

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PEAK BIAS FACTORS



EFFECTIVE LOCAL BIAS EXPANSION (II)

$\xi_{pk}(\nu_c, r)$ can be computed from the local bias expansion

$$\begin{aligned}\delta_{pk}(\mathbf{x}) = & \sigma_0 b_{10} \nu(\mathbf{x}) + \sigma_2 b_{01} u(\mathbf{x}) + \frac{1}{2} \sigma_0^2 b_{20} \nu^2(\mathbf{x}) + \sigma_0 \sigma_2 b_{11} \nu(\mathbf{x}) u(\mathbf{x}) + \frac{1}{2} \sigma_2^2 b_{02} u^2(\mathbf{x}) \\ & + \sigma_1^2 \chi_{10} \eta^2(\mathbf{x}) + \sigma_2^2 \chi_{01} \zeta^2(\mathbf{x}) + \dots\end{aligned}$$

provided we ignore all correlators involving zero-lag moments

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Speculation: exact to all orders and valid for any point process of a Gaussian random field specified as a set of "localized" constraints

COUNTING DARK MATTER HALOS



Image credit: Aquarius Project (Springel et. al. 2008)

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EFFECTIVE LOCAL BIAS EXPANSION (II)

$\zeta_{\mu}(u, v)$ can be computed from the local bias expansion

$$\delta_{\mu}(x) = \sigma_0 b_{10} u(x) + \sigma_2 b_{02} v(x) + \frac{1}{2} \sigma_0^2 b_{20} u^2(x) + \sigma_0 \sigma_2 b_{11} u(x) v(x) + \frac{1}{2} \sigma_2^2 b_{02} v^2(x) + \sigma_1^2 \chi_{10} u^2(x) + \sigma_1^2 \chi_{01} v^2(x) + \dots$$

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$$S_{ph} = b_1 H_1(\phi) + \frac{1}{2} b_2 H_2(\phi) + \dots + H_2(\phi)$$
$$S_{ph}(\cdot) = \langle \dots \rangle$$

$$\delta_{pk} = \psi_1 H_1(\delta) + \frac{1}{2} \psi_2 H_2(\delta) + \frac{1}{3!} \psi_3 H_3(\delta) + \dots$$

$$\Sigma_{pk}(r) \equiv \langle \delta_{pk}(x_1) \delta_{pk}(x_2) \rangle$$

$$\equiv \psi_1^2 \Sigma(r)$$

COUNTING DARK MATTER HALOS



Image credit: Aquarius Project (Springel et. al. 2008)

EXCURSION SET PEAKS (I)

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- *Smooth the linear density field with a tophat filter on scale $R_s \propto (M/\bar{\rho})^{1/3}$*
- *Select all the peaks with height $\nu = \nu_c$ and (Appel & Jones 1991)*

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Approximation to the first upcrossing condition

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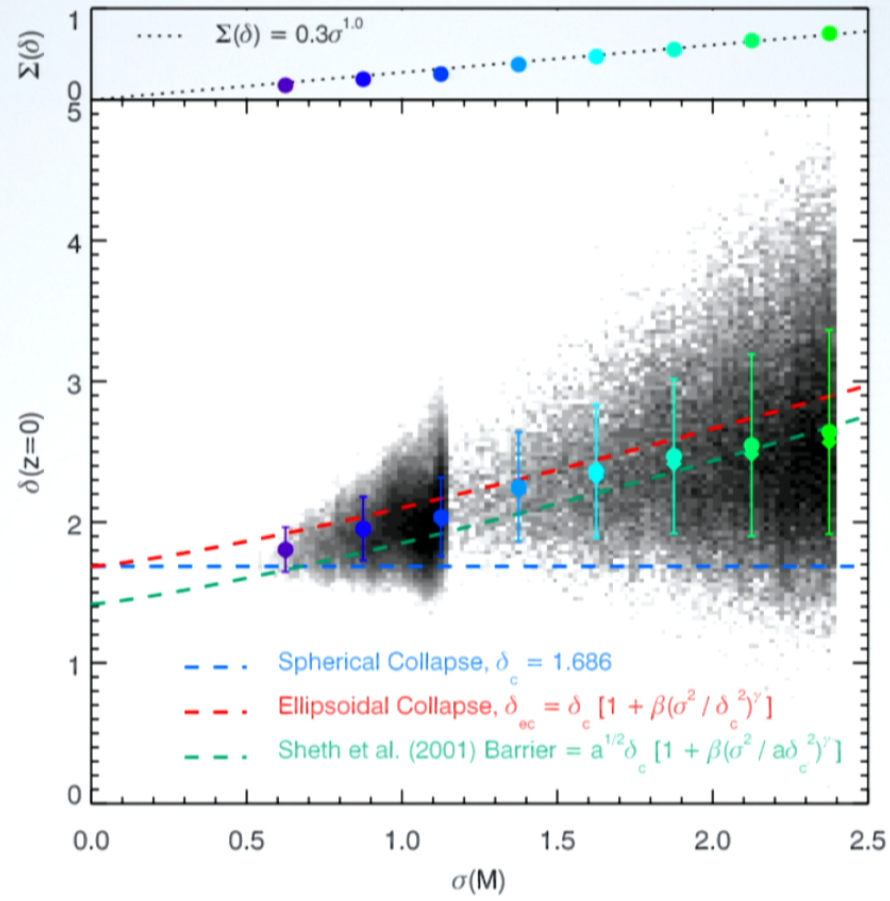
HALO MASS FUNCTION (I)

- The first (up)crossing distribution is

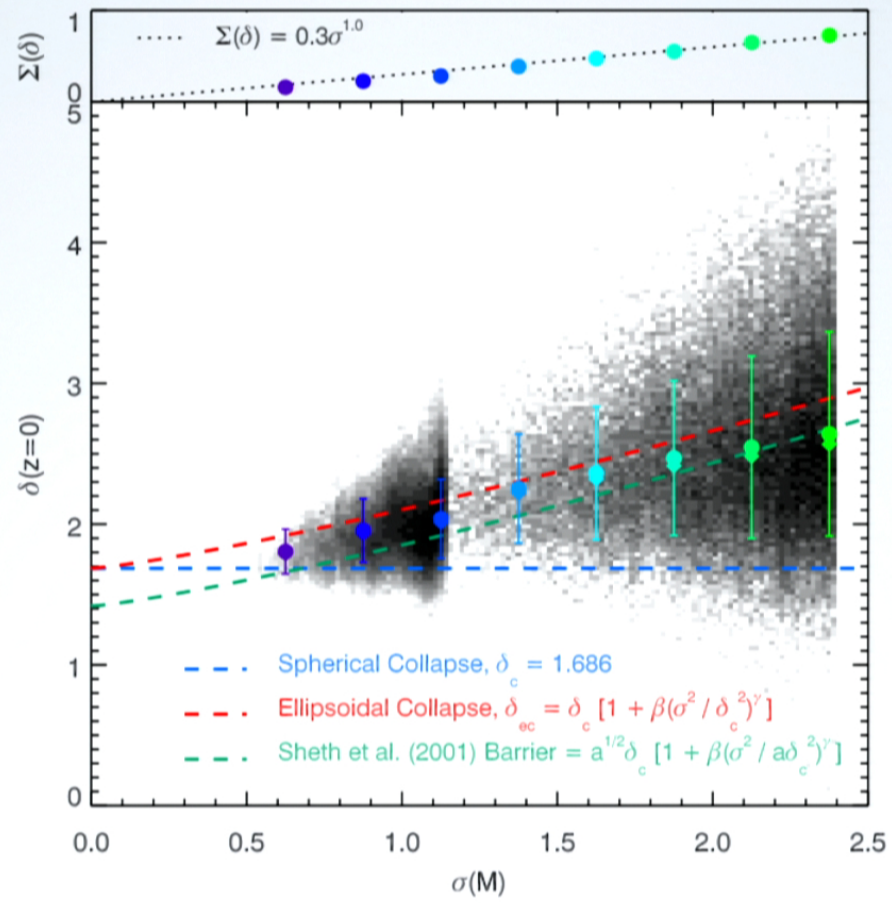
$$f_{\text{ESP}}(\nu_c) = \frac{M}{\bar{\rho}} \bar{n}_{\text{UC}}(\nu_c, R_s) \frac{dR_s}{d\nu_c}$$

and the halo abundance

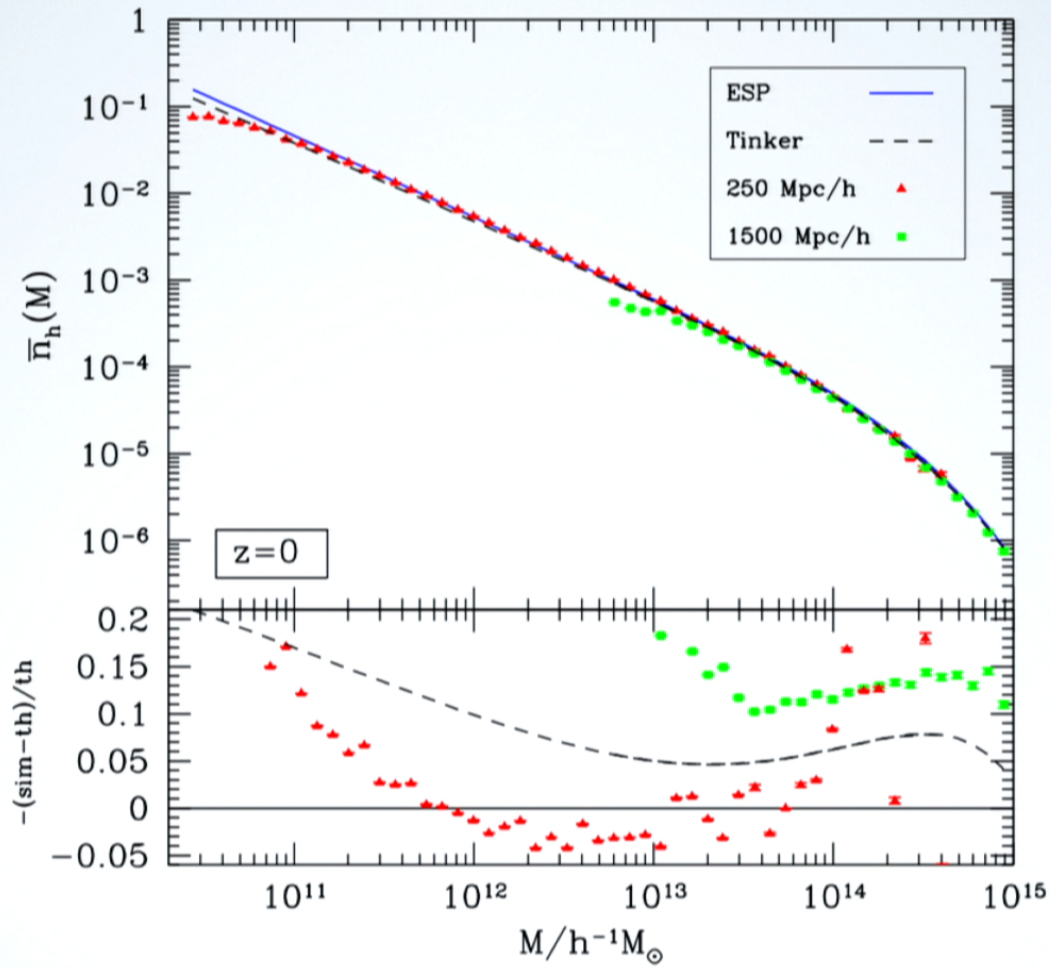
$$\bar{n}_{\text{h}}(M) = \frac{\bar{\rho}}{M^2} \nu_c f_{\text{ESP}}(\nu_c, R_s) \frac{d \ln \nu_c}{d \ln M}$$



Robertson et. al. (2009)



Robertson et. al. (2009)



Paranjape, Sheth, VD (2013), Biagetti et. al. (in prep.)