Title: Feature-finding in the cosmic microwave background

Date: Feb 09, 2012 11:40 AM

URL: http://pirsa.org/12020144

Abstract:

Pirsa: 12020144 Page 1/47

Feature finding in the CMB

Matthew C. Johnson Perimeter Institute

In collaboration with:

Stephen Feeney Daniel Mortlock Hiranya Peiris

Pirsa: 12020144 Page 2/47

The data Lambda-CDM: very successful at describing the CMB power spectrum. WMAP 7-year data 100 500 Multipole moment / Is there room for anything else?

Pirsa: 12020144 Page 3/47

The data Lambda-CDM: very successful at describing the CMB power spectrum. WMAP 7-year data 100 500 Multipole moment I Is there room for anything else?

Pirsa: 12020144 Page 4/47

Theories with features

 A number of theories predict specific features in the CMB on top of those predicted by LCDM:

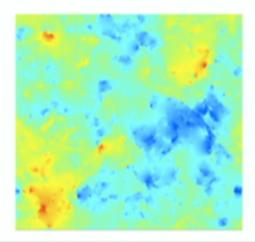
$$\frac{\Delta T(\hat{\mathbf{n}})}{T} = (1 + f(\hat{\mathbf{n}})) (1 + \delta_{\text{LCDM}}(\hat{\mathbf{n}})) - 1$$
$$\simeq f(\hat{\mathbf{n}}) + \delta_{\text{LCDM}}(\hat{\mathbf{n}})$$

Theories with features

A number of theories predict specific features in the CMB on top of those predicted by LCDM:

$$\frac{\Delta T(\hat{\mathbf{n}})}{T} = (1 + f(\hat{\mathbf{n}})) (1 + \delta_{\text{LCDM}}(\hat{\mathbf{n}})) - 1$$
$$\simeq f(\hat{\mathbf{n}}) + \delta_{\text{LCDM}}(\hat{\mathbf{n}})$$

Topological defects: cosmic strings



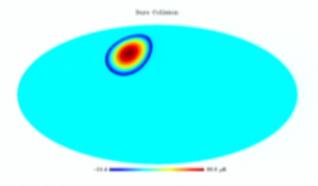
Pirsa: 12020144 Page 6/47

Theories with features

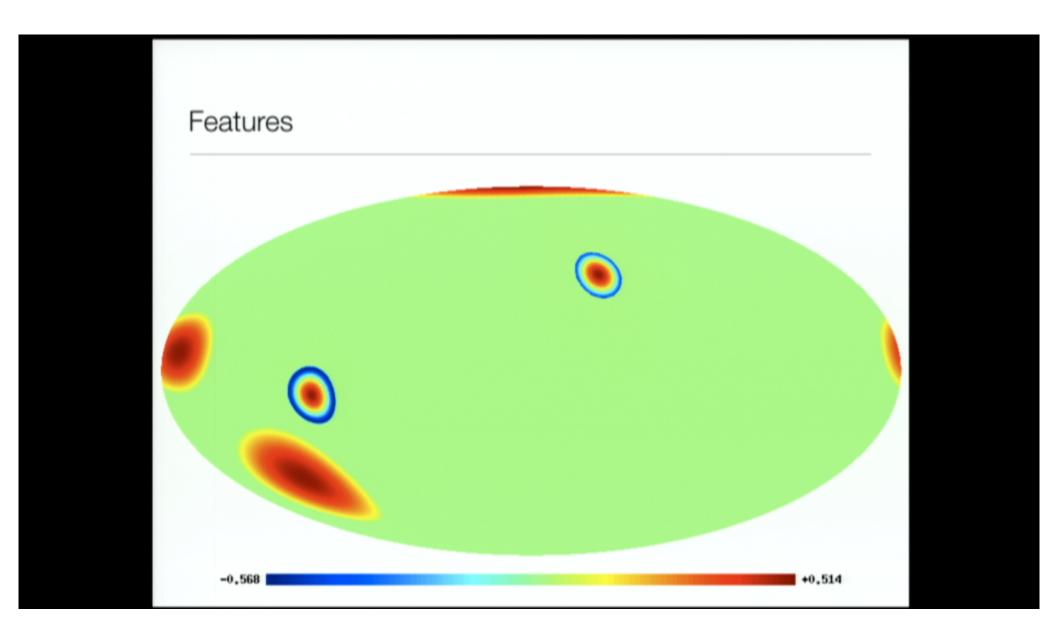
A number of theories predict specific features in the CMB on top of those predicted by LCDM:

$$\frac{\Delta T(\hat{\mathbf{n}})}{T} = (1 + f(\hat{\mathbf{n}})) (1 + \delta_{\text{LCDM}}(\hat{\mathbf{n}})) - 1$$
$$\simeq f(\hat{\mathbf{n}}) + \delta_{\text{LCDM}}(\hat{\mathbf{n}})$$

- Topological defects: cosmic strings and cosmic textures.
- Cosmic bubble collisions: observational signature of eternal inflation.



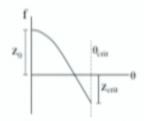
Pirsa: 12020144 Page 7/47



Pirsa: 12020144 Page 8/47

A theory that predicts features:

Template with free parameters **m** size, location, amplitude....



PDF for the number of features and template parameters

$$\Pr(\bar{N}_s, \mathbf{m})$$

s, m)

How many features of each type do I expect to see?

Pirsa: 12020144

Bayes' theorem:

$$\Pr(\bar{N}_{s}|\boldsymbol{d}) = \frac{\Pr(\bar{N}_{s})\Pr(\boldsymbol{d}|\bar{N}_{s})}{\Pr(\boldsymbol{d})}$$

Pirsa: 12020144 Page 10/47

Bayes' theorem:

$$\Pr(\bar{N}_{s}|\boldsymbol{d}) = \frac{\Pr(\bar{N}_{s})\Pr(\boldsymbol{d}|\bar{N}_{s})}{\Pr(\boldsymbol{d})}$$

• The actual number of features is drawn from a Poisson distribution:

$$\Pr(\boldsymbol{d}|\bar{N}_{\mathrm{s}}) = \sum_{N_{\mathrm{s}}=0}^{\infty} \Pr(N_{\mathrm{s}}|\bar{N}_{\mathrm{s}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}})$$

$$\Pr(N_{\mathrm{s}}|ar{N}_{\mathrm{s}}) = rac{ar{N}_{\mathrm{s}}^{N_{\mathrm{n}}}e^{-ar{N}_{\mathrm{n}}}}{N_{\mathrm{s}}!}$$

Pirsa: 12020144 Page 11/47

Bayes' theorem:

$$\Pr(\bar{N}_{s}|\boldsymbol{d}) = \frac{\Pr(\bar{N}_{s})\Pr(\boldsymbol{d}|\bar{N}_{s})}{\Pr(\boldsymbol{d})}$$

• The actual number of features is drawn from a Poisson distribution:

$$\Pr(\boldsymbol{d}|\bar{N}_{\mathrm{s}}) = \sum_{N_{\mathrm{s}}=0}^{\infty} \Pr(N_{\mathrm{s}}|\bar{N}_{\mathrm{s}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}})$$

$$\Pr(N_{\mathrm{s}}|ar{N}_{\mathrm{s}}) = rac{ar{N}_{\mathrm{s}}^{N_{\mathrm{s}}}e^{-ar{N}_{\mathrm{s}}}}{N_{\mathrm{s}}!}$$

ullet The likelihood for N_s templates is given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}}) = \int \mathrm{d}\boldsymbol{m}_1 \, \mathrm{d}\boldsymbol{m}_2 \dots \mathrm{d}\boldsymbol{m}_{N_{\mathrm{s}}} \Pr(\boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}}, \boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}})$$

We've marginalized over template parameters

$$\boldsymbol{m}_i = \{\theta_0^{(i)}, \phi_0^{(i)}, \theta_{\text{crit}}^{(i)}, z_0^{(i)}, z_{\text{crit}}^{(i)}\}$$

The likelihoods are given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}},\boldsymbol{m}_{1},\boldsymbol{m}_{2},\ldots\boldsymbol{m}_{N_{\mathrm{s}}}) = \frac{1}{(2\pi)^{N_{\mathrm{px}}/2}|\mathsf{C}|} \exp\left(-\frac{1}{2}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]\mathsf{C}^{-1}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]^{\mathrm{T}}\right)$$

d vector containing the data at each pixel

 $\mathbf{t}(\mathbf{m}_i)$ vector containing the value of the template at each pixel

C covariance matrix

Neglecting instrumental noise:

$$C_{ij} = \sum_{\ell} \frac{2\ell + 1}{4\pi} C_{\ell} P_{\ell}(\cos \theta_{ij})$$

The likelihoods are given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}},\boldsymbol{m}_{1},\boldsymbol{m}_{2},\ldots\boldsymbol{m}_{N_{\mathrm{s}}}) = \frac{1}{(2\pi)^{N_{\mathrm{px}}/2}|\mathsf{C}|} \exp\left(-\frac{1}{2}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]\mathsf{C}^{-1}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]^{\mathrm{T}}\right)$$

d vector containing the data at each pixel

 $\mathbf{t}(\mathbf{m}_i)$ vector containing the value of the template at each pixel

C covariance matrix

Neglecting instrumental noise:

$$C_{ij} = \sum_{\ell} \frac{2\ell + 1}{4\pi} C_{\ell} P_{\ell}(\cos \theta_{ij})$$

Bayes' theorem:

$$\Pr(\bar{N}_{\mathrm{s}}|\boldsymbol{d}) = \frac{\Pr(\bar{N}_{\mathrm{s}})\Pr(\boldsymbol{d}|\bar{N}_{\mathrm{s}})}{\Pr(\boldsymbol{d})}$$

• The actual number of features is drawn from a Poisson distribution:

$$\Pr(\boldsymbol{d}|\bar{N}_{\mathrm{s}}) = \sum_{N_{\mathrm{s}}=0}^{\infty} \Pr(N_{\mathrm{s}}|\bar{N}_{\mathrm{s}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}})$$

$$\Pr(N_{\mathrm{s}}|ar{N}_{\mathrm{s}}) = rac{ar{N}_{\mathrm{s}}^{N_{\mathrm{s}}}e^{-ar{N}_{\mathrm{s}}}}{N_{\mathrm{s}}!}$$

ullet The likelihood for N_s templates is given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}}) = \int \mathrm{d}\boldsymbol{m}_1 \, \mathrm{d}\boldsymbol{m}_2 \dots \mathrm{d}\boldsymbol{m}_{N_{\mathrm{s}}} \Pr(\boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}}, \boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}})$$

We've marginalized over template parameters

$$\boldsymbol{m}_i = \{\theta_0^{(i)}, \phi_0^{(i)}, \theta_{\mathrm{crit}}^{(i)}, z_0^{(i)}, z_{\mathrm{crit}}^{(i)}\}$$

Bayes' theorem:

$$\Pr(\bar{N}_{s}|\boldsymbol{d}) = \frac{\Pr(\bar{N}_{s})\Pr(\boldsymbol{d}|\bar{N}_{s})}{\Pr(\boldsymbol{d})}$$

• The actual number of features is drawn from a Poisson distribution:

$$\Pr(\boldsymbol{d}|\bar{N}_{\mathrm{s}}) = \sum_{N_{\mathrm{s}}=0}^{\infty} \Pr(N_{\mathrm{s}}|\bar{N}_{\mathrm{s}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}})$$

$$\Pr(N_{\mathrm{s}}|ar{N}_{\mathrm{s}}) = rac{ar{N}_{\mathrm{s}}^{N_{\mathrm{s}}}e^{-ar{N}_{\mathrm{s}}}}{N_{\mathrm{s}}!}$$

ullet The likelihood for N_s templates is given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}}) = \int \mathrm{d}\boldsymbol{m}_1 \, \mathrm{d}\boldsymbol{m}_2 \dots \mathrm{d}\boldsymbol{m}_{N_{\mathrm{s}}} \Pr(\boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}}) \Pr(\boldsymbol{d}|N_{\mathrm{s}}, \boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_{\mathrm{s}}})$$

We've marginalized over template parameters

$$\boldsymbol{m}_i = \{\theta_0^{(i)}, \phi_0^{(i)}, \theta_{\text{crit}}^{(i)}, z_0^{(i)}, z_{\text{crit}}^{(i)}\}$$

The likelihoods are given by:

$$\Pr(\boldsymbol{d}|N_{\mathrm{s}},\boldsymbol{m}_{1},\boldsymbol{m}_{2},\ldots\boldsymbol{m}_{N_{\mathrm{s}}}) = \frac{1}{(2\pi)^{N_{\mathrm{px}}/2}|\mathsf{C}|} \exp\left(-\frac{1}{2}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]\mathsf{C}^{-1}\left[\boldsymbol{d} - \sum_{i=1}^{N_{\mathrm{s}}}\boldsymbol{t}(\boldsymbol{m}_{i})\right]^{\mathrm{T}}\right)$$

d vector containing the data at each pixel

 $\mathbf{t}(\mathbf{m}_i)$ vector containing the value of the template at each pixel

C covariance matrix

Neglecting instrumental noise:

$$C_{ij} = \sum_{\ell} \frac{2\ell + 1}{4\pi} C_{\ell} P_{\ell}(\cos \theta_{ij})$$

- Problems:
 - Covariance matrix is very large and non-diagonal, hard to invert!
 - Need to sample a many-dimensional parameter space!
 - Arbitrary number of templates!

Pirsa: 12020144 Page 18/47

The likelihoods are given by:

$$\Pr(\boldsymbol{d}|N_s, \boldsymbol{m}_1, \boldsymbol{m}_2, \dots \boldsymbol{m}_{N_s}) = \frac{1}{(2\pi)^{N_{\mathrm{px}}/2}|\mathsf{C}|} \exp\left(-\frac{1}{2}\left[\boldsymbol{d} - \sum_{i=1}^{N_s} \boldsymbol{t}(\boldsymbol{m}_i)\right] \mathsf{C}^{-1} \left[\boldsymbol{d} - \sum_{i=1}^{N_s} \boldsymbol{t}(\boldsymbol{m}_i)\right]^{\mathrm{T}}\right)$$

d vector containing the data at each pixel

 $\mathbf{t}(\mathbf{m}_i)$ vector containing the value of the template at each pixel

C covariance matrix

Neglecting instrumental noise:

$$C_{ij} = \sum_{\ell} \frac{2\ell + 1}{4\pi} C_{\ell} P_{\ell}(\cos \theta_{ij})$$

Problems:

- Covariance matrix is very large and non-diagonal, hard to invert!
- Need to sample a many-dimensional parameter space!
- Arbitrary number of templates!

Solution:

- Locate candidate features with a blind analysis.
- Find an approximation to the posterior by integrating only over the regions of parameter space where the likelihood is large.

Pirsa: 12020144 Page 20/47

Problems:

- Covariance matrix is very large and non-diagonal, hard to invert!
- Need to sample a many-dimensional parameter space!
- Arbitrary number of templates!

Solution:

- Locate candidate features with a blind analysis.
- Find an approximation to the posterior by integrating only over the regions of parameter space where the likelihood is large.



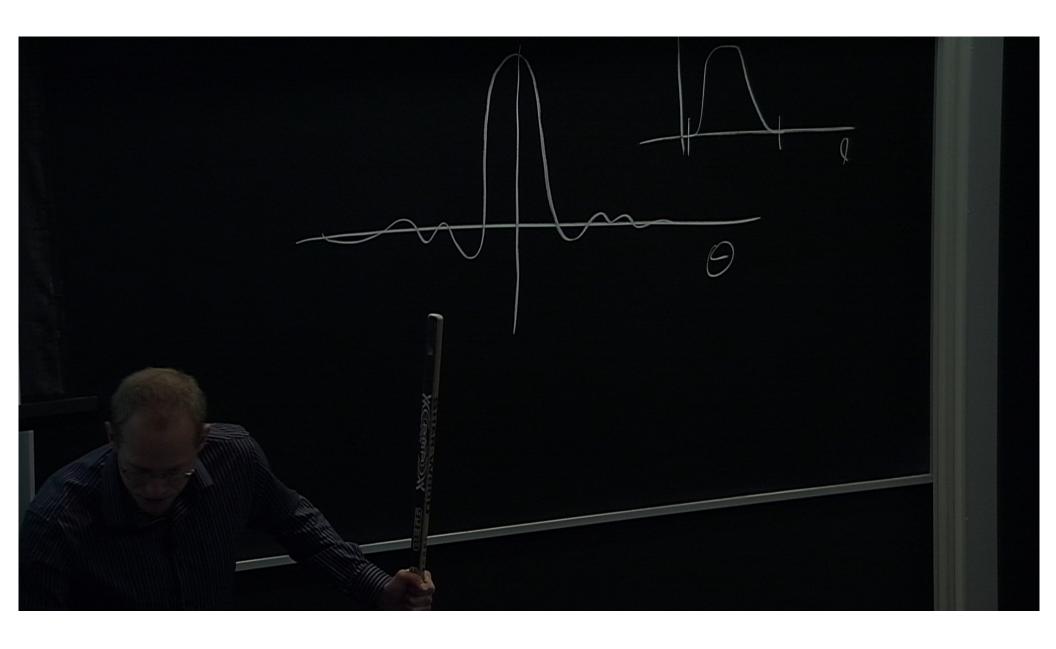
Pirsa: 12020144 Page 21/47

 The current pipeline uses needlets (a type of wavelet transform) to identify candidates: sensitive to scale (j) and location (k).

$$T(\hat{\gamma}) = \sum_{j,k} \beta_{jk} \psi_{jk}(\hat{\gamma}) \qquad \beta_{jk} = \int T(\hat{\gamma}) \psi_{jk}(\hat{\gamma}) d\Omega.$$

The coefficients are computed in harmonic space.

Pirsa: 12020144 Page 22/47

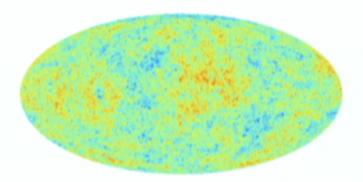


Pirsa: 12020144 Page 23/47

 The current pipeline uses needlets (a type of wavelet transform) to identify candidates: sensitive to scale (j) and location (k).

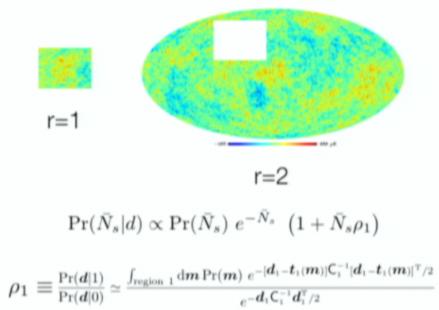
$$T(\hat{\gamma}) = \sum_{j,k} \beta_{jk} \psi_{jk}(\hat{\gamma}) \qquad \beta_{jk} = \int T(\hat{\gamma}) \psi_{jk}(\hat{\gamma}) d\Omega.$$

The coefficients are computed in harmonic space.



Pirsa: 12020144 Page 24/47

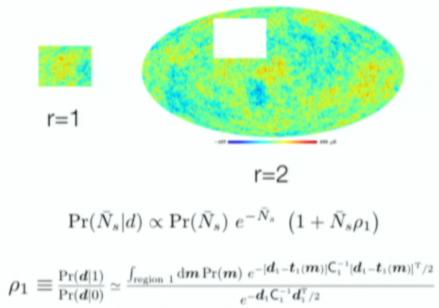
For one candidate:



Evidence ratio in the patch: how much better does one describe the data by adding a template

Pirsa: 12020144 Page 25/47

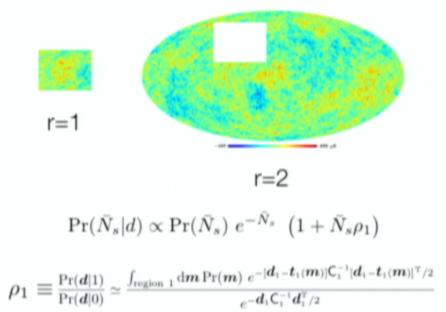
For one candidate:



Evidence ratio in the patch: how much better does one describe the data by adding a template

Pirsa: 12020144 Page 26/47

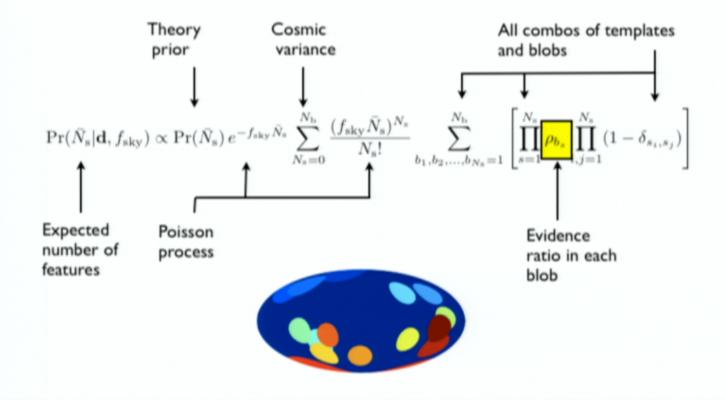
For one candidate:



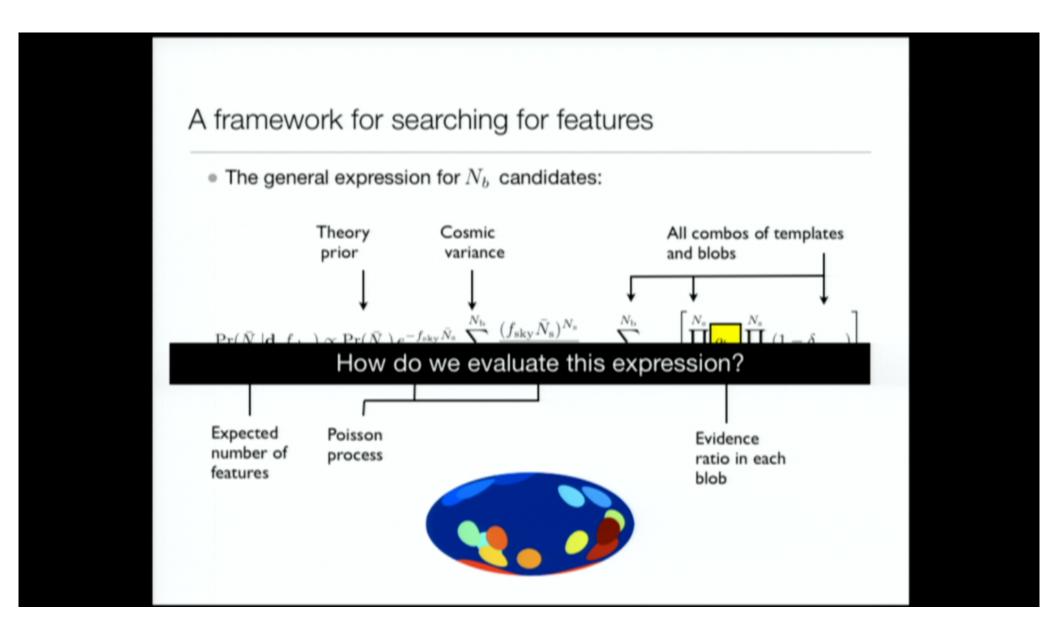
Evidence ratio in the patch: how much better does one describe the data by adding a template

Pirsa: 12020144 Page 27/47

ullet The general expression for N_b candidates:



Pirsa: 12020144 Page 28/47



Pirsa: 12020144 Page 29/47

- Data is publicly available from the LAMBDA web site.
- Full sky data is available at 5 frequencies: useful to remove foregrounds.
- The data is represented using the Healpix pixelization scheme:



Pirsa: 12020144 Page 30/47

- Data is publicly available from the LAMBDA web site.
- Full sky data is available at 5 frequencies: useful to remove foregrounds.
- The data is represented using the Healpix pixelization scheme:



Hierarchical, Equal Area, isoLatitude Pixelization.

$$N_{\rm pix} = 12 \ N_{\rm side}^2 \qquad N_{\rm side} = 1, 2, \dots$$

 \bullet WMAP data: $N_{\rm side}=512$, $\,N_{\rm pix}=3,145,728$.

Nyquist: 2 pixels per beam width (.2 degrees)

- Data is publicly available from the LAMBDA web site.
- Full sky data is available at 5 frequencies: useful to remove foregrounds.
- The data is represented using the Healpix pixelization scheme:



Hierarchical, Equal Area, isoLatitude Pixelization.

$$N_{\rm pix} = 12 \ N_{\rm side}^2 \qquad N_{\rm side} = 1, 2, \dots$$

 \bullet WMAP data: $N_{\rm side}=512$, $\,N_{\rm pix}=3,145,728$.

Nyquist: 2 pixels per beam width (.2 degrees)

- HEALPix software: public tools for the analysis and simulation of data.
- Fortran 90, C++, IDL...
- Parallelized.
 - Pixelization allows fast and accurate spherical harmonic transform.
 - Subroutines to manipulate data, perform forward and inverse spherical harmonic transforms, perform file I/O....
 - Facilities to make simulated maps, find power spectra, visualize data....

Pirsa: 12020144 Page 33/47

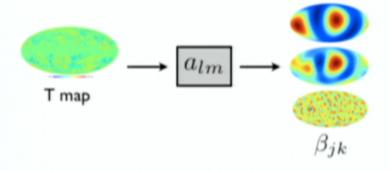
- HEALPix software: public tools for the analysis and simulation of data.
- Fortran 90, C++, IDL...
- Parallelized.
 - Pixelization allows fast and accurate spherical harmonic transform.
 - Subroutines to manipulate data, perform forward and inverse spherical harmonic transforms, perform file I/O....
 - Facilities to make simulated maps, find power spectra, visualize data....

Pirsa: 12020144 Page 34/47

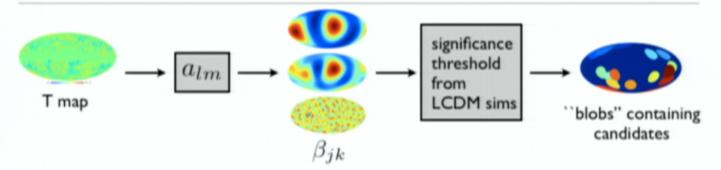


T map

Pirsa: 12020144 Page 35/47



Pirsa: 12020144 Page 36/47



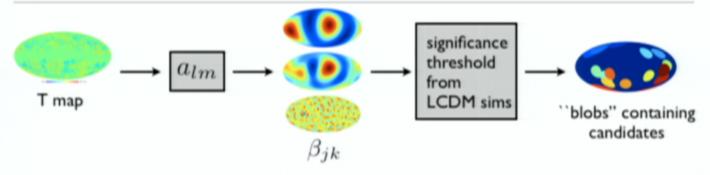
For each blob:



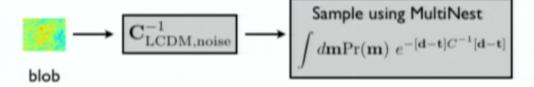
blob

Pirsa: 12020144 Page 37/47

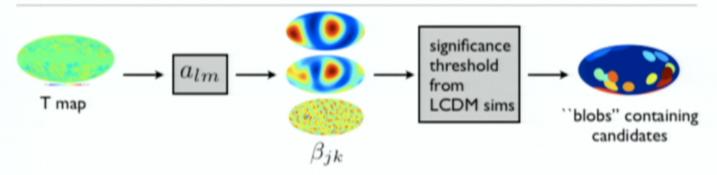




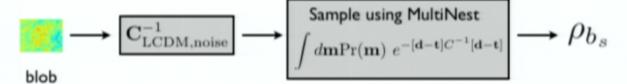
For each blob:



Pirsa: 12020144 Page 38/47



For each blob:



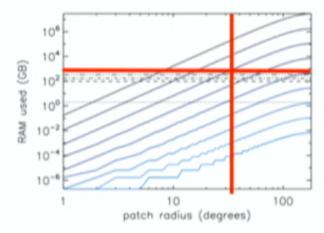
Combine according to:

$$\Pr(\bar{N}_{\rm s}|\mathbf{d}, f_{\rm sky}) \propto \Pr(\bar{N}_{\rm s}) \, e^{-f_{\rm sky}\bar{N}_{\rm s}} \sum_{N_{\rm s}=0}^{N_{\rm b}} \frac{(f_{\rm sky}\bar{N}_{\rm s})^{N_{\rm s}}}{N_{\rm s}!} \sum_{b_1, b_2, \dots, b_{N_{\rm s}}=1}^{N_{\rm b}} \left[\prod_{s=1}^{N_{\rm s}} \rho_{b_s} \prod_{i,j=1}^{N_{\rm s}} (1 - \delta_{s_i, s_j}) \right]$$

Pirsa: 12020144 Page 39/47

Limitations

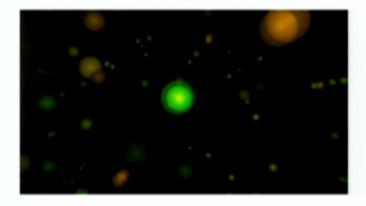
- Each blob contains a huge amount of data.
- Storing and inverting the covariance matrix requires lots of RAM:



- Currently running on a 672GB shared memory machine.
- The code is parallelized.
- The resolution inside a blob is degraded according to the resources available.

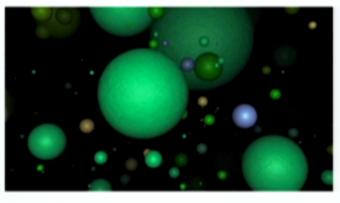
Pirsa: 12020144 Page 40/47

Results: Cosmic bubble collisions

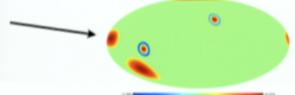


Pirsa: 12020144 Page 41/47

Results: Cosmic bubble collisions

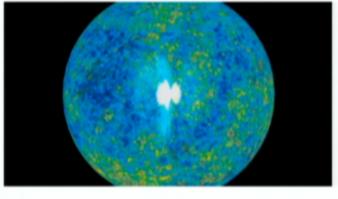


 The collision of our bubble with others provides an observational test of eternal inflation.

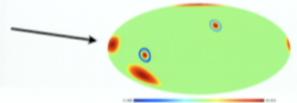


Pirsa: 12020144 Page 42/47

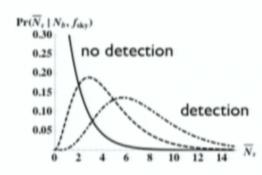
Results: Cosmic bubble collisions



 The collision of our bubble with others provides an observational test of eternal inflation.

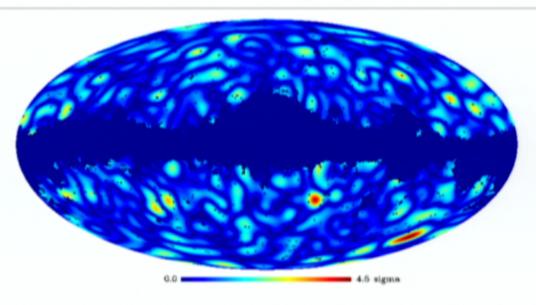


Some possible outcomes:



Pirsa: 12020144 Page 43/47

WMAP7 W-Band (94 GHz): Needlets

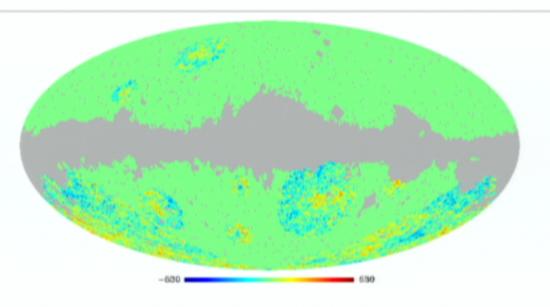


(sensitive to 5-14 degree templates)

11 features pass thresholds, with detections in multiple needlet types/frequencies

Pirsa: 12020144 Page 44/47





Divide the sky into "blobs" containing each candidate collision.

Pirsa: 12020144 Page 45/47

WMAP7 W-Band (94 GHz): Posterior

ullet The posterior is peaked around $ar{N}_s=0$

The data does not support the bubble collision hypothesis.

From the shape of the posterior, we can rule out

$$\bar{N}_{s} < 1.6 \text{ at } 68\% \text{ CL}$$

Pirsa: 12020144 Page 46/47

WMAP7 W-Band (94 GHz): Posterior

ullet The posterior is peaked around $ar{N}_s=0$

The data does not support the bubble collision hypothesis.

From the shape of the posterior, we can rule out

$$\bar{N}_{s} < 1.6 \text{ at } 68\% \text{ CL}$$

Pirsa: 12020144 Page 47/47