Title: Cosmological Model Selection and the Inflationary Cosmology

Date: Oct 24, 2005 03:00 PM

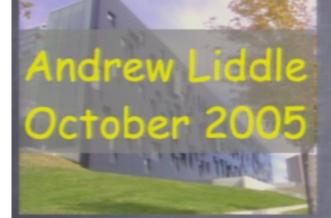
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Abstract:

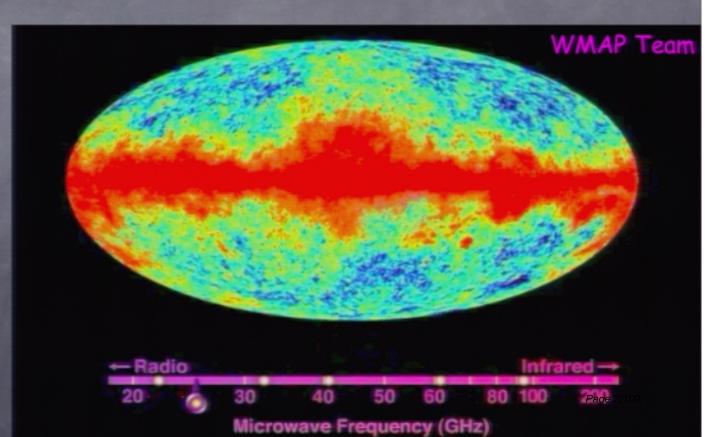
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Cosmological model selection and inflation



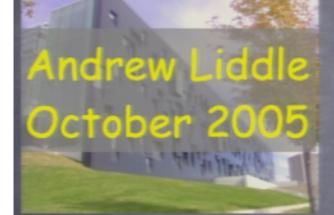




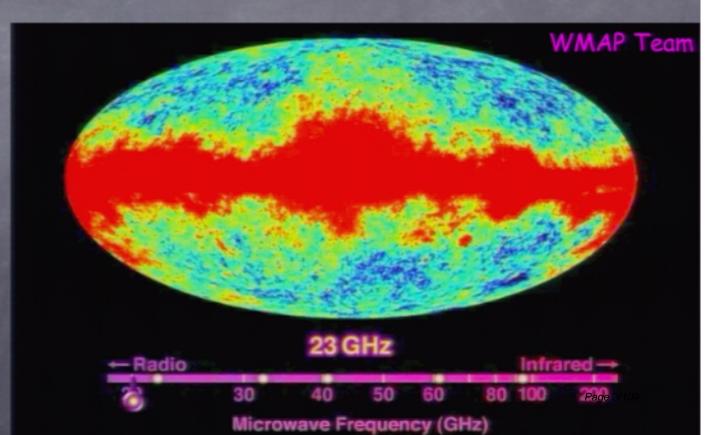


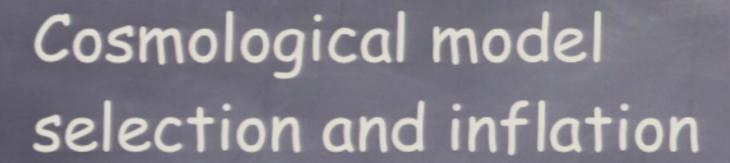
Cosmological model selection and inflation







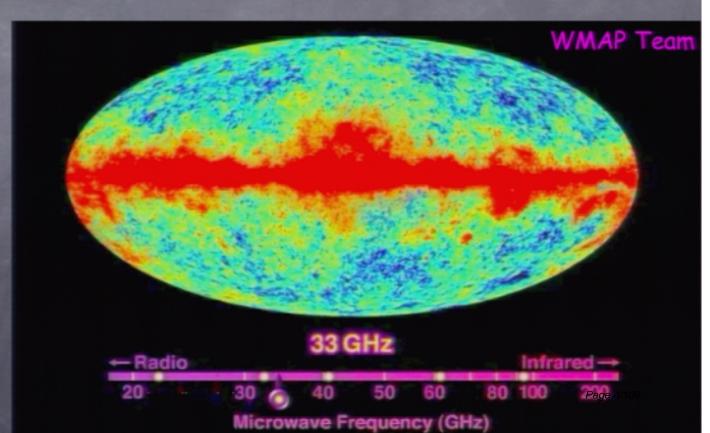






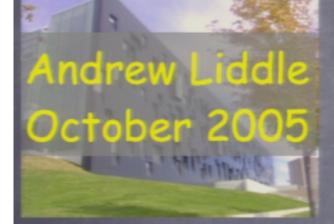




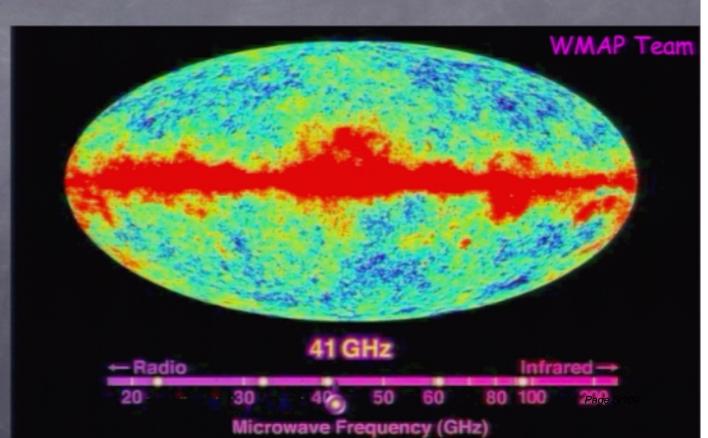


Cosmological model selection and inflation



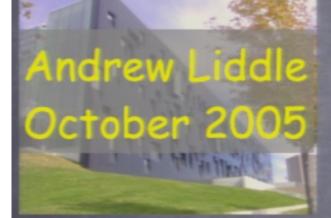




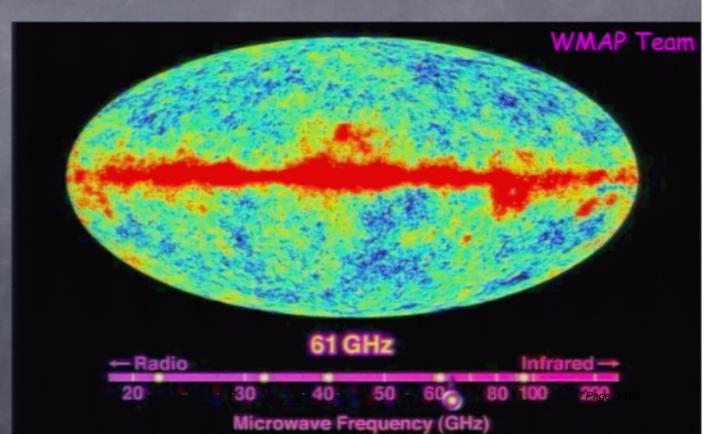


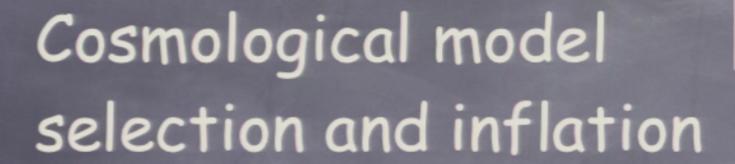
Cosmological model selection and inflation



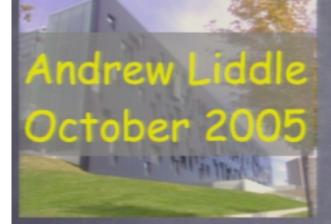




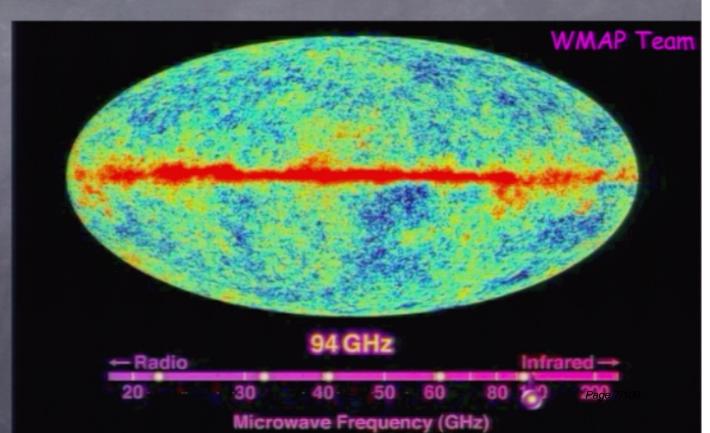






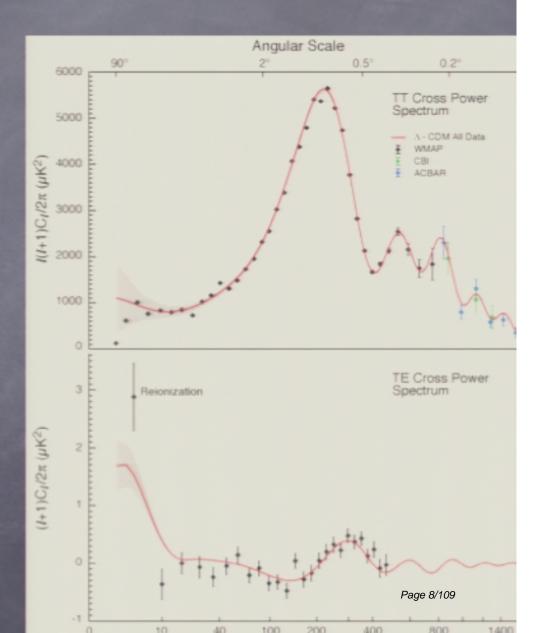






Conclusions from WMAP:

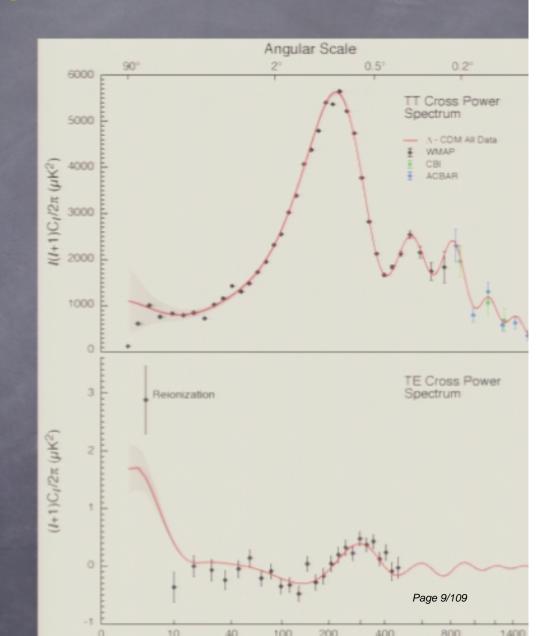
If you want to explain this data, the simplest way is ...



Conclusions from WMAP:

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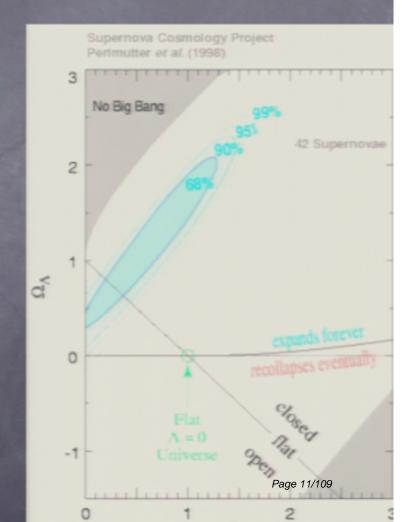
- A spatially-flat Universe
- Dark matter and dark energy
- Initial perturbations which are gaussian, adiabatic and nearly scale-invariant, e.g. as given by inflation.



Inflation is any period of the Universe's evolution during which the Universe is accelerating

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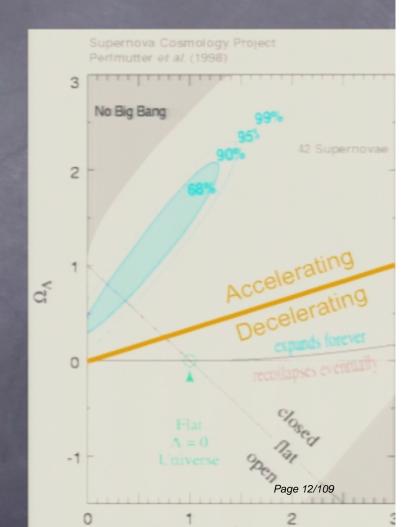
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This can also be written in terms of the comoving Hubble length as

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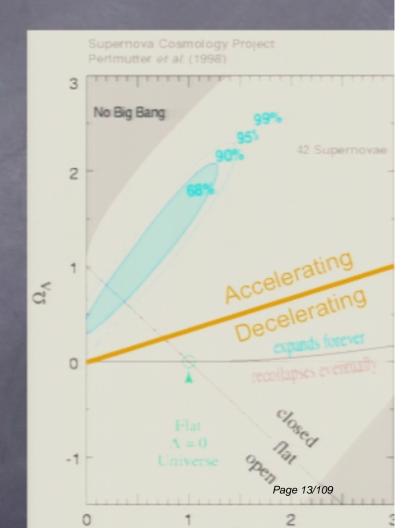


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Early Universe inflation is the most plausible exlanation we have from the origin of structure.



Predictions of the simplest models

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The simplest models of inflation predict nearly scale-invariant spectra of adiabatic gaussian density perturbations and gravitational waves, in their growing mode, in a spatially-flat Universe.

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The simplest models of inflation predict nearly scale-invariant spectra of adiabatic gaussian density perturbations and gravitational waves, in their growing mode, in a spatially-flat Universe.

WMAP does not provide any evidence against any of these, and gives support to all but the gravitational waves. As such, it gives strong general support to the sinflationary paradigm (but not uniquely to inflation).

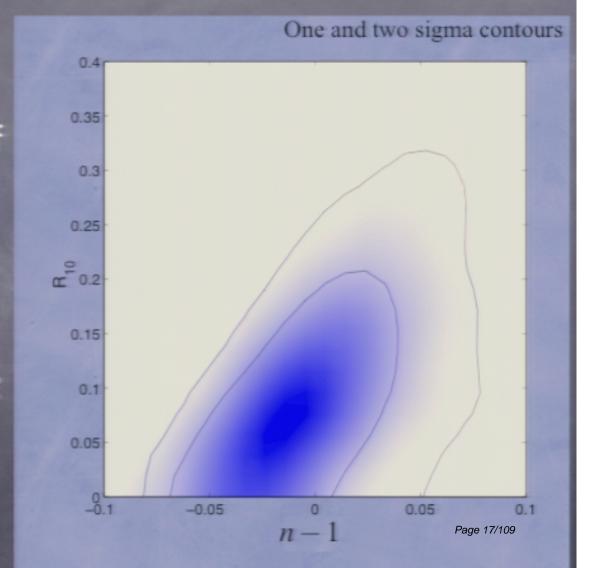
Current constraints

Comparison with observations:

- Fit to data compilation of WMAP, other CMB experiments (VSA, CBI and ACBAR), and 2dF galaxy survey.
- Use CAMB plus CosmoMC plus WMAP likelihood code plus slow-roll

 Pirsa: 0510041 lation module.

Leach & Liddle, PRD, astro-ph/0306305

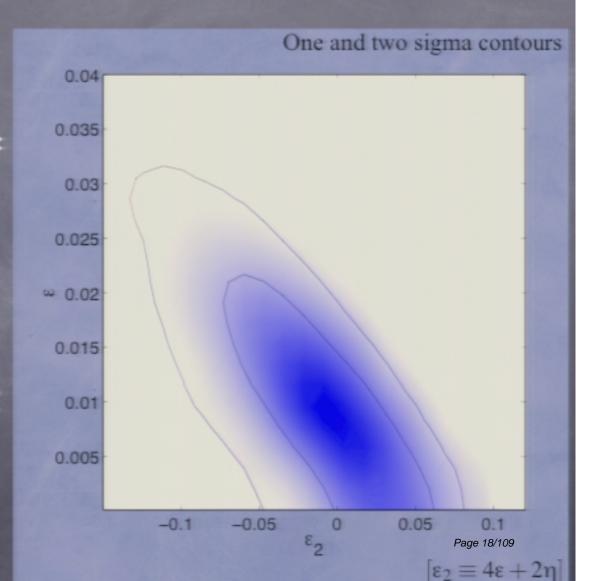


Current status of single-field inflation models

Leach & Liddle, PRD, astro-ph/0306305

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No unambiguous evidence of primordial non-gaussianity.

Gravitational waves

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Good fit to data assuming no decaying mode. Temperature-polarization anti-correlation.

 $\Omega_{\text{tot}} = 1.02 \pm 0.02$

Spatial flatness

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Even if effects from these more complex models are never seen,

Pitch 5000/41 introduce degeneracies in interpretting observations. Page 29/109

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The precise constraints obtained depend on

- The observational datasets used.
- The set of cosmological parameters used to define the cosmological model.

There have been a variety of choices made for both of these.

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Table 7. Best Fit Parameters: Power Law Λ CDM

	WMAP	WMAPext ¹⁶ a	WMAPext+2dFGRS	WMAPext+ 2dFGRS+ Lyman α
A	0.9 ± 0.1	0.8 ± 0.1	0.8 ± 0.1	0.75+0.08
n_s	0.99 ± 0.04	0.97 ± 0.03	0.97 ± 0.03	0.96 ± 0.02
τ	$0.166^{+0.076}_{-0.071}$	$0.143^{+0.071}_{-0.062}$	$0.148^{+0.073}_{-0.071}$	0.117+0.057
h	0.72 ± 0.05	0.73 ± 0.05	0.73 ± 0.03	0.72 ± 0.03
$\Omega_m h^2$	0.14 ± 0.02	0.13 ± 0.01	0.134 ± 0.006	0.133 ± 0.006
$\Omega_b h^2$	0.024 ± 0.001	0.023 ± 0.001	0.023 ± 0.001	0.0226 ± 0.0008
χ^2_{eff}/ν	1429/1341	1440/1352	1468/1381	Ь

WMAP: Spergel et al

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Table 8. Best Fit Parameters for the Running Spectral Index ACDM Model

	WMAP	WMAPext	WMAPext+2dFGRS	WMAPext+ 2dFGRS+ Lyman α
A n_s $dn_s/d\ln k$ τ	0.92 ± 0.12 $0.93^{+0.07}_{-0.07}$ -0.047 ± 0.04 0.20 ± 0.07	0.9 ± 0.1 0.91 ± 0.06 -0.055 ± 0.038 0.20 ± 0.07	0.84 ± 0.09 $0.93^{+0.04}_{-0.05}$ $-0.031^{+0.023}_{-0.025}$ 0.17 ± 0.06	$0.83^{+0.09}_{-0.08}$ 0.93 ± 0.03 $-0.031^{+0.016}_{-0.017}$ 0.17 ± 0.06
a; 05100041	0.70 ± 0.05 0.14 ± 0.02 0.023 ± 0.002 1431/1342	0.71 ± 0.06 0.14 ± 0.01 0.022 ± 0.001 1437/1350	0.71 ± 0.04 0.136 ± 0.009 0.022 ± 0.001 1465/1380	0.71 ^{+0.04} 0.135 ^{+0.008} 0.0224 ± 0.0009 **

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Table 8. Best Fit Parameters for the Running Spectral Index Λ CDM Model

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A	0.92 ± 0.12 $0.93^{+0.07}_{-0.07}$	0.9 ± 0.1 0.91 ± 0.06	0.84 ± 0.09 0.93+0.04 0.95	$0.83^{+0.09}_{-0.08}$ 0.93 ± 0.03
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a: 05100041	1431/1342	1437/1350	1465/1380	Page 39/

Parameter Estimation

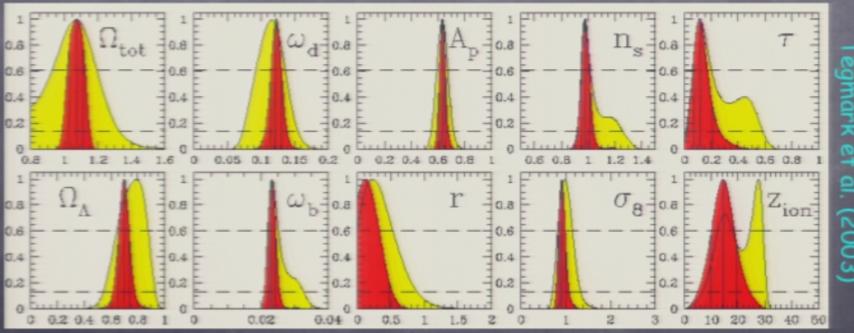
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al. (2003)

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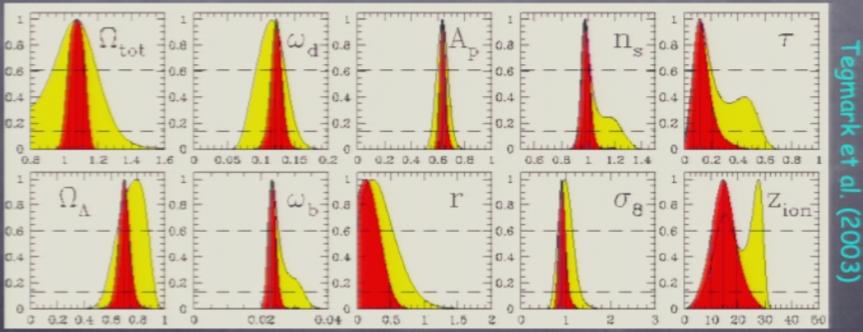
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The maximum likelihood gives the best values for the parameters, Page 42/109 and the neighbouring behaviour gives the confidence limits.

Model Selection

In model selection, we aim to distinguish different cosmological models, meaning different choices of the parameters to be varied. In particular we need to allow for model dimensionality: that different models may have different numbers of parameters.

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A suitable baseline cosmological model to consider is the simplest one giving an adequate fit to current data. It is a spatially-flat adiabatic ΛCDM model with five fundamental parameters and two phenomenological ones.

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Ω_{m}	matter density
$\Omega_{\rm b}$	baryon density
$\Omega_{\rm r}$	radiation density
h	hubble parameter
A	adiabatic density perturbation amplitude
τ	reionization optical depth
h	hias parameter (or parameters)

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There are many, many ways in which this base cosmological model can be extended.

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Table 2. Candidate parameters: those which might be relevant for cosmological observations, but for which there is presently no convincing evidence requiring them. They are listed so as to take the value zero in the base cosmological model. Those above the line are parameters of the background homogeneous cosmology, and those below describe the perturbations. Of the latter set, the first six refer to adiabatic perturbations, the next three to tensor perturbations, and the remainder to isocurvature perturbations.

Ω_k	spatial curvature
$N_{\nu} - 3.04$	effective number of neutrino species (CMBFAST definition)
m_{ν_i}	neutrino mass for species 'i'
	[or more complex neutrino properties]
$m_{ m dm}$	(warm) dark matter mass
w+1	dark energy equation of state
dw/dz	redshift dependence of w
	[or more complex parametrization of dark energy evolution]
$c_{S}^{2}-1$	effects of dark energy sound speed
$1/r_{\rm top}$	topological identification scale
	[or more complex parametrization of non-trivial topology]
$d\alpha/dz$	redshift dependence of the fine structure constant
dG/dz	redshift dependence of the gravitational constant
n-1	scalar spectral index
$dn/d \ln k$	running of the scalar spectral index
$k_{\rm cut}$	large-scale cut-off in the spectrum
Afeature	amplitude of spectral feature (peak, dip or step)
k_{feature}	and its scale
	[or adiabatic power spectrum amplitude parametrized in N bins]
fnl	quadratic contribution to primordial non-gaussianity
	[or more complex parametrization of non-gaussianity]
r	tensor-to-scalar ratio
$r + 8n_{\mathrm{T}}$	violation of the inflationary consistency equation
$dn_{\rm T}/d\ln k$	running of the tensor spectral index
\mathcal{P}_S	CDM isocurvature perturbation
n_S	and its spectral index
PSR	and its correlation with adiabatic perturbations
$n_{SR} - n_S$	and the spectral index of that correlation
	[or more complicated multi-component isocurvature perturbation]
C.	cosmic string component of parturbations

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Problem 2: as we add extra parameters, the uncertainties on existing parameters increase, and eventually we learn nothing useful about anything.

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Problem 1: if we add extra parameters, typically the maximum likelihood will increase, even if the new parameter actually has no physical relevance.

Problem 2: as we add extra parameters, the uncertainties on existing parameters increase, and eventually we learn nothing useful about anything.

We need a way of penalizing use of extra parameters - an implementation of Ockham's razor

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Akaike information criterion (Akaike 1974)

Bayesian information criterion (Schwarz 1978)

Bayesian evidence

(Jeffreys 1961 etc)

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Akaike information criterion

$$AIC = -2\ln \mathcal{L}_{max} + 2k$$

Bayesian information criterion

$$BIC = -2 \ln \mathcal{L}_{max} + k \ln N$$

Bayesian evidence

(Akaike 1974)

k = number of parameters

(Schwarz 1978)

N = number of datapoints

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Bayesian evidence

$$E = \int d\theta \, \mathcal{L}(\theta) \, \mathrm{pr}(\theta)$$

 θ = parameter vector, pr = prior

The preferred model is the one which minimizes the information criterion, or maximizes the evidence.

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The Bayesian evidence is the most powerful of these. It is a full implementation of Bayesian inference, and literally gives the probability of the data given the model (note not the probability of particular parameter values). If multiplied by the prior model probability it gives the posterior model probability. However it can be hard to calculate, being a highly-peaked multi-dimensional integral.

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The Akaike Information Criterion was derived using information theory techniques. It gives an approximate minimization of the so-called Kullback-Leibler information entropy, which is a measure of the difference between two probability distributions and the difference between two probability distributions are distributed and the difference between two probability distributions and the difference between two probability distributions are distributed as a distribution of the difference between two probabilities and the difference between two probabilities an

Model selection techniques are essential when considering whether or not new data requires the addition of new parameters to describe it.

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Assuming that the density is the only parameter, with a uniform prior from 0.1 to 2, and likelihood $(\Omega - 1.02)^2$

- Flat: Evidence = $L(\Omega = 1) = 0.6L_0$
- Curved: Evidence = $\frac{1}{1.9} \int \mathcal{L}(\Omega) d\Omega \simeq 0.03 \mathcal{L}_0$

According to the evidence, the flat model is a better description of the data, with odds of about 20:1 against the curved model.

Note that this assumes flat and curved were thought equally likely before the data came along.

WMAP says
$$\Omega_{\text{tot}} = 1.02 \pm 0.02$$

This has been widely interpretted as supporting the idea of a flat Universe, but actually favouring a slightly closed Universe.

Assuming that the density is the only parameter, with a uniform prior from 0.1 to 2, and likelihood $(\Omega - 1.02)^2$

■ Flat: Evidence = $L(\Omega = 1) = 0.6L_0$

■ Curved: Evidence = $\frac{1}{1.9} \int \mathcal{L}(\Omega) d\Omega \simeq 0.03 \mathcal{L}_0$

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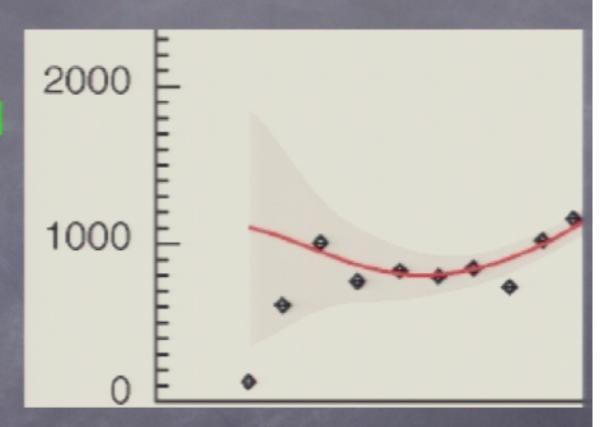
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New physics from low quadrupole??

[Argument roughly following Niarchou et al]



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New physics from low quadrupole??

[Argument roughly following Niarchou et al]

If you want to explain this with new physics, you have to introduce new parameters, for



which you will be penalized. As the discrepancy is only at the 95% level, the gain in fit will never compensate for this penalty.

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Something like 95% of all 95% confidence "detections" turn out to be wrong. Why?

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Statistical fluke: By definition important only if people

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- Statistical fluke: By definition important only if people do their error analysis wrongly.
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- Inappropriate "a posteriori" reasoning: choosing "interesting" features from the data and assessing their significance via Monte Carlo analyses.
- Neglect of model dimensionality: using parameter estimation rather than model selection.

Beltran, Garcia-Bellido, Lesgourgues, Liddle, Slosar, PRD, astro-ph/0501477

Even if the real perturbations are adiabatic, some level of isocurvature perturbations will always be allowed.

While parameter estimation techniques can only place upper limits on the isocurvature modes, model selection can give positive support to simpler models.

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While parameter estimation techniques can only place upper limits on the isocurvature modes, model selection can give positive support to simpler models.

We consider the three observationally-distinct classes of isocurvature mode, CDI, NID and NIV. Only one type of mode is permitted per model, but with arbitrary spectral index and correlation to adiabatic: 4 extra parameters. We compare with two adiabatic models, one with n=1 and one with n varying.

The Bayesian Evidence was computed using a technique called thermodynamic integration. This is an MCMC method where the chains are heated in order to fully explore the prior space (parameter estimation chains sample the posterior which is usually localized to a small fraction of the prior).

We tested several variants on this scheme. Accurate determination of the evidence required approximately 10⁷ likelihood evaluations per model, making it a supercomputer class problem.

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Jeffreys Scale:

 $\Delta \ln E < 1$

Not worth more than a bare mention

 $1 < \Delta \ln E < 2.5$

Substantial evidence

 $2.5 < \Delta \ln E < 5$

Strong to very strong evidence

 $5 < \Delta \ln E$

Decisive evidence

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Note that the results depend on the priors chosen. Our prior range covers the complete range from all adiabatic to all isocurvature using the relative fraction. We use two different parametrizations to test robustness.

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Model	In(Evidence)*		
	Parametrization 1	Parametrization 2	
AD-HZ	0.0 ± 0.1		
AD-n	0.0 ± 0.1		
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AD-n	0.0 ± 0.1		
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NID	-1.0 ± 0.2	-2.0 ± 0.2	
NIV	-1.0 ± 0.3	-2.3 ± 0.2	

*Normalized to the AD-HZ value In(Evidence)=-854.1

Mukherjee, Parkinson and Liddle, astro-ph/0508461

The main lesson from that work is that a more efficient algorithm is needed to make computations feasible. We have recently implemented Skilling's Nested Sampling algorithm for cosmology.

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Mukherjee, Parkinson and Liddle, astro-ph/0508461

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Skilling (2004) rewrote the evidence as

$$E = \int \mathcal{L}(\theta) \operatorname{pr}(\theta) d\theta = \int_0^1 \mathcal{L}(X) dX$$

where X is the fractional prior mass.

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Mukherjee, Parkinson and Liddle, astro-ph/0508461

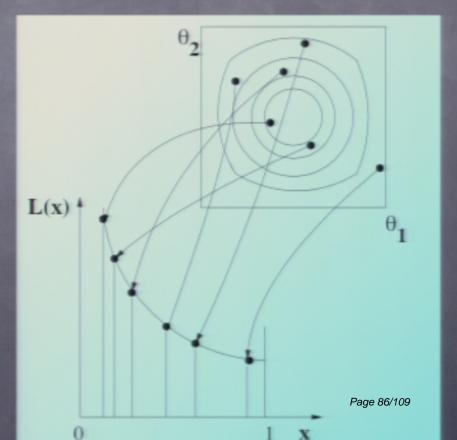
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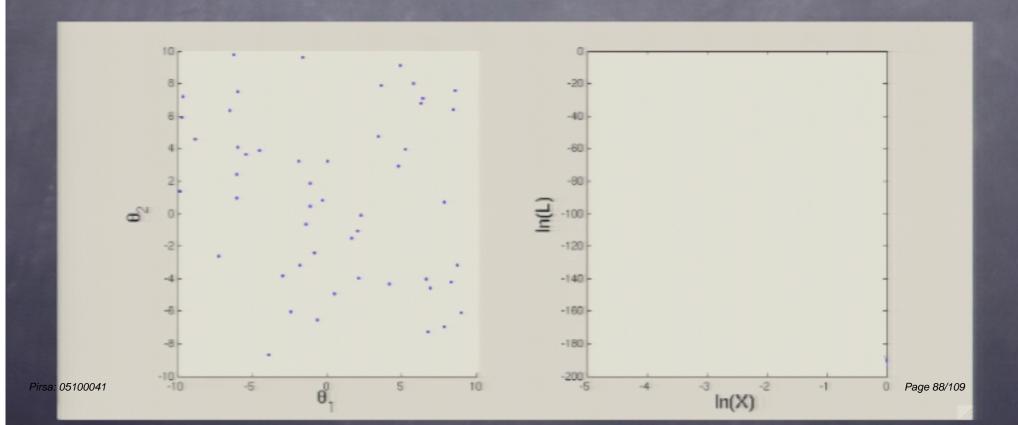
This can then be evaluated using Monte Carlo samples to trace the variation of likelihood with prior mass, peeling away thin nested isosurfaces of equal likes 15100010d.



The method `walks' a set of points (eg 300) into the high-likelihood region using replacement. The main difficulty in implementing the algorithm successfully is in efficiently generating replacement points which are uniformly sampled from the remaining prior volume.

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A model selection example

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Model	$\Lambda \text{CDM+HZ}$	$\Lambda \text{CDM} + n_{\text{s}}$	$\Lambda {\rm CDM} + n_{\rm s}$	HZ+w	$w + n_{\rm s}$
			(wide prior)		
$n_{ m s}$	1	0.8 - 1.2	0.6 - 1.4	1	0.8 - 1.2
w	-1	-1	-1	$-\frac{1}{3}$ 1	$-\frac{1}{3}$ 1
e.f	1.5	1.7	1.7	1.7	1.8
$N_{\rm like}(\times 10^4)$	8.4	17.4	16.7	10.6	18.0
$\ln E$	0.00 ± 0.08	-0.58 ± 0.09	-1.16 ± 0.08	-0.45 ± 0.08	-1.52 ± 0.08

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At the moment the more complex models are not excluded, but they are mildly disfavoured against the simplest model.

Model selection for survey comparison/design

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- Fisher matrix approach:
 - simulate data for a fiducial model (eg LambdaCDM); estimate expected parameter uncertainties about that model; interpret as excluding models outside the contours
- Bayes factor approach:

simulate data at each point in parameter plane; compute Bayes factor (ie evidence ratio) of full model versus

Pirsa: 05100041 LambdaCDM at each point.

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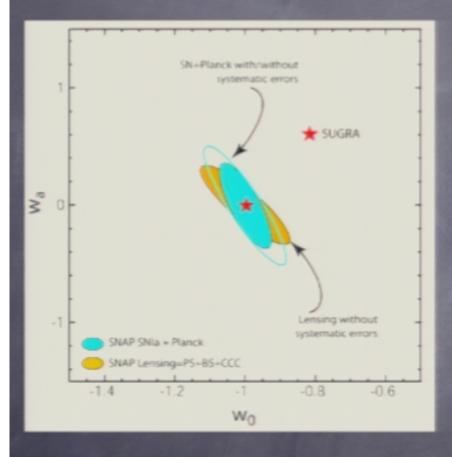
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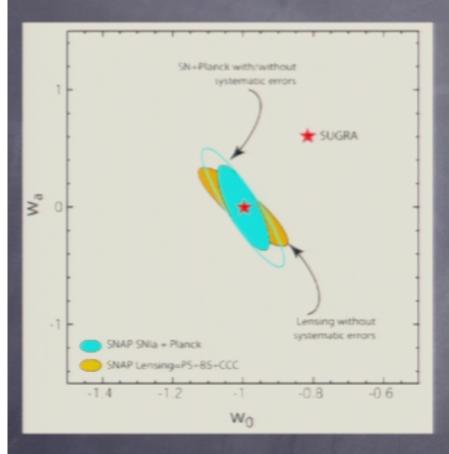
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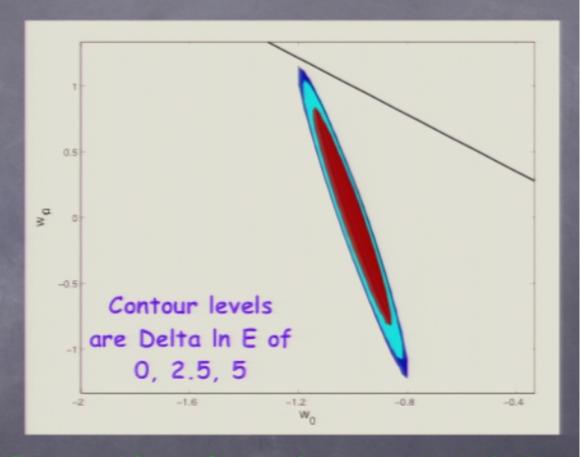
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Projected Fisher matrix uncertainties about LambdaCDM Pirsa: 05100041 (SNAP collaboration)

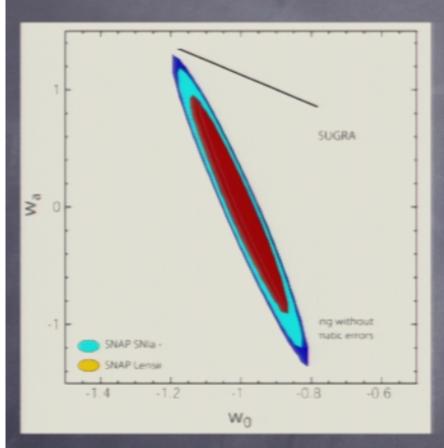
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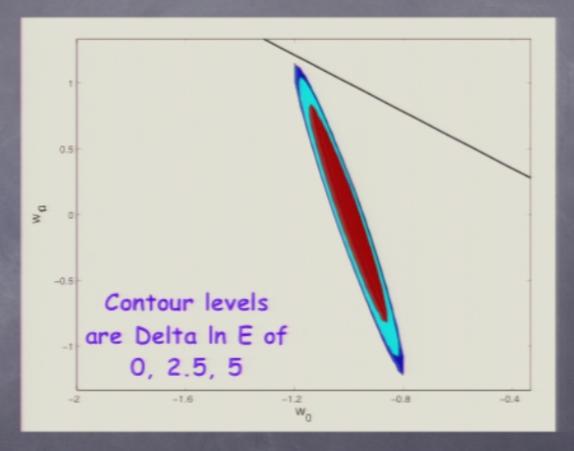




Projected Fisher matrix uncertainties about LambdaCDM Pirsa: 05100041 (SNAP collaboration) Projected Bayes factor plot against LambdaCDM, SNAP supernovae only with Omega_matter prior (myself and collaborators, in press 102/109

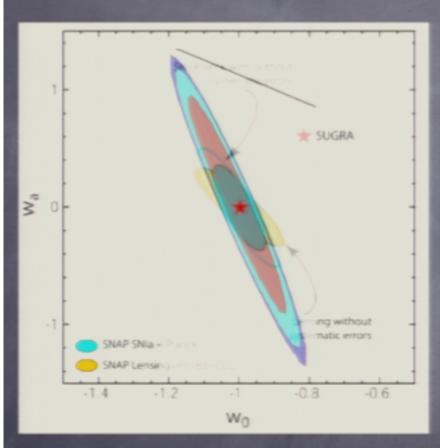
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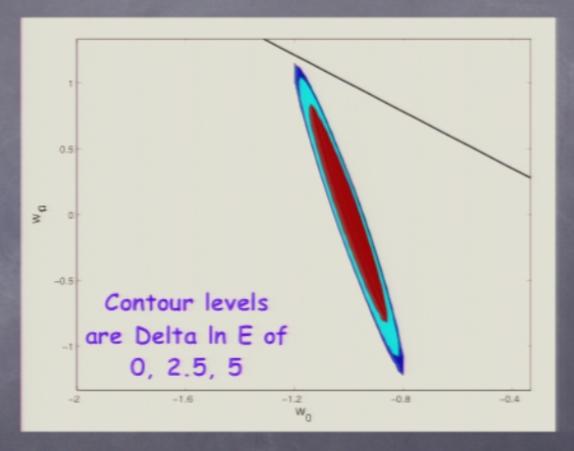




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- The Bayesian evidence is the most powerful available tool. It is challenging to compute but nested sampling makes it feasible.
- An application to adiabatic models shows current data are comparably well explained by the Harrison-Zel'dovich model and a varying spectral index model (prior 0.8 < n < 1.2).</p>

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